Towards Principled Training and Serving of Large Language Models



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By

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Abstract

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Large language models (LLMs), powered by neural networks with billions to trillions of parameters, face critical challenges in training efficiency and deployment scalability due to their computational demands. This thesis addresses these challenges through two key contributions: advancing reinforcement learning from human feedback (RLHF) for post-training and optimizing LLM serving via novel caching strategies.

First, we provide a comprehensive theoretical analysis of RLHF, proposing algorithms with near-optimal sample complexity for reward learning. We validate these proposed algorithms through real-world case studies, including the development of Starling-7B, an RLHF-aligned model that demonstrates strong performance in human preference benchmarks.

Second, we design near-optimal caching algorithms tailored for LLM inference, reducing computational overhead while preserving output quality. Our framework achieves significant latency reductions in LLM serving environments.

Our work bridges theoretical analysis with practical implementation, offering insights into scalable alignment techniques and efficient deployment strategies. The results highlight the viability of RLHF for LLM post-training and the importance of system-level optimizations for sustainable LLM adoption.

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Chapter 1 Introduction

Large Language Models (LLMs) have emerged as a transformative technology in artificial intelligence (AI), demonstrating unprecedented capabilities in tasks ranging from translation and summarization to code generation, complex reasoning, and agentic tasks [266, 205, 28, 160, 118, 6, 12, 256, 206, 53, 262, 29, 228, 218, 179, 120]. Recent progress in LLMs is reshaping not only natural language processing but also broader AI applications. In this chapter, we explore the comprehensive landscape of LLMs, from their theoretical foundations to practical deployment challenges.

1.1 Overview of Large Language Models

The theoretical foundations of language modeling trace back to Claude Shannon's pioneering work in information theory [236, 235], which introduced fundamental concepts like entropy and established methodologies for next-symbol prediction in sequences. Early Natural Language Processing (NLP) systems employed basic preprocessing techniques such as punctuation removal and stemming [171], combined with statistical methods including N-gram models [23] and Hidden Markov Models (HMMs) for part-of-speech tagging [204, 125]. These approaches, while effective for their time, faced fundamental limitations in capturing semantic relationships and long-range dependencies.

A paradigm shift occurred with the introduction of distributed word representations through embeddings like Word2Vec [180] and GloVe [198], which encoded semantic relationships in dense vector spaces. The field subsequently witnessed the rise of neural architectures, beginning with Recurrent Neural Networks (RNNs) [217, 132] and their enhanced variants Long Short-Term Memory (LSTM) networks [108] and Gated Recurrent Units (GRUs) [51], which addressed the vanishing gradient problem through sophisticated gating mechanisms.

The attention revolution began with Bahdanau et al.'s seminal work on sequence-tosequence learning [14], culminating in the Transformer architecture [266] that replaced recurrence with self-attention mechanisms. This innovation enabled parallel computation and superior handling of long-range dependencies through positional encoding strategies like absolute positional embeddings [266] and Rotary Position Embeddings (RoPE) [253]. The Transformer's scalability fueled an era of massive language models, beginning with GPT [205] and BERT [118], which demonstrated unprecedented few-shot learning capabilities through scaled-up architectures [135].

Two key components have led to the development of modern LLMs: First, the scaling of data size and model size during pre-training creates strong base models. Second, new advanced post-training paradigms like Reinforcement Learning from Human Feedback (RLHF) significantly increases the helpfulness and harmlessness of the model as a chatbot.

Recent years have witnessed explosive growth in both proprietary and open-source LLMs. The GPT series progressed through GPT-2 [206], GPT-3 [28], and GPT-4 [6], incorporating multimodal capabilities and improved reasoning. Competing proprietary models like Anthropic's Claude series [12] and Google's Gemini series [255] have pushed boundaries in capability and multimodality. The open-source community has responded with models series like MPT [258], Falcon [11], LLaMA [262, 68], Mistral [127], Qwen [16, 291], OLMo [92], Gemma [179, 257], Phi [96, 4, 5] and Deepseek [166, 167], while coding-focused models such as StarCoder [160], CodeLlama [216], Qwen-Coder [114] and agent-focused models such as Gorilla [196], NexusRaven [249], xLAM [304] demonstrate domain specialization. There have also been efforts in further post-training existing base models to improve the capability and safety, including but not limited to Alpaca [254, 69], Vicuna [49], OpenChat [268], Starling [319], WizardLM [286], Zephyr [264], Tulu [151] and Athene [81].

As the scaling of the size for the language models approaches diminishing returns, the research community has shifted focus toward optimizing test-time computation [245, 95]. Notable advancements in this direction include OpenAI's o1 [120] and Deepseek's R1 [97], which demonstrate significant performance gains on reasoning-intensive tasks such as code generation and mathematical problem-solving. Regarding methodological advancements, recent post-training paradigms have increasingly favored reinforcement learning (RL) over supervised fine-tuning (SFT), a trend attributed to evidence suggesting that online RL approaches may generalize better than offline training methods like SFT. Concurrently, researchers have transitioned from human-feedback-trained reward models to directly applying deterministic and verifiable reward signals for complex domains like math and coding. This shift stems from the demonstrated limitations of both LLMs and conventional reward models in reliably evaluating such domains.

1.2 Model Architecture and Components

Most of the current LLMs build upon the Transformer architecture with several sophisticated enhancements. Advanced tokenization methods using Byte Pair Encoding (BPE) [231] and WordPiece [281] have improved the handling of vocabulary and rare words. The architecture incorporates various positional embedding approaches, from absolute positional encodings [266] to more recent innovations like RoPE [253]. Attention mechanisms have been optimized through innovations such as grouped-query attention [10] and sparse attention patterns [300]. The introduction of Mixture of Experts (MoE) architectures [119, 70, 76, 155] has further enhanced parameter efficiency and model scalability. There have been some recent attempts in improving original attention mechanism from quadratic in the context length to linear in the context length, including state-space-model-based approaches [133, 99, 94, 93], and other variants of RNNs or attentions [202, 136, 52].

1.3 Training LLMs

The training process for LLMs can usually be decomposed into two major phases, pre-training and post-training.

Pre-Training

Pre-training serves as the foundational stage, where the model is exposed to vast, diverse text corpora (e.g., web pages, books, and scientific articles) through self-supervised learning. During this phase, the model learns to predict masked tokens or subsequent words in a sequence, enabling it to internalize linguistic patterns, world knowledge, and rudimentary reasoning abilities. The pre-training process equips the model with a broad, general-purpose understanding of language and context, albeit without task-specific or safety-aligned behaviors.

The training pipeline for large language models (LLMs) begins with tokenization, a preprocessing step where raw text is converted into discrete subword units (e.g., morphemes, common phrases) mapped to integer indices. Subword tokenization algorithms like Byte-Pair Encoding (BPE) [231] or WordPiece [281] balance vocabulary efficiency with the ability to handle rare or out-of-vocabulary terms.

After tokenization, the pre-training phase leverages self-supervised learning on internetscale datasets. The model learns to predict the next token in a sequence conditioned on preceding tokens, a paradigm popularized by autoregressive architectures like GPT [205] and masked language modeling in BERT [65]. For a given set of text corpus with tokens u_1, u_2, \dots, u_N , the pre-training phase for an autoregressive architecture is directly based on minimizing cross entropy loss:

$$\min_{\theta \in \Theta} \mathcal{L}^{\mathsf{pre}}(\theta) = \min_{\theta \in \Theta} \sum_{i=1}^{N} -\log P_{\theta}(u_i \mid u_1, u_2, \cdots, u_{i-1}),$$

where P_{θ} is a parameterized language model. Due to computational constraints, sequences are typically truncated to a fixed context window (with a common range of 4,000-2,000,000 tokens), discarding distant dependencies while preserving local coherence.

Empirical studies have established that pre-training loss reduction correlates with qualitative improvements in downstream task performance [267, 57]. Notably, as model size (parameters) and training data scale, LLMs exhibit emergent capabilities—abilities absent in smaller models, such as complex reasoning, in-context learning, and instruction following

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[273]. This phase equips the model with a foundational understanding of syntax, semantics, and domain-agnostic world knowledge, which subsequent post-training stages refine into task-specific, safe, and aligned behaviors.

Post-Training

Post-training encompasses subsequent stages aimed at improving the model's capabilities and aligning it with human preferences. Unlike pre-training, post-training leverages smaller, high-quality datasets to achieve objectives like:

- Helpfulness: Optimizing responses to understand user intent and follow user instructions, as seen in chatbot applications.
- Safety and ethics: Mitigating harmful, biased, or untruthful outputs through human-in-the-loop oversight.
- Reasoning: Improving the reasoning capability of the model via training it to output longer thinking process before answering the question.
- Task-specific adaptation: Enhancing performance on specialized domains (e.g., coding or function calling).

Post-training phase often includes techniques such as supervised fine-tuning (SFT) on curated prompt-response datasets, reinforcement learning from human feedback (RLHF), and Reinforcement Learning (RL) from verifiable signals. A crucial advancement in LLM training has been the introduction of RL. The creation of ChatGPT has benefited mostly from the application of RLHF [55, 191]. The initial RL algorithm used in ChatGPT training is Proximal Policy Optimization (PPO) [229]. After the release of ChatGPT, SFT has notably advanced the development of open LLMs, particularly with the utilization of high-quality data distilled from GPT-4 or Deepseek R1. This is demonstrated effectively by models such as Alpaca [254], Vicuna [49], Openchat-3.5 [268] and Sky T-1 [260].

The field has also seen the development of alternative RL algorithms to PPO, such as Direct Preference Optimization (DPO) [207], Group Relative Policy Optimization (GRPO) [237], Kahneman-Tversky Optimization (KTO) [72], SimPO [177], Pairwise Proximal Policy Optimization (P3O) [280], Advantage-induced Policy Alignment (APA) [318] and Reinforce LeaveOne-Out (RLOO) [142, 7] etc. Based on early experiments, there have been observations where online reinforcement learning like PPO tends to outperform offline reinforcement learning methods like DPO and produces models with higher quality [287, 116].

The computational demands of training these models have spurred innovations in distributed training strategies. Advanced optimization techniques and codebases like DeepSpeed ZeRO [208] and FSDP [308] have enabled efficient training across multiple devices. Efficient IO-aware exact attention algorithm like FlashAttention [62, 61, 232] and parameter-efficient fine-tuning approaches such as LoRA [111] and QLoRA [64] have made model adaptation more accessible.

1.4 Evaluating LLMs

The rapid advancement of LLMs necessitates equally sophisticated evaluation methodologies. Effective evaluation frameworks not only differentiate the capabilities of current models, but also identify directions for improvement through measurable benchmarks. While traditional metrics like perplexity, BLEU [194], and ROUGE [164] offer quantitative insights, they increasingly fail to capture the full spectrum of capabilities exhibited by modern LLMs. This gap has spurred the development of comprehensive evaluation frameworks tailored to complex language understanding and generation tasks.

LLM benchmarks are primarily static, ground-truth-based assessments across diverse domains such as general knowledge (MMLU [104], MMLU-Pro [271]), mathematics (MATH [105], GSM-8K [58]), coding (HumanEval [40], Bigcode Bench [321]), and reasoning (DROP [67], BigBench [250]). Domain-specific evaluations like AGIEval [313] for human exams, GPQA [213] for expert-level questions, and HellaSwag [303] for commonsense reasoning further extend this landscape. Holistic initiatives such as HELM [163] aggregate multiple dimensions of model performance. Task-oriented benchmarks have also emerged, exemplified by IFEval [315] for instruction following and SWE-Bench [129] for software engineering. There are also benchmarks focusing on LLMs as an Agent, including AgentBench [169], NexusBench [259], Berkeley Function Calling Leaderboard [290] and Tool Sandbox [172] and τ -Bench [293].

Static benchmarks are susceptible to test set contamination, as demonstrated by studies [33, 226, 292]. DynaBench [137] formalizes this need through its proposal for continuously evolving benchmarks. There have been similar evolving benchmarks like LiveCodeBench [122] for coding, LiveBench [275] for general tasks, and R2E [123] for reasoning. Community-driven evaluation frameworks like Chatbot Arena [48] mitigate these limitations by enabling real-time human comparison of LLM responses to identical prompts. The platform employs the Bradley-Terry-Luce model to statistically derive model rankings from crowd-sourced preference judgments.

To address scalability challenges in human evaluation, recent work explores LLM-as-ajudge paradigms. Frameworks such as AlpacaFarm [69], MT-bench [309], and AlpacaEval [162] show promising alignment with human preferences, while Arena Hard [161] targets harder queries sampled from Chatbot Arena. However, these approaches exhibit limitations in domains requiring precise verification, particularly for mathematical reasoning and complex coding tasks, where human expertise remains essential for reliable assessment.

1.5 Serving LLMs

The production deployment of LLMs introduces multifaceted engineering challenges [63], requiring careful optimization across three key dimensions: computational efficiency (inference latency, memory footprint), operational costs, and model quality preservation [278]. Modern solutions employ a layered optimization strategy:

- Model Compression: Fundamental architectural adaptations through knowledge distillation into a smaller model [227], structured pruning [79, 42], dynamic sparsification techniques [22, 158], and precision quantization [117, 305] reduce baseline resource requirements. These techniques minimize the size of active parameters during inference.
- Inference Optimization: Serving infrastructure now integrates new algorithms like speculative decoding [38, 157] to accelerate token generation. Hardware-aware implementations leverage flash attention mechanisms [62, 61, 232] and streaming execution [283] to maximize throughput. Specialized inference systems including vLLM [149] and SGLang [312] optimize memory management through KV cache optimization and request batching, while frameworks like TensorRT enable hardware-specific kernel fusion.
- System-Level Enhancements: Advanced resource schedulers [184, 239] and adaptive caching strategies [320] address scalability challenges in multi-tenant deployments. By caching the outputs of previously processed inputs, inference times can be significantly reduced for recurring requests, thus improving both latency and throughput. This thesis will explore optimal caching algorithms for scheduling inference requests.

1.6 Organization of the Thesis

This thesis addresses two fundamental challenges in Large Language Models (LLMs):posttraining with RLHF and optimizing inference efficiency via near-optimal caching strategies. The structure progresses from theoretical foundations to practical implementations, culminating in real-world validation.

The remainder of this work is organized as follows:

- Chapter 2 (Theoretical Analysis of RLHF): Presents a rigorous mathematical framework for RLHF, analyzing its convergence properties and alignment capabilities through the lens of preference optimization.
- Chapter 3 (Practical RLHF Implementation): Introduces pessimism-based RLHF algorithms to address reward overoptimization, with experimental validation through the Starling-7B case study demonstrating improved ethical alignment.
- Chapter 4 (Real-world RLHF Experiments: Starling-7B): A real-world RLHF experiments that first collects 7-wise comparison data, trains the reward model, and fine-tunes an existing language model with the proposed RLHF algorithm.
- Chapter 5 (Efficient LLM Serving): Develops optimal caching strategies for inference scheduling, combining LRU/LFU fundamentals with novel cost-aware algorithms that adapt to LLM-specific computation patterns.

Chapter 2

Theoretical Analysis of RLHF

2.1 Introduction

RLHF first learns a reward from human feedback, in the form of pairwise or K-wise comparisons between actions (responses), and trains the LLM with RL on the learned reward. In this chapter, we take the first step towards providing a theoretical framework for RLHF, with a specific focus on reward learning. We provide theoretical analysis that justifies the empirical success of RLHF in InstructGPT and ChatGPT, along with new insights for algorithm design.

Taking InstructGPT [191] as an example, a typical deployment of RLHF for language modeling includes the following steps:

(a) Fine-tune an LLM using supervised fine-tuning.

(b) Train a reward model based on the LLM in the first step with human feedback.

(c) Further fine-tune the existing LLM based on the learned reward model using Proximal Policy Optimization (PPO).

During the reward training step, the prompts are first sampled from a pre-collected dataset. Then K responses are sampled by executing existing models on the sampled prompts. Based on the prompt provided, a human labeler ranks all the responses according to her own preference. The reward model is trained based on a maximum likelihood estimator (MLE), also known as the learning-to-rank algorithm or cross-entropy minimization [168, 282, 32, 55, 191].

In the setting of InstructGPT, the ranking of responses is based purely on the current prompt, which can be viewed as the state in a contextual bandit. We accordingly start with the setting of a contextual bandit, and later generalize our results to Markov Decision Process (MDP) where there are transitions between states. Let S be the set of states (prompts), and A be the set of actions (responses). For each state-action pair (s, a), we assume that the reward is parametrized by $r_{\theta}(s, a) = \langle \theta, \phi(s, a) \rangle$ for some known and fixed feature function $\phi(s, a) : S \times A \mapsto \mathbb{R}^d$. In an LLM, such a ϕ is usually derived by removing the last layer of the pre-trained model.¹ We denote the ground-truth reward provided by a human as $r_{\theta^{\star}}(s, a)$ for some parameter $\theta^{\star} \in \mathbb{R}^d$.

We are interested in the sample complexity for learning a reward model r_{θ^*} from pairwise or *K*-wise comparison data. For the *i*-th sample, a state s^i is first sampled from some fixed distribution ρ . Given the state s^i , *K* actions $(a_0^i, a_1^i, \dots, a_{K-1}^i)$ are sampled from some joint distribution $\mathbb{P}(a_0, \dots, a_{K-1} \mid s^i)$. Let $\sigma^i : [K] \mapsto [K]$ denote the output of the human labeller, which is a permutation function representing the ranking of the actions. Here $\sigma^i(0)$ represents the most preferred action. We assume that the distribution of σ^i follows a Plackett-Luce (PL) model [200, 173]:

$$\mathbb{P}(\sigma^{i} \mid s^{i}, a_{0}^{i}, a_{1}^{i}, \cdots, a_{K-1}^{i}) = \prod_{k=0}^{K-1} \frac{\exp(r_{\theta^{\star}}(s^{i}, a_{\sigma^{i}(k)}^{i}))}{\sum_{j=k}^{K-1} \exp(r_{\theta^{\star}}(s^{i}, a_{\sigma^{i}(j)}^{i}))}$$

When K = 2, this reduces to the pairwise comparison of the Bradley-Terry-Luce (BTL) model [25], which is widely applied in existing RLHF algorithms [55, 191].

Since the learned reward model is mainly used for downstream policy training, we measure the correctness of the estimated reward model via the performance of a greedy policy trained from a reward model $r_{\hat{\theta}}$. Concretely, for a greedy policy $\hat{\pi}(s) = \arg \max_a r_{\hat{\theta}}(s, a)$, we compute a performance gap compared to the optimal policy:

$$\mathsf{SubOpt}(\hat{\pi}) \coloneqq \mathbb{E}_{s \sim \rho}[r_{\theta^{\star}}(s, \pi^{\star}(s)) - r_{\theta^{\star}}(s, \hat{\pi}(s)].$$

Here $\pi^* = \arg \max_a r_{\theta^*}(s, a)$ is the optimal policy under the true reward r_{θ^*} .

It has been observed in [85] that in the reward model trained from practice, there exists an over-optimization phenomenon where the true reward first increases and then decreases during the policy optimization stage. In this chapter, we study the potential sub-optimality of the MLE in the RLHF setting. As a by-product, we also provide guarantee of the estimation error on the semi-norm of the parameter estimation error, $\|\hat{\theta} - \theta^{\star}\|_{\Sigma}$, for a query-dependent covariance matrix Σ .

From a broader perspective, the framework of RLHF can be viewed as a special case of reward learning from pre-collected data, which has been a primary focus in Inverse Reinforcement Learning (IRL) and offline reinforcement learning. Our techniques also provide theoretical guarantee for the max-entropy IRL [322] and action-based IRL algorithms [210, 186, 78].

Main Results

Pairwise Comparison. We start with the setting of a contextual bandit with pairwise comparison. We focus on two algorithms, MLE and pessimistic MLE. The following result

¹In InstructGPT, the function ϕ is still parametrized can be further trained in the reward learning step. However, for simplicity of theoretical analysis we assume in this chapter that ϕ is fixed and one only fine-tunes the last layer with parameter θ .

from dueling bandits and RL [74, 192] shows that under a semi-norm $\|\cdot\|_{\Sigma}$, MLE converges to the true parameter.

Lemma 1 (Informal). Under certain regularity conditions, the MLE satisfies the following with probability at least $1 - \delta$,

$$\|\hat{\theta}_{\mathsf{MLE}} - \theta^{\star}\|_{\Sigma_{\mathcal{D}}} \leq C \cdot \sqrt{\frac{d + \log(1/\delta)}{n}}.$$

Here $\Sigma_{\mathcal{D}} = \frac{1}{n} \sum_{i=1}^{n} (\phi(s^{i}, a^{i}_{1}) - \phi(s^{i}, a^{i}_{0}))(\phi(s^{i}, a^{i}_{1}) - \phi(s^{i}, a^{i}_{0}))^{\top}.$

However, when we consider the performance of the induced policy, MLE provably fails while pessimistic MLE gives a near-optimal rate. In essence, the pessimism principle discounts actions that are less represented in the observed dataset, and hence is conservative in outputting a policy.

Theorem 2 (Informal). Under certain coverage assumption, one can design a pessimistic MLE such that the induced greedy policy $\hat{\pi}_{\mathsf{PE}}$ is good; i.e., with probability at least $1 - \delta$,

$$\mathsf{SubOpt}(\hat{\pi}_{\mathsf{PE}}) = \Theta\left(\sqrt{\frac{d + \log(1/\delta)}{n}}\right)$$

In contrast, under the same assumption, one can find instances such that the greedy policy w.r.t. MLE $\hat{\pi}_{MLE}$ fails:

$$\forall n > 1, \mathbb{E}[\mathsf{SubOpt}(\hat{\pi}_{\mathsf{MLE}})] \ge 0.1.$$

K-wise Comparison. For K-wise comparison, we analyze both the MLE and the algorithm in InstructGPT [191] which splits the ranking data into K(K-1)/2 pairwise comparison data and runs an MLE based on the BTL model. We show that both converge in terms of the estimation error under the semi-norm, and give a near-optimal policy when combined with pessimism. More importantly, we show that although both estimators are unbiased, the asymptotic variance of MLE is smaller than that of the splitted estimator in InstructGPT [191], which belongs to the family of M-estimators. Thus the MLE is more efficient than the existing algorithm used in InstructGPT. We also conduct experiments to verify the theoretical prediction.

Let the estimated parameter for the splitted estimator be $\hat{\theta}$ and the induced policy be $\hat{\pi}_{\mathsf{PE}}$. We have:

Theorem 3 (Informal). Under certain coverage and regularity conditions, the following holds separately with probability at least $1 - \delta$:

$$\|\hat{\theta} - \theta^{\star}\|_{\Sigma_{\mathcal{D}}} \le C \cdot \sqrt{\frac{d + \log(1/\delta)}{n}},$$

SubOpt $(\hat{\pi}_{\mathsf{PE}}) \le C' \cdot \sqrt{\frac{d + \log(1/\delta)}{n}}$

Here $\Sigma_{\mathcal{D}} = \frac{2}{K(K-1)n} \left(\sum_{i=1}^{n} \sum_{j=0}^{K-1} \sum_{k=j+1}^{K-1} (\phi(s^i, a^i_j) - \phi(s^i, a^i_k)) (\phi(s^i, a^i_j) - \phi(s^i, a^i_k))^\top \right).$

We also extend our results to the case of MDP and IRL; see the detailed presentation in Section 2.4 and Section 2.5. Let the estimated parameter be $\hat{\theta}$ and the induced pessimistic policy be $\hat{\pi}_{\mathsf{PE}}$. For pairwise comparison we have:

Theorem 4 (Informal). In the MDP setting with horizon H, under certain coverage and regularity conditions, the following holds separately with probability at least $1 - \delta$:

$$\begin{split} \|\hat{\theta} - \theta^{\star}\|_{\Sigma_{\mathcal{D}}} &\leq C \cdot \sqrt{\frac{d + \log(1/\delta)}{n}},\\ \mathsf{SubOpt}(\hat{\pi}_{\mathsf{PE}}) &\leq C' \cdot \sqrt{\frac{d + \log(1/\delta)}{n}},\\ Here \ \Sigma_{\mathcal{D}} &= \frac{1}{n} \sum_{i=1}^{n} (\sum_{h=0}^{H} (\phi(s_{h}^{i}, a_{h}^{i}) - \phi(s_{h}^{i\prime}, a_{h}^{i\prime}))) \ (\sum_{h=0}^{H} (\phi(s_{h}^{i}, a_{h}^{i}) - \phi(s_{h}^{i\prime}, a_{h}^{i\prime})))^{\top}. \end{split}$$

Our results not only explain the correctness of existing algorithms, but also provide new insights for algorithm design in RLHF. In particular, it suggests the importance of introducing pessimism in the reward learning part, which can be implemented via adding regularization in policy training steps as in [191], or using existing offline RL algorithms, including but not limited to Conservative Q-Learning [145], Implicit Q-Learning [144] and Adversarially Trained Actor Critic [46]. On the other hand, it also shows that MLE is a more efficient estimator than that in [191].

Related Work

Learning and Estimation from Pairwise Comparison and Ranking. The problem of estimation and ranking from pairwise or K-wise comparisons has been studied extensively in the literature. In the literature of *dueling bandit*, one compares two actions and aims to minimize regret based on pairwise comparisons [299, 324, 298, 297, 225, 88, 223, 9, 325, 141, 83, 222, 224, 74]. [189, 289] analyze the sample complexity of dueling RL under the tabular case, which is extended to linear case and function approximation by the recent work [192, 44]. [37] studies a close setting where in each episode only binary feedback is received. However, most of the work focuses on regret minimization. We take a first step towards the theoretical analysis for function approximation for K-wise comparisons with policy learning as the target.

On the other hand, in the literature of ranking, most of the theoretical work focuses on the tabular case where the rewards for different actions are uncorrelated [77, 233, 234, 103, 176, 124, 43, 45, 209, 185, 98, 102]. And a majority of the empirical literature focuses on the framework of learning to rank (MLE) under general function approximation, especially when the reward is parameterized by a neural network [168, 282, 32, 55, 191, 27, 241, 31, 276, 277, 56, 3]. Similar idea of RL with AI feedback also learns a reward model from preference [17], except for that the preference is labeled by another AI model instead of human. **Inverse Reinforcement Learning and Offline Reinforcement Learning.** RLHF, IRL and offline learning are all approaches that can be used to incorporate human preferences or expertise into the decision-making process of an agent. However, they differ in the way that they use human input to guide the agent's behavior. In IRL and imitation learning, we only observe an expert's behavior and would like to infer the expert's preferences or goals [187, 2, 322, 210, 186, 107, 78, 115]. In offline learning, we directly observe the cardinal rewards for the state. But the actions are likely to be sub-optimal. In RLHF, we observe ordinal comparisons between pairs or a set of actions. In one of the popular IRL frameworks, max-entropy IRL [322], it is also assumed that human choice follows a PL model. We unify the problem of RLHF and max-entropy IRL, and provide the first sample complexity analysis for max-entropy IRL.

Pessimism in Offline RL. The idea of introducing pessimism for offline RL has been studied in recent year [131, 212, 159, 285, 301, 302, 284, 288]. In this chapter, we connect RLHF with offline RL and show that pessimism also helps in RLHF.

Preliminaries

We begin with the notation that we use in the paper. Then we discuss our formulations of contextual bandits and Markov decision processes, and introduce the data collection model and the BTL and PL models.

Notations. We use calligraphic letters for sets, e.g., S and A. Given a set S, we write |S| to represent the cardinality of S. For vectors x and y, we use $\langle x, y \rangle = x^{\top}y$ to denote their inner product. We use [K] to denote the set of integers from 0 to K - 1. We write $||x||_{\Sigma} = \sqrt{x^{\top}\Sigma x}$ as a semi-norm of x when Σ is some positive-semidefinite matrix. We write $\Sigma \succeq \Sigma'$ if $\Sigma - \Sigma'$ is positive semidefinite.

Markov Decision Processes. We consider a finite-horizon MDP described by a tuple $M = (S, \mathcal{A}, H, \{P_h\}_{h=1}^H, \{R_h\}_{h=1}^H, \rho)$, where S is a (possibly infinite) state space, \mathcal{A} is a (possibly infinite) action space, H is the horizon length, $P_h : S \times \mathcal{A} \mapsto \Delta(S)$ is a probability transition matrix at step $h, R_h : S \times \mathcal{A} \mapsto \Delta([0, 1])$ encodes a family of reward distributions with $r_h : S \times \mathcal{A} \mapsto [0, 1]$ as the expected reward function, $\rho : S \mapsto \Delta(S)$ is the initial state distribution. At step h, upon executing action a from state s, the agent receives a deterministic reward $r_h(s, a)$ and transits to the next state s' with probability $P_h(s'|s, a)$. The MDP transits to an absorbing termination state with zero reward at step H. When H = 1 and there is no transition, the model reduces to the contextual bandit problem.

A deterministic policy $\pi_h : \mathcal{S} \to \mathcal{A}$ is a function that maps a state to an action at step $h \in [H]$. We use π to denote the family of policies $\{\pi_h\}_{h=1}^H$. Correspondingly, the value function $V^{\pi} : \mathcal{S} \to \mathbb{R}$ of the policy family $\{\pi_h\}_{h \in [H]}$ is defined as the expected sum of rewards starting at state s and following policy π_h at step h. More precisely, we have for any $s \in \mathcal{S}$, $V^{\pi}(s) := \mathbb{E}\left[\sum_{h=0}^{H} r_h(s_h, a_h) \mid s_0 = s, a_h = \pi_h(s_h), \forall h \ge 0\right]$, where the expectation is taken

over the trajectory generated according to the transition kernel $s_{h+1} \sim P_h(\cdot | s_h, a_h)$ and reward distribution $r_h \sim R_h(\cdot | s_h, a_h)$. The Q-function $Q^{\pi} : \mathcal{S} \times \mathcal{A} \to \mathbb{R}$ of policy π is defined analogously: $Q^{\pi}(s, a) \coloneqq \mathbb{E}\left[\sum_{h=0}^{H} r_h(s_h, a_h) | s_0 = s, a_0 = a, a_h = \pi_h(s_h), \forall h \ge 0\right]$. Note that although we work with undiscounted episodic case, it is straightforward to extend the framework and analysis to discounted MDP. We define the expected value of a policy π :

$$J(\pi) \coloneqq \mathbb{E}_{s \sim \rho}[V^{\pi}(s)] = \sum_{s \in \mathcal{S}} \rho(s) V^{\pi}(s).$$

We use shorthands $V^* \coloneqq V^{\pi^*}$ and $Q^* \coloneqq Q^{\pi^*}$ to denote the optimal value function and the optimal Q-function. We define the sub-optimality of any policy π as

$$\mathsf{SubOpt}(\pi) \coloneqq J(\pi^*) - J(\hat{\pi}).$$

We use shorthands $V^* \coloneqq V^{\pi^*}$ and $Q^* \coloneqq Q^{\pi^*}$ to denote the optimal value function and the optimal Q-function. We define the sub-optimality of any policy π as

$$\mathsf{SubOpt}(\pi) \coloneqq J(\pi^{\star}) - J(\hat{\pi}).$$

We also define the state occupancy measures $d^{\pi} : S \mapsto [0, H]$ and state-action occupancy measures $d^{\pi} : S \times \mathcal{A} \mapsto [0, H]$ as $d^{\pi}(s) \coloneqq \sum_{h=0}^{H} \mathbb{P}_h(s_h = s \mid \pi), d^{\pi}(s, a) \coloneqq \sum_{h=0}^{H} \mathbb{P}_h(s_h = s; a_h = a \mid \pi)$, where we use $\mathbb{P}_h(s_h = s \mid \pi)$ to denote the probability of visiting state $s_h = s$ (and similarly $s_h = s, a_h = a$) at step h after executing policy π and starting from $s_0 \sim \rho(\cdot)$.

Throughout the chapter, we make the following assumption on the parameterization of the reward:

Assumption 5. The reward lies in the family of linear functions $r_{\theta}(s, a) = \theta^{\top} \phi(s, a)$ for some known $\phi(s, a)$ with $\max_{s,a} \|\phi(s, a)\|_2 \leq L$. Let θ^* be the true parameter. To ensure the identifiability of θ^* , we let $\theta^* \in \Theta_B$, where

$$\Theta_B = \{ \theta \in \mathbb{R}^d \mid \langle 1, \theta \rangle = 0, \|\theta\|_2 \le B \}.$$

Sampling Procedure and Comparison Model. As in [191], we assume that both the states and actions in the training set come from a pre-collected dataset. In a contextual bandit, for the *i*-th sample, a state (prompt) s^i is first sampled from some fixed distribution ρ . Given the state s^i , K actions $(a_0^i, a_1^i, \dots, a_{K-1}^i)$ are sampled from some joint distribution $\mathbb{P}(a_0, \dots, a_{K-1} \mid s^i)^2$. Let $\sigma^i : [K] \mapsto [K]$ be the output of the human labeller, which is a permutation function that denotes the ranking of the actions. Here $\sigma^i(0)$ represents the most preferred action. We use $a_0 > a_1$ to denote the event that the action a_0 is more preferred compared to a_1 . A common model on the distribution of σ under K-ary comparisons is

²Indeed, it is not necessary to only compare actions under the same state. Our results can be easily generalized to the case when the states for K queries are completely different.

a Plackett-Luce model [200, 173]. The Plackett-Luce model defines the probability of a state-action pair (s, a_i) being the largest among a given set $\{(s, a_i)\}_{i=0}^{K-1}$ as

$$\mathbb{P}(a_i > a_j, \forall j \neq i \mid s) = \frac{\exp(r_{\theta}(s, a_i))}{\sum_{j=0}^{K-1} \exp(r_{\theta}(s, a_j))}$$

Moreover, one can calculate the probability of observing the permutation σ as³

$$\mathbb{P}(\sigma \mid s, \{a_i\}_{i=0}^{K-1}) = \prod_{i=0}^{K-1} \frac{\exp(r_{\theta^{\star}}(s, a_{\sigma(i)}))}{\sum_{j=i}^{K-1} \exp(r_{\theta^{\star}}(s, a_{\sigma(j)}))}.$$

When K = 2, this reduces to the pairwise comparison considered in the BTL model, which is used in existing RLHF algorithms. In this case, the permutation σ can be reduced to a Bernoulli random variable, representing whether a_0 is preferred compared to a_1 . Concretely, for each queried state-actions pair (s, a_0, a_1) , we observe a sample y from a Bernoulli distribution with parameter $\frac{\exp(r_{\theta^{\star}}(s,a_1))}{\exp(r_{\theta^{\star}}(s,a_0)) + \exp(r_{\theta^{\star}}(s,a_1))}$; i.e., for any $l \in \{0, 1\}$,

$$\mathbb{P}(y = l \mid s, a_0, a_1) = \frac{\exp(r_{\theta^*}(s, a_l))}{\exp(r_{\theta^*}(s, a_0)) + \exp(r_{\theta^*}(s, a_1))}.$$

Organization Section 2.2 presents the problem of learning with pairwise comparisons under the contextual bandit framework, we provide upper and lower bounds for MLE and pessimistic MLE. We extend the result into K-wise comparisons in Section 2.3 and MDP in Section 2.4. We discuss the guarantee for IRL in Section 2.5. We present our experimental results on simulated dataset in Section 2.6. We also discuss the analysis for nonlinear rewards in Appendix A.1.

2.2 Learning from Pairwise Comparison

We begin with the problem of learning from pairwise comparisons under the BTL model.

Algorithms: MLE and Pessimistic MLE

We first bound the estimation error for MLE, the most common algorithm in learning to rank and RLHF [168, 282, 32, 55, 191]. For any query-observation dataset $\{(s^i, a_1^i, a_2^i, y^i)\}_{i=1}^n$,

³In practice, one may introduce an extra temperature parameter σ and replace all r_{θ^*} with r_{θ^*}/σ . Here we take $\sigma = 1$.

MLE aims at minimizing the negative log likelihood, defined as:

$$\begin{aligned} \theta_{\mathsf{MLE}} &\in \arg\min_{\theta\in\Theta_B} \ell_{\mathcal{D}}(\theta), \\ \ell_{\mathcal{D}}(\theta) &= -\sum_{i=1}^n \log\Big(\frac{1(y^i=1)\cdot\exp(r_{\theta}(s^i,a_1^i))}{\exp(r_{\theta}(s^i,a_0^i)) + \exp(r_{\theta}(s^i,a_1^i))} + \frac{1(y^i=0)\cdot\exp(r_{\theta}(s^i,a_0^i))}{\exp(r_{\theta}(s^i,a_0^i)) + \exp(r_{\theta}(s^i,a_1^i))}\Big) \\ &= -\sum_{i=1}^n \log\Big(1(y^i=1)\cdot\operatorname{sigmoid}(\langle\theta,\,\phi(s^i,a_1^i) - \phi(s^i,a_0^i)\rangle) \\ &+ 1(y^i=0)\cdot\operatorname{sigmoid}(\langle\theta,\,\phi(s^i,a_0^i) - \phi(s^i,a_1^i)\rangle)\Big). \end{aligned}$$

When the minimizer is not unique, we take any of the $\hat{\theta}$ that achieve the minimum. Let $\mathcal{D} = \{(s^i, a_1^i, a_2^i)\}_{i=1}^n$ denote the queried state-action pairs. In this chapter, we study how one can utilize \mathcal{D} to learn a near-optimal reward model and policy. We first present a lemma on the estimation error conditioned on the data \mathcal{D} . The lemma is a generalization of the upper bound in Theorem 1 of [233] and the analysis follows a similar structure. The main difference is that [233] focus on the tabular case when $\phi(s, a)$ is always a standard basis vector, while in our case $\phi(s, a)$ can be an arbitrary *d*-dimensional vector. This confidence bound guarantee is also similar to the guarantee for dueling bandits and RL in [74, 192], except for that we have better rate in logarithmic factors since union bound is not needed in our case.

Lemma 6. For any $\lambda > 0$, with probability at least $1 - \delta$,

$$\|\hat{\theta}_{\mathsf{MLE}} - \theta^{\star}\|_{\Sigma_{\mathcal{D}} + \lambda I} \leq C \cdot \sqrt{\frac{d + \log(1/\delta)}{\gamma^2 n}} + \lambda B^2.$$

Here $\Sigma_{\mathcal{D}} = \frac{1}{n} \sum_{i=1}^n (\phi(s^i, a^i_1) - \phi(s^i, a^i_0))(\phi(s^i, a^i_1) - \phi(s^i, a^i_0))^{\top}, \ \gamma = 1/(2 + \exp(-LB) + \exp(LB)).$

The proof is deferred to Appendix A.2. The optimality of the bound can be seen via a lower-bound argument akin to that in Theorem 1 of [233].

Now consider the set of parameters

$$\Theta(\hat{\theta}_{\mathsf{MLE}}, \lambda) = \Big\{ \theta \in \Theta_B \mid \|\hat{\theta}_{\mathsf{MLE}} - \theta\|_{\Sigma_{\mathcal{D}} + \lambda I} \le C \cdot \sqrt{\frac{d + \log(\frac{1}{\delta})}{\gamma^2 n}} + \lambda B^2 \Big\}.$$

Lemma 6 shows that with probability at least $1 - \delta$, one has $\theta^* \in \Theta(\hat{\theta}_{\mathsf{MLE}})$. We thus consider the pessimistic MLE in Algorithm 1, which takes the lower confidence bound (LCB) as the reward estimate. In the context of LLM, the features of meaningful prompts and responses usually lie on a low-dimensional manifold. The idea of pessimism is to assign larger reward for the responses that lie on the manifold, and penalize the rarely seen responses that do not lie on manifold. We have the following guarantee for pessimistic MLE:

Algorithm 1 Pessimistic MLE

Input: The current estimator $\hat{\theta}$, the data covariance $\Sigma_{\mathcal{D}}$, the regularization parameter λ , the bound on the semi-norm $f(n, d, \delta, \lambda)$, a reference vector $v \in \mathbb{R}^d$, state distribution q Construct the confidence set

$$\Theta(\hat{\theta},\lambda) = \Big\{ \theta \in \Theta_B \mid \|\hat{\theta} - \theta\|_{\Sigma_{\mathcal{D}} + \lambda I} \le f(n,d,\delta,\lambda) \Big\}.$$

Compute the pessimistic expected value function

$$\hat{J}(\pi) = \min_{\theta \in \Theta(\hat{\theta}, \lambda)} \mathbb{E}_{s \sim q} [\theta^{\top}(\phi(s, \pi(s)) - v)]$$

= $(\mathbb{E}_{s \sim q}[\phi(s, \pi(s))] - v)^{\top} \hat{\theta} - \|(\Sigma_{\mathcal{D}} + \lambda I)^{-\frac{1}{2}} (\mathbb{E}_{s \sim q}[\phi(s, \pi(s))] - v)\|_2 \cdot f(n, d, \delta, \lambda)$

Return: $\hat{\pi} = \arg \max_{\pi} \hat{J}(\pi).$

Theorem 7. Let $\hat{\pi}_{\mathsf{PE}}$ be the output of Algorithm 1 when taking $\hat{\theta} = \hat{\theta}_{\mathsf{MLE}}$, $f(n, d, \delta, \lambda) = C \cdot \sqrt{\frac{d + \log(1/\delta)}{\gamma^2 n} + \lambda B^2}$, $q = \rho$. For any $\lambda > 0$ and $v \in \mathbb{R}^d$, with probability at least $1 - \delta$,

$$\mathsf{SubOpt}(\hat{\pi}_{\mathsf{PE}}) \le C \cdot \sqrt{\frac{d + \log(1/\delta)}{\gamma^2 n} + \lambda B^2} \cdot \|(\Sigma_{\mathcal{D}} + \lambda I)^{-1/2} \mathbb{E}_{s \sim \rho}[(\phi(s, \pi^{\star}(s)) - v)]\|_2.$$

The proof is deferred to Appendix A.2. We make several remarks.

Remark 8 (The single concentratability coefficient assumption). When v = 0, the term $\|(\Sigma_{\mathcal{D}} + \lambda I)^{-1/2} \mathbb{E}_{s \sim \rho}[\phi(s, \pi^*(s))]\|_2$ is referred to as a "single concentratability coefficient", which is assumed to be bounded in most of the literature on offline learning [212, 159, 285, 301, 302]. A bounded concentratability coefficient can be understood as certifying good coverage of the target vector $\mathbb{E}_{s \sim \rho}[\phi(s, \pi^*(s))]$ from the dataset \mathcal{D} in the feature space. The performance guarantee also holds when we replace π^* with any reference policy π on both sides.

Remark 9 (The choice of λ). When $\Sigma_{\mathcal{D}}$ is invertible, or when any $\theta \in \Theta_B$ is orthogonal to the nullspace of $\Sigma_{\mathcal{D}}$, the above inequality holds for the case of $\lambda = 0$. In other cases, one may minimize λ on the right-hand side, or simply take $\lambda = (d + \log(1/\delta)/(B^2\gamma^2 n))$ to achieve a near-optimal rate up to a constant factor.

Remark 10 (**The choice of** v). Compared to the traditional pessimism principle [212, 159, 285, 301, 302], we subtract an extra reference vector v in all the feature vectors ϕ . Subtracting a constant vector in feature space will not change the induced policy, but may affect the concentratability coefficient $\|(\Sigma_{\mathcal{D}} + \lambda I)^{-1/2}(\mathbb{E}_{s \sim \rho}[\phi(s, \pi(s))] - v)\|_2$.

We briefly describe the reason for introducing v here. Consider the case where the differences between features lie in the same subspace, while the feature ϕ itself does not. As a concrete example, consider a single state s and two actions a_0, a_1 , we let $\phi(s, a_0) = (1, 1)$ and

 $\phi(s, a_1) = (1, 0)$. The data covariance is $(\phi(s^i, a_1^i) - \phi(s^i, a_0^i))(\phi(s^i, a_1^i) - \phi(s^i, a_0^i))^{\top} = [0, 0; 0, 1]$. Thus $\|(\Sigma_{\mathcal{D}} + \lambda I)^{-1/2}\phi(s, a_0)\|_2$ can be arbitrarily large as $\lambda \to 0$ when v = 0. On the other hand, when we take $v = \phi(s, a_1)$, one can verify that $\|(\Sigma_{\mathcal{D}} + \lambda I)^{-1/2}(\phi(s, a_0) - v)\|_2 \leq 1$.

The above example illustrates the importance of choosing an appropriate v. A good rule of thumb for choosing v is the most common feature vector ϕ that appears in the data, so that more features can be covered. This also affords additional design latitude for other pessimism algorithms.

Remark 11 (Implementation for neural network). When r_{θ} is a neural network, Algorithm 1 may not be directly implementable. As an alternative, there has been a number of heuristic approximations considered, including Conservative Q-Learning [145], Implicit Q-Learning [144] and Adversarially Trained Actor Critic [46]. Furthermore, one may also introduce pessimism in the policy training procedure. For example, [191] add regularization terms in policy training, which enforces that the policy stays close to the original policy, and within the coverage of the pre-trained dataset. Our analysis supplies a theoretical rationale for such regularization terms.

Remark 12 (Implications for online learning). Although we mainly focus on offline learning, Lemma 6 also gives a straightforward online learning algorithm when combined with an optimism-based algorithm. In particular, a pure exploration-based active learning scheme would seek to compare pairs of actions whose feature difference is poorly covered by the past observations; i.e., find (s, a_1, a_2) such that $\|\phi(s, a_1) - \phi(s, a_2)\|_{(\Sigma_{\mathcal{D}} + \lambda I)^{-1}}$ is maximized. As a corollary of Lemma 6 and exploration results for linear bandits [1, 246], one can derive tight regret bound for online learning.

Remark 13 (Special Case: Multi-Armed Bandit). For multi-armed bandits we have only a single state, such that the feature $\phi(s, a)$ reduces to $\vec{1}_a$, which is a unit vector with 1 on its *a*-th element. In this case, the data covariance reduces to a Laplacian matrix, defined as $\Sigma_{\mathcal{D}} = \frac{1}{n} \sum_{i=1}^{n} (\vec{1}_{a_1} - \vec{1}_{a_0}) (\vec{1}_{a_1} - \vec{1}_{a_0})^{\top}$. This is precisely the problem considered in [233]. The Laplacian matrix is positive semidefinite and always has a zero eigenvalue, corresponding to an all ones eigenvector. When the graph induced by the Laplacian matrix is connected, any θ with $\langle 1, \theta \rangle = 0$ is orthogonal to the nullspace of $\Sigma_{\mathcal{D}}$, thus the theorem holds for the case of $\lambda = 0$.

Failure of MLE and Lower Bounds

We also show that there exists a simple linear bandit where MLE fails and pessimistic MLE succeeds. Let $\hat{\pi}_{\mathsf{MLE}} = \arg \max_{\pi} \mathbb{E}[r_{\hat{\theta}_{\mathsf{MLE}}}(s, \pi(s))]$ be the greedy policy with respect to the MLE.

Theorem 14. There exists a linear bandit with four actions and a sampling distribution such that for any n > 1,

$$\mathbb{E}[\mathsf{SubOpt}(\hat{\pi}_{\mathsf{MLE}})] \geq 0.1.$$

On the other hand, with probability at least $1 - \delta$,

$$\mathsf{SubOpt}(\hat{\pi}_{\mathsf{PE}}) \leq \frac{C \cdot \log(1/\delta)}{\sqrt{n}}$$

Here C is some universal constant.

The proof is deferred to Appendix A.2. The results show a separation between MLE and pessimistic MLE when the concentratability coefficient is bounded. The failure of MLE has also been empirically observed in [85], which leads to overoptimization with the trained reward model.

We also show that for the problems with bounded concentratability coefficient, pessimistic MLE is minimax-rate optimal up to a constant factor. Consider the family of contextual bandit instances as follows:

$$\mathsf{CB}(\Lambda) = \{\rho, \{(s_i, a_{i1}, a_{i2})\}_{i=1}^n, \theta^* \mid \|\Sigma_{\mathcal{D}}^{-1/2} \mathbb{E}_{s \sim \rho}[\phi(s, \pi^*(s))]\|_2 \le \Lambda\}.$$

Here we assume that $\Sigma_{\mathcal{D}}$ is invertible to simplify the presentation of the lower bound. For any $\mathcal{Q} \in \mathsf{CB}(\Lambda)$, we let $\mathsf{SubOpt}_{\mathcal{Q}}(\pi)$ be the sub-optimality under instance \mathcal{Q} . We have the following lower bound result, the proof of which is deferred to Appendix A.2.

Theorem 15. For any $d > 6, n \ge Cd\Lambda^2, \Lambda \ge 2$, there exists a feature mapping ϕ such that the following lower bound holds.

$$\inf_{\hat{\pi}} \sup_{\mathcal{Q} \in \mathsf{CB}(\Lambda)} \mathsf{SubOpt}_{\mathcal{Q}}(\hat{\pi}) \ge C\Lambda \cdot \sqrt{\frac{d}{n}}.$$

Comparing with the upper bound in Theorem 7, we see that the pessimistic MLE is minimax-optimal up to constant factors for the sub-optimality of induced policy.

2.3 Learning from *K*-wise comparisons

We now consider learning from K-wise comparisons under the PL model. In this case, we design two different estimators based on MLE. One involves directly maximizing the likelihood under the PL model, denoted as MLE_K . The other involves splitting the K-wise comparison data with pairwise comparisons and running MLE for pairwise comparisons. We denote this estimator as MLE_2 .

Algorithms

Guarantee for MLE_K. Let $\mathcal{D} = \{(s^i, a_0^i, \cdots, a_K^i)\}_{i=1}^n$ be the set of queried states and actions, and the permutation function σ^i be the output of the *i*-th query. We can compute

its maximum likelihood estimator as

$$\theta_{\mathsf{MLE}_{K}} \in \underset{\theta \in \Theta_{B}}{\arg \min} \ell_{\mathcal{D}}(\theta),$$

where $\ell_{\mathcal{D}}(\theta) = -\frac{1}{n} \sum_{i=1}^{n} \sum_{j=0}^{K-1} \log \left(\frac{\exp(\langle \theta, \phi(s^{i}, a^{i}_{\sigma_{i}(j)}) \rangle)}{\sum_{k=j}^{K-1} \exp(\langle \theta, \phi(s^{i}, a^{i}_{\sigma_{i}(k)}) \rangle)} \right)$

Similar to [233], we restrict our attention to $K = \mathcal{O}(1)$ since it is known that it is difficult for human to compare more than a small number of items due to a limited information storage and processing capacity [181, 138, 240, 220]. For instance, [220] recommend eliciting preferences over no more than seven options. We have the following result for K-wise comparisons.

Theorem 16. Under the K-wise PL model, for any $\lambda > 0$, with probability at least $1 - \delta$,

$$\|\hat{\theta}_{\mathsf{MLE}_{K}} - \theta^{\star}\|_{\Sigma_{\mathcal{D}} + \lambda I} \le C \cdot \sqrt{\frac{K^{4}(d + \log(1/\delta))}{\gamma^{2}n}} + \lambda B^{2}$$

Here $\Sigma_{\mathcal{D}} = \frac{2}{K(K-1)n} \left(\sum_{i=1}^{n} \sum_{j=0}^{K-1} \sum_{k=j+1}^{K-1} (\phi(s^{i}, a^{i}_{j}) - \phi(s^{i}, a^{i}_{k}))(\phi(s^{i}, a^{i}_{j}) - \phi(s^{i}, a^{i}_{k}))^{\top} \right)$, and $\gamma = \exp(-4LB)$. As a consequence, let $\hat{\pi}_{\mathsf{PE}_{K}}$ be the output of Algorithm 1 when taking $\hat{\theta} = \hat{\theta}_{\mathsf{MLE}_{K}}$, $f(n, d, \delta, \lambda) = C \cdot \sqrt{\frac{K^{4}(d + \log(1/\delta))}{\gamma^{2}n} + \lambda B^{2}}$. For any $\lambda > 0$ and $v \in \mathbb{R}^{d}$, with probability at least $1 - \delta$,

$$\mathsf{SubOpt}(\hat{\pi}_{\mathsf{PE}_K}) \le C \cdot \sqrt{\frac{K^4(d + \log(1/\delta))}{\gamma^2 n} + \lambda B^2} \cdot \|(\Sigma_{\mathcal{D}} + \lambda I)^{-1/2} \mathbb{E}_{s \sim \rho}[(\phi(s, \pi^{\star}(s)) - v)]\|_2.$$

The proof of Theorem 16 is provided in Appendix A.2. [233] also study the extension from pairwise to K-wise comparisons. However, they focus on the setting where only the maximum is selected, where we assume a complete ranking among K items is given. Also, they only provide an expectation bound while we provide a high-probability bound.

Compared to the pairwise comparison result in Theorem 7, the covariance matrix $\Sigma_{\mathcal{D}}$ now takes the sum over the feature differences between all pairs of actions among K-wise comparisons. As a cost, the right-hand side bound also introduces extra dependence on K. Our bound is likely to be loose in terms of the dependence on K. However, since we mainly focus on the case of $K = \mathcal{O}(1)$, such a bound is still near-optimal due to the minimax lower bound for pairwise comparisons. Furthermore, the gap between MLE and pessimistic MLE for sub-optimality still exists since Theorem 14 holds as a special case of K-wise comparison.

Guarantee for MLE₂ Besides the standard MLE approach, another option is to replace the joint distribution of K-ranking data with K(K-1)/2 pairs of pairwise comparisons. This

can be understood as replacing the true probability in MLE_K with the product of marginals:

$$\hat{\theta}_{\mathsf{MLE}_2} \in \underset{\theta \in \Theta_B}{\arg \min} \ell_{\mathcal{D}}(\theta),$$

where $\ell_{\mathcal{D}}(\theta) = -\frac{1}{n} \sum_{i=1}^{n} \sum_{j=0}^{K-1} \sum_{k=j+1}^{K-1} \log \left(\frac{\exp(\langle \theta, \phi(s^i, a^i_{\sigma_i(j)}) \rangle)}{\exp(\langle \theta, \phi(s^i, a^i_{\sigma_i(j)}) \rangle) + \exp(\langle \theta, \phi(s^i, a^i_{\sigma_i(k)}) \rangle)} \right).$

This estimator is also applied in the current RLHF for LLM (see e.g. [191]). We show that it also leads to a good induced policy, as is shown in the theorem below.

Theorem 17. Under the K-wise PL model, for any $\lambda > 0$, with probability at least $1 - \delta$,

$$\|\hat{\theta}_{\mathsf{MLE}_2} - \theta^\star\|_{\Sigma_{\mathcal{D}} + \lambda I} \le C \cdot \sqrt{\frac{d + \log(1/\delta)}{\gamma^2 n} + \lambda B^2}.$$

Here $\Sigma_{\mathcal{D}} = \frac{2}{K(K-1)n} \left(\sum_{i=1}^{n} \sum_{j=0}^{K-1} \sum_{k=j+1}^{K-1} (\phi(s^{i}, a^{i}_{j}) - \phi(s^{i}, a^{i}_{k}))(\phi(s^{i}, a^{i}_{j}) - \phi(s^{i}, a^{i}_{k}))^{\top} \right)$, and $\gamma = 1/(2 + \exp(-2LB) + \exp(2LB))$. As a consequence, let $\hat{\pi}_{\mathsf{PE}_{2}}$ be the output of Algorithm 1 when taking $\hat{\theta} = \hat{\theta}_{\mathsf{MLE}_{2}}$, $f(n, d, \delta, \lambda) = C \cdot \sqrt{\frac{d + \log(1/\delta)}{\gamma^{2}n} + \lambda B^{2}}$, $q = \rho$. For any $\lambda > 0$ and $v \in \mathbb{R}^{d}$, with probability at least $1 - \delta$,

$$\mathsf{SubOpt}(\hat{\pi}_{\mathsf{PE}_2}) \le C \cdot \sqrt{\frac{d + \log(1/\delta)}{\gamma^2 n} + \lambda B^2} \cdot \|(\Sigma_{\mathcal{D}} + \lambda I)^{-1/2} \mathbb{E}_{s \sim \rho}[(\phi(s, \pi^{\star}(s)) - v)]\|_2.$$

The proof of Theorem 17 is provided in Appendix A.2. Our theoretical analysis validates the empirical performance of MLE_2 in [191]. Compared to the guarantee for MLE_K , MLE_2 seems to has better nonasymptotic upper bound in terms of the dependence on K. However, it is likely that this comes from a loose analysis of MLE_K . The MLE_2 belongs to the family of the M-estimators, whose asymptotic variance is known to be larger than that of MLE [90, 154]. Thus, asymptotically, MLE_K is more efficient than MLE_2 . We can calculate the asymptotic variance of both estimators as follows:

Theorem 18. We have

$$\frac{\sqrt{n}(\hat{\theta}_{\mathsf{MLE}_{K}} - \theta^{\star}) \to \mathcal{N}(0, \mathcal{I}(\theta^{\star})^{-1})}{\sqrt{n}(\hat{\theta}_{\mathsf{MLE}_{2}} - \theta^{\star}) \to \mathcal{N}(0, V)}.$$

where

$$\begin{aligned} \mathcal{I}(\theta^{\star}) &= \mathbb{E}_{\theta^{\star}} \left[\sum_{j=0}^{K-1} \sum_{k=j}^{K-1} \sum_{k'=j}^{K-1} \frac{\exp(\langle \theta^{\star}, \phi(s^{i}, a^{i}_{\sigma_{i}(k)}) + \phi(s^{i}, a^{i}_{\sigma_{i}(k')}) \rangle)}{(\sum_{k'=j}^{K-1} \exp(\langle \theta^{\star}, \phi(s^{i}, a^{i}_{\sigma_{i}(k')}) \rangle))^{2}} \\ &\cdot (\phi(s^{i}, a^{i}_{\sigma_{i}(k)}) - \phi(s^{i}, a^{i}_{\sigma_{i}(k')}))(\phi(s^{i}, a^{i}_{\sigma_{i}(k)}) - \phi(s^{i}, a^{i}_{\sigma_{i}(k')}))^{\top} \right], \\ V &= \Sigma^{-1} \mathbb{E}_{\theta^{\star}} \left[GG^{\top} \right] \Sigma^{-1}, \\ \Sigma &= \mathbb{E}_{\theta^{\star}} \left[\sum_{k=1}^{K-1} \sum_{j=1}^{K-1} \frac{\exp(-\langle \theta^{\star}, x^{i}_{\sigma_{i}(j)\sigma_{i}(k)} \rangle)}{(1 + \exp(-\langle \theta^{\star}, x^{i}_{\sigma_{i}(j)\sigma_{i}(k)} \rangle))^{2}} \cdot \left(\phi(s^{i}, a^{i}_{\sigma_{i}(j)}) - \phi(s^{i}, a^{i}_{\sigma_{i}(k)}) \right) \end{aligned}$$

$$G = \sum_{j=0}^{K-1} \sum_{k=j+1}^{K-1} \frac{\exp(-\langle \theta^*, x_{\sigma_i(j)\sigma_i(k)}^i) \rangle^\top}{1 + \exp(-\langle \theta^*, x_{\sigma_i(j)\sigma_i(k)}^i) \rangle} \cdot \left(\phi(s^i, a_{\sigma_i(j)}^i) - \phi(s^i, a_{\sigma_i(k)}^i)\right)$$

The proof follows directly the gradient and Hessian computed in Appendix A.2 and A.2, combined with Section 5.3 of [265]. We also empirically verify the performances of both estimators in Section 2.6.

2.4 Extension to MDPs

Thus far we have considered only contextual bandits. We now extend our results to the MDP setting. Depending on whether the comparison is based on a single action or a whole trajectory, we have two regimes, namely action-based comparison and trajectory-based comparison.

Trajectory-based Comparison

In trajectory-based comparison, we assume that two trajectories that start from the same initial state are given, and the comparison is based on the cumulative reward of the two trajectories. Concretely, we first sample the initial state s_0 from some fixed distribution ρ , and then sample two trajectories $\tau_0 = (a_0, s_1, a_1, \dots, s_H, a_H)$ and $\tau_1 = (a'_0, s'_1, a'_1, \dots, s'_H, a'_H)$ from joint distributions $P_l(a_0, s_1, a_1, \dots, s_H, a_H|s_0) = \prod_i \pi_l(a_i|s_i)P(s_{i+1}|s_i, a_i)$, where $l \in \{0, 1\}$. For each queried state-trajectory pair, we observe a sample y from a Bernoulli distribution as follows:

$$\mathbb{P}(y=1 \mid s, \tau_0, \tau_1) = \frac{\exp(\sum_{h=0}^{H} r_{\theta^{\star}}(s_h, a_h))}{\exp(\sum_{h=0}^{H} r_{\theta^{\star}}(s_h, a_h)) + \exp(\sum_{h=0}^{H} r_{\theta^{\star}}(s'_h, a'_h)))}$$

Given the dataset $\{(s^i,\tau^i_0,\tau^i_1,y^i\}_{i=1}^n,$ the MLE is

$$\begin{aligned} \theta_{\mathsf{MLE}} &\in \argmin_{\theta \in \Theta_B} \ell_{\mathcal{D}}(\theta), \\ \text{where } \ell_{\mathcal{D}}(\theta) &= -\sum_{i=1}^{n} \log \Big(\frac{1(y^{i}=1) \cdot \exp(\sum_{h=0}^{H} r_{\theta}(s_{h}^{i}, a_{h}^{i}))}{\exp(\sum_{h=0}^{H} r_{\theta}(s_{h}^{i}, a_{h}^{i})) + \exp(\sum_{h=0}^{H} r_{\theta}(s_{h}^{i\prime}, a_{h}^{i\prime}))} \\ &+ \frac{1(y^{i}=0) \cdot \exp(\sum_{h=0}^{H} r_{\theta}(s_{h}^{i\prime}, a_{h}^{i\prime}))}{\exp(\sum_{h=0}^{H} r_{\theta}(s_{h}^{i}, a_{h}^{i})) + \exp(\sum_{h=0}^{H} r_{\theta}(s_{h}^{i\prime}, a_{h}^{i\prime}))} \Big). \end{aligned}$$

Compared to the pairwise comparison in the contextual bandit, the exponent changes from a single reward to the cumulative reward. Similarly, we provide the following guarantee for the estimation error of MLE:

Lemma 19. Assume that $\|\phi(\cdot, \cdot)\|_{\infty} \leq L$ for any s, a. Then for any $\lambda > 0$, with probability at least $1 - \delta$,

$$\|\hat{\theta}_{\mathsf{MLE}} - \theta^{\star}\|_{\Sigma_{\mathcal{D}} + \lambda I} \le C \cdot \sqrt{\frac{d \log(1/\delta)}{\gamma^2 n} + \lambda B^2}.$$

Here $\Sigma_{\mathcal{D}} = \frac{1}{n} \sum_{i=1}^{n} (\sum_{h=0}^{H} (\phi(s_h^i, a_h^i) - \phi(s_h^{i\prime}, a_h^{i\prime}))) (\sum_{h=0}^{H} (\phi(s_h^i, a_h^i) - \phi(s_h^{i\prime}, a_h^{i\prime})))^{\top}$, and $\gamma = 1/(2 + \exp(-2HLB) + \exp(2HLB))$.

The proof is deferred to Appendix A.2. Compared to the guarantee for contextual bandits in Lemma 6, the features in the covariance is now the difference between the cumulative feature in trajectory τ and the cumulative feature in trajectory τ' . The result reduces to Lemma 6 when H = 1.

In order to bound the sub-optimality of the induced policy, one needs to plug-in a pessimistic version of the reward estimate. Note that from the definition of d^{π} , one has

$$\mathbb{E}_{s \sim \rho}[V^{\pi}(s)] = \mathbb{E}_{s, a \sim d^{\pi}}[r(s, a)]$$

In the case when the transition distribution P is known, one may directly compute d^{π} for any policy π and replace the initial distribution ρ in the algorithm for contextual bandit. This gives the following result:

Theorem 20. Let $\hat{\pi}_{\mathsf{PE}}$ be the output of Algorithm 1 when taking $\hat{\theta} = \hat{\theta}_{\mathsf{MLE}}$, $f(n, d, \delta, \lambda) = C \cdot \sqrt{\frac{d + \log(1/\delta)}{\gamma^2 n} + \lambda B^2}$, $q = d^{\pi}$. For any $\lambda > 0$ and $v \in \mathbb{R}^d$, with probability at least $1 - \delta$,

$$\begin{aligned} \mathsf{SubOpt}(\hat{\pi}_{\mathsf{PE}}) &\leq C \cdot \sqrt{\frac{d + \log(1/\delta)}{\gamma^2 n} + \lambda B^2} \\ & \cdot \| (\Sigma_{\mathcal{D}} + \lambda I)^{-1/2} \mathbb{E}_{s \sim d^{\pi^\star}} [(\phi(s, \pi^\star(s)) - v)] \|_2. \end{aligned}$$

The proof is deferred to Appendix A.2. The result can be generalized to the case of K-wise comparisons following the same argument in Section 2.3.

Action-based Comparison

 $\phi(s^i, a^i_0))^{\top}$.

In action-based comparison, we assume that two actions are sampled for each state, and the comparison is based on the expected cumulative return starting from such state-action pair.

Concretely, assume that the optimal Q-function is parameterized as $Q_{\theta}^{\star}(s, a) = \theta^{\top} \phi(s, a)$ for some given $\phi(s, a)$. Let θ^{\star} be the true parameter. During the training, we first sample the state s from some fixed distribution ρ , and then sample a pair of actions a_0, a_1 from a joint distribution $P(a_0, a_1|s)$. For each queried state-actions pair (s, a_0, a_1) , we observe a sample yfrom a Bernoulli distribution with parameter $\frac{\exp(Q_{\theta^{\star}}(s,a_1))}{\exp(Q_{\theta^{\star}}(s,a_0)) + \exp(Q_{\theta^{\star}}(s,a_1))}$, i.e.

$$\mathbb{P}(y=1 \mid s, a_0, a_1) = \frac{\exp(Q_{\theta^{\star}}(s, a_1))}{\exp(Q_{\theta^{\star}}(s, a_0)) + \exp(Q_{\theta^{\star}}(s, a_1))},$$

and $\mathbb{P}(y=0 \mid s, a_0, a_1) = \frac{\exp(Q_{\theta^{\star}}(s, a_0))}{\exp(Q_{\theta^{\star}}(s, a_0)) + \exp(Q_{\theta^{\star}}(s, a_1))}.$

In this case, one may use the same MLE to estimate θ^* , which results in an estimator \hat{Q} for the Q^* -function. The following lemma follows exactly the same analysis as Lemma 6:

Lemma 21. Under the BTL model for action-based RLHF, for any $\lambda > 0$, with probability at least $1 - \delta$,

$$\|\hat{\theta}_{\mathsf{MLE}} - \theta^{\star}\|_{\Sigma_{\mathcal{D}} + \lambda I} \leq C \cdot \sqrt{\frac{d + \log(1/\delta)}{\gamma^2 n} + \lambda B^2}.$$

Here $\gamma = 1/(2 + \exp(-LB) + \exp(LB)).$ $\Sigma_{\mathcal{D}} = \frac{1}{n} \sum_{i=1}^n (\phi(s^i, a^i_1) - \phi(s^i, a^i_0))(\phi(s^i, a^i_1))$

When $\Sigma_{\mathcal{D}}$ is invertible and covers all the directions well, this will lead to a valid confidence bound for Q^* , which implies a good performance of the induced greedy policy without pessimism. However, when $\Sigma_{\mathcal{D}}$ does not provide good coverage, introducing pessimism in this case can be hard. The reason is that one needs to construct lower confidence bound for Q^{π} for any π . However, given such confidence bound of $\hat{\theta}_{\mathsf{MLE}}$, one can only construct confidence bound for Q^* .

2.5 Connection with Inverse Reinforcement Learning

In Inverse Reinforcement Learning (IRL), BTL and PL model are also popular model of human behavior. However, in IRL it is assumed that we only observe the human behavior, which is sampled from the distribution under PL model. Thus no comparison is queried. Depending on the comparison is action-based on trajectory-based, one has max-entropy IRL or action-based IRL, discussed in details below.

Trajectory-based IRL

In max-entropy IRL [322], it is also assumed that the human selection of trajectory follows a PL model. A common assumption in IRL or IL is that the observed trajectory collected by human behavior is likely to be the optimal policy. Assume that the transitions are deterministic. For any trajectory $\tau = (s_0, a_0, \dots, s_H, a_H)$, it is assumed that the expert chooses trajectory τ under the following model:

$$\mathbb{P}(\tau) = \frac{\exp(\sum_{h=0}^{H} \langle \theta^{\star}, \phi(s_h, a_h) \rangle)}{\sum_{\tau' \in \mathcal{T}(s_0)} \exp(\sum_{h=0}^{H} \langle \theta^{\star}, \phi(s'_h, a'_h) \rangle)}.$$

Here the set $\mathcal{T}(s_0)$ denotes the set for all possible trajectories that start from s_0 . Each trajectory is represented by $\tau' = \{(s'_h, a'_h)\}_{h=1}^H$. Assume that we are given a set of trajectories $\{s_h^i, a_h^i\}_{i \in [n], h \in [H]}$ that are sampled from the distribution $\mathbb{P}(\tau)$. When the denominator can be computed exactly, the algorithm of max entropy IRL also reduces to the MLE, which can be written as

$$\hat{\theta}_{\mathsf{MLE}} \in \underset{\theta \in \Theta_B}{\arg\min} \ell_{\mathcal{D}}(\theta),$$

where $\ell_{\mathcal{D}}(\theta) = -\frac{1}{n} \sum_{i=1}^{n} \log \left(\frac{\exp(\sum_{h=0}^{H} \langle \theta, \phi(s_h^i, a_h^i) \rangle)}{\sum_{\tau' \in \mathcal{T}(s_0^i)} \exp(\sum_{h=0}^{H} \langle \theta, \phi(s_h', a_h') \rangle)} \right)$

Although the enumeration of all trajectories $\mathcal{T}(s_0^i)$ is not possible due to exponential growth of the possible trajectories with respect to horizon H, [322] provides an alternative way of computing the gradient via calculating the expected state frequency. This enables the efficient implementation of MLE. One can show the performance guarantee for max entropy IRL as follows:

Lemma 22. Under the PL model, for any $\lambda > 0$, with probability at least $1 - \delta$,

$$\|\hat{\theta}_{\mathsf{MLE}} - \theta^{\star}\|_{\Sigma_{\mathcal{D}} + \lambda I} \le C \cdot \sqrt{\frac{\sup_{s} |\mathcal{T}(s)|^{2} \cdot (d + \log(1/\delta))}{\gamma^{2}n} + \lambda B^{2}}.$$

Here $\gamma = \exp(-4LB)/2m$, and $\Sigma_{\mathcal{D}}$ is defined as

$$\frac{1}{n \sup_{s} |\mathcal{T}(s)|^{2}} \sum_{i=1}^{n} \sum_{\{(s_{h}, a_{h})\} \in \mathcal{T}(s_{0}^{i})} \sum_{\{(s_{h}^{'}, a_{h}^{'})\} \in \mathcal{T}(s_{0}^{i})} (\sum_{h=0}^{H} (\phi(s_{h}, a_{h}) - \phi(s_{h}^{'}, a_{h}^{'}))) (\sum_{h=0}^{H} (\phi(s_{h}, a_{h}) - \phi(s_{h}^{'}, a_{h}^{'})))^{\top}$$

Given such guarantee for MLE, we also show that IRL, when combined with pessimism principle, will lead to a good policy.
Theorem 23. Let $\hat{\pi}_{\mathsf{PE}}$ be the output of Algorithm 1 when taking $\hat{\theta} = \hat{\theta}_{\mathsf{MLE}}$, $f(n, d, \delta, \lambda) = C \cdot \sqrt{\frac{\sup_s |\mathcal{T}(s)|(d+\log(1/\delta))}{\gamma^2 n} + \lambda B^2}$, $q = d^{\pi}$. For any $\lambda > 0$ and $v \in \mathbb{R}^d$, with probability at least $1 - \delta$,

$$\mathsf{SubOpt}(\hat{\pi}_{\mathsf{PE}}) \leq C \cdot \sqrt{\frac{\sup_{s} |\mathcal{T}(s)|^2 (d + \log(1/\delta))}{\gamma^2 n}} + \lambda B^2$$
$$\cdot \| (\Sigma_{\mathcal{D}} + \lambda I)^{-1/2} \mathbb{E}_{s \sim \rho} [(\phi(s, \pi^*(s)) - v)] \|_2.$$

The proof of Lemma 22 and Theorem 23 is provided in Appendix A.2. For IRL we have the dependence of $\sup_s |\mathcal{T}(s)|$ in our bound, which can be much larger than d. Similar to the case of K-wise comparison, one may also split the one observation into $\sup_s |\mathcal{T}(s)|$ pairwise comparisons, which can help improve the dependence on $\sup_s |\mathcal{T}(s)|$ in the current analysis.

Action-based IRL

Similar to action-based RLHF, action-based IRL also models human choice based on Q^* instead of cumulative reward [210, 186, 78]. Concretely, the human behavior is assumed to be based on the Q function $Q^*(s, a) = \langle \theta^*, \phi(s, a) \rangle$, i.e.

$$\pi^{\star}(a|s) = \frac{\exp(\langle \theta^{\star}, \phi(s, a) \rangle)}{\sum_{a' \in \mathcal{A}} \exp(\langle \theta^{\star}, \phi(s, a') \rangle)}$$

Here the denominator takes all possible actions. Unlike RLHF where a pair of actions are observed, in IRL or IL, only a single human behavior is observed in each round and there is no comparison, i.e. the observed actions a are sampled from $\pi^*(a \mid s)$. Given such observation, one can still run MLE and gives similar performance guarantee. In particular, the MLE is given by

$$\hat{\theta}_{\mathsf{MLE}} \in \underset{\theta \in \Theta_B}{\arg \min} \ell_{\mathcal{D}}(\theta),$$

where $\ell_{\mathcal{D}}(\theta) = -\frac{1}{n} \sum_{i=1}^{n} \log \left(\frac{\exp(\langle \theta, \phi(s^i, a^i) \rangle)}{\sum_{a' \in \mathcal{A}} \exp(\langle \theta, \phi(s^i, a') \rangle)} \right)$

The following lemma follows a similar analysis as Lemma 6 and Lemma 22:

Lemma 24. Under the PL model for action-based IRL, for any $\lambda > 0$, with probability at least $1 - \delta$,

$$\|\hat{\theta}_{\mathsf{MLE}} - \theta^{\star}\|_{\Sigma_{\mathcal{D}} + \lambda I} \le C \cdot \sqrt{\frac{|\mathcal{A}|^2 (d + \log(1/\delta))}{\gamma^2 n}} + \lambda B^2.$$

Here $\Sigma_{\mathcal{D}} = \frac{1}{n|\mathcal{A}|^2} \sum_{i=1}^n \sum_{a \in \mathcal{A}} \sum_{a' \in \mathcal{A}} (\phi(s^i, a) - \phi(s^i, a')) (\phi(s^i, a) - \phi(s^i, a')))^\top$, $\gamma = \exp(-4LB)/2$.

Similar to the case of action-based RLHF, it remains an interesting open problem how one can introduce provable lower confidence bound algorithm for policy learning.

2.6 Experiments

There has been a large amount of empirical work that demonstrates the success of MLE and pessimistic MLE in RLHF for game playing [140, 175, 55, 272], robotics [27, 241] and language models [323, 252, 279, 183, 191, 178, 89, 85, 18, 84, 211]. Notably, the concurrent work [241] proposes Offline Preference-Based Reward Learning (OPRL), which trains pessimistic policy from the learned reward and shows empirically the superior performance of pessimistic based method (which can be viewed as an approximation of pessimistic MLE).



Figure 2.1. Left: the convergence of MLE under the semi-norm $\|\cdot\|_{\Sigma}$; Right: the comparison between MLE and pessimistic MLE under sub-optimality metric.

In this section, we provide experiments for the contextual bandit case. In particular, we conduct both MLE and pessimistic MLE on the example constructed in Appendix A.2. The results are included in Fig. 2.1. We range the number of samples n from 10 to 500. Each sample size is repeated 100 times. The result verifies our theoretical analysis: MLE converges under the semi-norm but fails to give good policy. On the other hand, pessimistic MLE gives vanishing rate when considering the sub-optimality of the induced policy. Note that in the left figure we do not include pessimistic MLE, since both MLE and pessimistic MLE rely on the same parameter $\hat{\theta}_{MLE}$, and they only defer in how the induced policy is trained.

On the other hand, we compare the performance of MLE_2 and MLE_K when learning from K-wise comparisons. We take K = 4 and K = 9, and range samples from 10 to 500. We randomly generate ϕ and θ^* as independent samples from 3-dimensional Gaussian distribution. The result is shown in Figure 2.2. One can see that as n grows larger, both estimators converge, while MLE_K has smaller estimation error than MLE_2 . The gap grows larger when K becomes larger. This is consistent with our theoretical prediction in Section 2.3: since MLE_K is the true MLE and MLE_2 belongs to the family of M-estimators, asymptotically MLE_K shall be more efficient than MLE_2 .

2.7 Conclusion

We have provided a theoretical analysis of the sample complexity of RLHF. Our main results involve two insights: (i) pessimism is important to guarantee a good policy; (ii) in



Figure 2.2. The comparison of estimation error between MLE_2 and MLE_K , with K = 4 in the left and K = 9 in the right.

K-wise comparison, both MLE_K and MLE_2 converge. Moreover, MLE_K is asymptotically more efficient.

While we have made progress in understanding the reward learning aspect of RLHF, there are many additional questions that remain to be answered.

- 1. We assumed that the policy trained is greedy with respect to the learned reward. However, in practice the reward is mostly used to fine-tune the pre-trained policy. This requires a more extensive theory that considers the whole procedure of pre-training the policy, learning a reward model and then fine-tuning the policy with policy gradient or PPO.
- 2. Although we focused on the BTL and PL models, there have been a number of other models considered for the modeling of human behavior, including the Thurstone model and cardinal models. It would be interesting to extend our analysis to cover these additional models and begin to provide a general characterization of behavioral models for RLHF.
- 3. Our constructed confidence bound is based on a fixed feature ϕ . In the practical fine-tuning scenario, ϕ is not fixed but may change slowly. It is interesting to see how the constructed confidence bound helps in the practical fine-tuning scenario for online (active) learning or offline learning, and how one can design valid confidence bound for slowly changing ϕ .

Chapter 3

Practical Implementation of Pessimism in RLHF

3.1 Introduction

The initial phase of RLHF involves learning human values using a reward model from ranking data. It is observed that the performance of the reward model degrades after one epoch of training, and optimizing too much against the learned reward model eventually hinders the true objective. This chapter analyzes potential reasons behind the issues, and designs improved reward learning algorithm termed 'Iterative Data Smoothing' (IDS). The core idea is that during each training epoch, we not only update the model with the data, but also update the date using the model, replacing hard labels with soft labels. Our empirical findings highlight the superior performance of this approach over the traditional methods.

Following on from a supervised learning stage, a typical RLHF protocol involves two main steps:

- **Reward learning:** Sample prompts from a prompt dataset and generate multiple responses for the same prompt. Collect human preference data in the form of pairwise or multi-wise comparisons of different responses. Train a reward model based on the preference data.
- **Policy learning:** Fine-tune the current LLM based on the learned reward model with reinforcement learning algorithms.

Although RLHF has been successful in practice [18, 191, 69], it is not without flaws, and indeed the current reward training paradigm grapples with significant value-reward mismatches. There are two major issues with the current paradigm:

• **Reward overfitting:** During the training of the reward model, it has been observed that the test cross-entropy loss of the reward model can deteriorate after one epoch of training [191].

• **Reward overoptimization:** When training the policy model to maximize the reward predicted by the learned model, it has been observed that the ground-truth reward can increase when the policy is close in KL divergence to the initial policy, but decrease with continued training [86].

In this chapter, we investigate these issues in depth. We simplify the formulation of RLHF to a multi-armed bandit problem and make attempt to explain and reproduce the overfitting and overoptimization phenomena. We leverage theoretical insights in the bandit setting to design new algorithms that help mitigate the issues and work well under practical fine-tuning scenarios.

Main Results

As our first contribution, we make the attempt to explain and analyze both reward overfitting and overoptimization from the simple setting of multi-armed bandit. We show that the **inadequacy of the cross-entropy loss for long-tailed preference datasets** can be one of the reasons for both overfitting and overoptimization. As illustrated in Figure 3.1, even a simple 3-armed bandit problem can succumb to overfitting and overoptimization when faced with such imbalanced datasets. Consider a scenario where we have three arms with true rewards given by $r_1^* = 1, r_2^* = r_3^* = 0$, and the preference distribution is generated by the Bradley-Terry-Luce (BTL) model [25], i.e. $\mathbb{P}(i \succ j) = \exp(r_i^*)/(\exp(r_i^*) + \exp(r_j^*))$. Suppose our preference dataset compares the first and second arms 1000 times but only compares the first and third arm once, and let $n(i \succ j)$ denote the number of times that arm *i* is preferred over arm *j*. The standar empirical cross-entropy loss used in the literature for learning the reward model [191, 316] can be written as follows:

$$-\sum_{i,j} n(i \succ j) \log \left(\frac{\exp(r_i)}{\exp(r_i) + \exp(r_j)} \right).$$

We know that the empirical values $n(1 \succ 2)$ and $n(2 \succ 1)$ concentrate around their means. However, we have with probability 0.73, $n(1 \succ 3) = 1$ and $n(3 \succ 1) = 0$, and with probability 0.27, $n(1 \succ 3) = 0$ and $n(3 \succ 1) = 1$. In either case, the minimizer of the empirical entropy loss will satisfy either $\hat{r}_1 - \hat{r}_3 = -\infty$ or $\hat{r}_1 - \hat{r}_3 = +\infty$. This introduces a huge effective noise when the coverage is imbalanced. Moreover, the limiting preference distribution is very different from the ground truth distribution, leading to reward overfitting. Furthermore, since there is 0.27 probability that $\hat{r}_1 - \hat{r}_3 = -\infty$, we will take arm 3 as the optimal arm instead of arm 1. This causes reward overoptimization during the stage of policy learning since the final policy converges to the wrong arm with reward zero.

To mitigate these effects, we leverage the pessimism mechanism from bandit learning to analyze and design a new algorithm, Iterative Data Smoothing (IDS), that simultaneously addresses both reward overfitting and reward overoptimization. The algorithm design is straightforward: in each epoch, beyond updating the model with the data, we also adjust the



Figure 3.1. Illustration of the problem of the vanilla empirical cross-entropy minimization for learning the ground truth reward. With a small number of samples comparing arm 1 and 3, the minimization converges to a solution which assigns $\hat{r}_1 - \hat{r}_3 = -\infty$ with constant probability. With the proposed Iterative Data Smoothing (IDS) algorithm, the estimator is able to recover the ground truth reward.

data using the model. Theoretically, we investigate the two phenomena in the tabular bandit case. We show that the proposed method, as an alternative to the lower-confidence-boundbased algorithm [131, 284, 212, 316], learns the ground truth distribution for pairs that are compared enough times, and ignores infrequently covered comparisons thereby mitigating issues introduced by long-tailed data. Empirically, we present experimental evidence that the proposed method improves reward training in both bandit and neural network settings.

Related Work

RLHF and Preference-based Reinforcement Learning. RLHF, or Preference-based Reinforcement Learning (PbRL), has delivered significant empirical success in the fields of game playing, robot training, stock prediction, recommender systems, clinical trials and natural language processing [189, 221, 56, 148, 121, 277, 140, 175, 55, 272, 27, 241, 323, 252, 279, 183, 191, 178, 89, 85, 18, 84, 211]. In the setting of the language models, there has been work exploring the efficient fine-tuning of the policy model [244, 247, 296, 318, 207, 280].

In the case of reward learning, [191] notes that in general the reward can only be trained for one epoch in the RLHF pipeline, after which the test loss can go up. [86] studies the scaling law of training the reward model, and notes that overoptimization is another problem in reward learning. To address the problem, [316] propose a pessimism-based method that improves the policy trained from the reward model when the optimal reward lies in a linear family. It is observed in [248] that the reward model tends to be identical regardless of whether the prompts are open-ended or closed-ended during the terminal phase of training, and they propose a prompt-dependent reward optimization scheme.

Another closely related topic is the problem of estimation and ranking from pairwise or K-wise comparisons. In the literature of *dueling bandit*, one compares two actions and aims to minimize regret based on pairwise comparisons [299, 324, 298, 297, 225, 88, 223, 9, 325, 141, 83, 222, 224, 74]. [189, 289] analyze the sample complexity of dueling RL agents in the tabular case, which is extended to linear case and function approximation by the recent work of [192, 44]. [37] studies a related setting where in each episode only binary feedback is received. Most of the theoretical work of learning from ranking focuses on regret minimization, while RLHF focuses more on the quality of the final policy.

Knowledge Distillation The literature of knowledge distillation focuses on transferring the knowledge from a teacher model to a student model [106, 82, 50, 306, 215, 294, 113, 195, 261, 263, 203, 47]. It is observed in this literature that the soft labels produced by the teacher network can help train a better student network, even when the teacher and student network are of the same size and structure [106]. [82] present a method which iteratively trains a new student network after the teacher network achieves the smallest evaluation loss. Both our iterative data smoothing algorithm and these knowledge distillation methods learn from soft labels. However, iterative data smoothing iteratively updates the same model and data, while knowledge distillation method usually focuses on transferring knowledge from one model to the other.

3.2 Formulation

We begin with the notation that we use in the chapter. Then we introduce the general formulation of RLHF, along with our simplification in the multi-armed bandit case.

Notations. We use calligraphic letters for sets, e.g., S and A. Given a set S, we write |S| to represent the cardinality of S. We use [K] to denote the set of integers from 1 to K. We use $\mu(a, a')$ to denote the probability of comparing a and a' in a preference dataset, and $\mu(a) = \sum_{a' \in A} \mu(a, a')$ to denote the probability of comparing a with any other arms. Similarly, we use n(a), n(a, a') to denote the number of samples that compare a with any other arms, and the number of samples that compare a with a', respectively. We use $a_1 \succ a_2$ to denote the event that the a_1 is more preferred compared to a_2 .

General Formulation of RLHF

The key components in RLHF consist of two steps: reward learning and policy learning. We briefly introduce the general formulation of RLHF below.

In the stage of reward learning, one collects a preference dataset based on a prompt dataset and responses to the prompts. According to the formulation of [316], for the *i*-th sample, a state (prompt) s_i is first sampled from some prompt distribution ρ . Given the state s_i , M actions (responses) $(a_i^{(1)}, a_i^{(2)}, \dots, a_i^{(M)})$ are sampled from some joint distribution $\mathbb{P}(a^{(1)}, \dots, a^{(M)} | s_i)$, Let $\sigma_i : [M] \mapsto [M]$ be the output of the human labeller, which is a permutation function that denotes the ranking of the actions. Here $\sigma_i(1)$ represents the most preferred action, and $\sigma_i(M)$ is the least preferred action. A common model for the distribution of σ under multi-ary comparisons is a Plackett-Luce model [200, 173]. The Plackett-Luce model defines the probability of a state-action pair (s, a_i) being the largest among a given set $\{(s, a_i)\}_{i=1}^M$ as

$$\mathbb{P}(a_i \succ a_j, \forall j \neq i \mid s) = \frac{\exp(r^{\star}(s, a_i))}{\sum_{j=1}^{M} \exp(r^{\star}(s, a_j))},$$

where $r^* : \mathcal{S} \times \mathcal{A} \mapsto \mathbb{R}$ is the ground-truth reward for the response given the prompt. Moreover, the probability of observing the permutation σ is defined as¹

$$\mathbb{P}(\sigma \mid s, \{a^{(i)}\}_{i=1}^{M}) = \prod_{i=1}^{M} \frac{\exp(r^{\star}(s, a^{(\sigma(i))}))}{\sum_{j=i}^{M} \exp(r^{\star}(s, a^{(\sigma(j))}))}.$$

When M = 2, this reduces to the pairwise comparison considered in the Bradley-Terry-Luce (BTL) model [25], which is used in existing RLHF algorithms. In this case, the permutation σ can be reduced to a Bernoulli random variable, representing whether one action is preferred compared to the other. Concretely, for each queried state-actions pair (s, a, a'), we observe a sample c from a Bernoulli distribution with parameter $\frac{\exp(r^*(s,a))}{\exp(r^*(s,a))+\exp(r^*(s,a'))}$. Based on the observed dataset, the cross-entropy loss is minimized to estimate the ground-truth reward for the case of pairwise comparison. The minimizer of cross-entropy loss is the maximum likelihood estimator:

$$\hat{r}_{\mathsf{MLE}} \in \arg\min_{r} \mathcal{L}_{\mathsf{CE}}(\mathcal{D}, r),$$

$$\mathcal{L}_{\mathsf{CE}}(\mathcal{D}, r) = -\sum_{i=1}^{n} \log\left(\frac{y_{i} \cdot \exp(r(s_{i}, a_{i}^{(1)}))}{\exp(r(s_{i}, a_{i}^{(1)})) + \exp(r(s_{i}, a_{i}^{(2)}))} + \frac{(1 - y_{i}) \cdot \exp(r(s_{i}, a_{i}^{(2)}))}{\exp(r(s_{i}, a_{i}^{(1)})) + \exp(r(s_{i}, a_{i}^{(2)}))}\right).$$
(3.1)

After learning the reward, we aim to learn the optimal policy under KL regularization with respect to an initial policy π_0 under some state (prompt) distribution ρ' .

$$\hat{\pi} = \arg \max_{\pi} \mathbb{E}_{s \sim \rho', a \sim \pi} [\hat{r}_{\mathsf{MLE}}(s, a)] - \lambda \cdot \mathbb{E}_{s \sim \rho'} [\mathsf{KL}(\pi(\cdot \mid s) \| \pi_0(\cdot \mid s))].$$

RLHF in Multi-Armed Bandits

To understand the overfitting and overoptimization problems, we simplify the RLHF problem to consider a single-state multi-armed bandit formulation with pairwise comparisons. Instead of fitting a reward model and policy model with neural networks, we fit a tabular reward model in a K-armed bandit problem.

¹In practice, one may introduce an extra temperature parameter σ and replace all r^* with r^*/σ . Here we take $\sigma = 1$.

Consider a multi-armed bandit problem with K arms. Each arm has a deterministic ground-truth reward $r^*(k) \in \mathbb{R}, k \in [K]$. In this case, the policy becomes a distribution supported on the K arms $\pi \in \Delta([K])$. The sampling process for general RLHF reduces to the following: we first sample two actions a_i, a'_i from a joint distribution $\mu \in \Delta([K] \times [K])$, and then observe a binary comparison variable y_i following a distribution

$$\mathbb{P}(y_i = 1) = \frac{\exp(r^*(a_i))}{\exp(r^*(a_i)) + \exp(r^*(a'_i))},\\ \mathbb{P}(y_i = 0) = 1 - \mathbb{P}(y_i = 1).$$

Assume that we are given n samples, which are sampled *i.i.d.* from the above process. Let n(a, a') be the total number of comparisons between actions a and a' in the n samples. Let the resulting dataset be $\mathcal{D} = \{a_i, a'_i, y_i\}_{i=1}^n$. The tasks in RLHF for multi-armed bandit setting can be simplified as:

- 1. Reward learning: Estimate true reward r^* with a proxy reward \hat{r} from the comparison dataset \mathcal{D} .
- 2. Policy learning: Find a policy $\pi \in \Delta([K])$ that maximizes the proxy reward under KL constraints.

In the next two sections, we discuss separately the reward learning phase and policy learning phase, along with the reasons behind overfitting and overoptimization.

Overfitting in Reward Learning

For reward learning, the commonly used maximum likelihood estimator is the estimator that minimizes empirical cross-entropy loss:

$$\hat{r}_{\mathsf{MLE}} = \arg\min_{r} \hat{\mathcal{L}}_{\mathsf{CE}}(\mathcal{D}, r), \text{ where}$$

$$\hat{\mathcal{L}}_{\mathsf{CE}}(\mathcal{D}, \hat{r}) = -\frac{1}{n} \sum_{i=1}^{n} y_i \log \left(\frac{\exp(\hat{r}(a_i))}{\exp(\hat{r}(a_i)) + \exp(\hat{r}(a'_i))} \right)$$

$$+ (1 - y_i) \log \left(\frac{\exp(\hat{r}(a'_i))}{\exp(\hat{r}(a_i)) + \exp(\hat{r}(a'_i))} \right).$$
(3.2)

By definition, \hat{r}_{MLE} is convergent point when we optimize the empirical cross entropy fully. Thus the population cross-entropy loss of \hat{r}_{MLE} is an indicator for whether overfitting exists during reward training. We define the population cross entropy loss $\mathcal{L}_{CE}(r)$ as

$$-\mathbb{E}_{(a,a')\sim\mu}\left[\frac{\exp(r^{\star}(a))}{\exp(r^{\star}(a))+\exp(r^{\star}(a'))}\log\left(\frac{\exp(r(a))}{\exp(r(a))+\exp(r(a'))}\right)\right.\\\left.+\frac{\exp(r^{\star}(a'))}{\exp(r^{\star}(a))+\exp(r^{\star}(a'))}\log\left(\frac{\exp(r(a))}{\exp(r(a))+\exp(\hat{r}(a'))}\right)\right].$$

For a fixed pairwise comparison distribution μ , it is known that the maximum likelihood estimator \hat{r}_{MLE} converges to the ground truth reward r^* as the number of samples n goes to infinity.

Theorem 25 (Consistency of MLE, see, e.g., Theorem 6.1.3. of [109]). Fix $r^{\star}(K) = \hat{r}(K) = 0$ for the uniqueness of the solution. For any fixed μ , and any given ground-truth reward r^{\star} , we have that \hat{r}_{MLE} converges in probability to r^{\star} ; i.e., for any $\epsilon > 0$,

$$\lim_{n \to +\infty} \mathbb{P}\left(\|\hat{r}_{\mathsf{MLE}} - r^{\star}\|_{\infty} \ge \epsilon \right) = 0.$$

Here we view \hat{r}_{MLE} and r^{\star} as K-dimensional vectors.

The proof is deferred to Appendix B.4. This suggests that the overfitting phenomenon does not arise when we have an infinite number of samples. However, in the non-asymptotic regime when the comparison distribution μ may depend on n, one may not expect convergent result for MLE. We have the following theorem.

Theorem 26 (Reward overfitting of MLE in the non-asymptotic regime). Fix $r^*(a) = 1(a = 1)$ and $\hat{r}(K) = 0$ for uniqueness of the solution. For any fixed n > 500, there exists some 3-armed bandit problem such that with probability at least 0.09,

$$\mathcal{L}_{\mathsf{CE}}(\hat{r}_{\mathsf{MLE}}) - \mathcal{L}_{\mathsf{CE}}(r^{\star}) \ge C$$

for any arbitrarily large C.

The proof is deferred to Appendix B.5. Below we provide a intuitive explanation. The constructed hard instance is a bandit where $r^*(a) = 1(a = 1)$. For any fixed *n*, we set $\mu(1,2) = 1 - 1/n$, $\mu(1,3) = 1/n$.

In this hard instance, there is constant probability that arm 3 is only compared with 1 once. And with constant probability, the observed comparison result between arm 1 and arm 3 will be different from the ground truth. The MLE will assign $r(3) = +\infty$ since the maximizer of $\log(\exp(x)/(1 + \exp(x)))$ is infinity when x is not bounded. Thus when optimizing the empirical cross entropy fully, the maximum likelihood estimator will result in a large population cross-entropy loss. We also validate this phenomenon in Section B.3 with simulated experiments.

This lower bound instance simulates the high-dimensional regime where the number of samples is comparable to the dimension, and the data coverage is unbalanced across dimensions. One can also extend the lower bound to more than 3 arms, where the probability of the loss being arbitrarily large will be increased to close to 1 instead of a small constant.

Overoptimization in Policy Learning

After obtaining the estimated reward function \hat{r} , we optimize the policy $\pi \in \Delta([K])$ to maximize the estimated reward. In RLHF, one starts from an initial (reference) policy π_0 , and optimizes the new policy π to maximize the estimated reward \hat{r} under some constraint in KL divergence between π and π_0 . It is observed in [85] that as we continue optimizing the policy to maximize the estimated reward, the true reward of the policy will first increase then decrease, exhibiting the reward overoptimization phenomenon.

Consider the following policy optimization problem for a given reward model \hat{r} :

$$\max_{\pi \in \Delta([K])} \mathbb{E}_{a \sim \pi(\cdot)}[\hat{r}(a)] - \frac{1}{\lambda} \cdot \mathsf{KL}(\pi \| \pi_0).$$
(3.3)

Assuming that the policy gradient method converges to the optimal policy for the above policy optimization problem, which has a closed-form solution:

$$\pi_{\lambda}(a) = \frac{\pi_0(a) \cdot \exp(\lambda \cdot \hat{r}(a))}{\sum_{a' \in \mathcal{A}} \pi_0(a') \cdot \exp(\lambda \cdot \hat{r}(a'))}.$$
(3.4)

In the tabular case, we can derive a closed form solution for how the KL divergence and ground-truth reward change with respect to λ , thus completely characterizing the reward-KL tradeoff. We compute the KL divergence and ground-truth reward of the policy as

$$\begin{aligned} \mathsf{KL}(\pi_{\lambda} \| \pi_{0}) &= \frac{\sum_{a \in \mathcal{A}} \pi_{0}(a) \cdot \exp(\lambda \cdot \hat{r}(a)) \cdot \log(\exp(\lambda \cdot \hat{r}(a)) / (\sum_{a' \in \mathcal{A}} \pi_{0}(a') \cdot \exp(\lambda \cdot \hat{r}(a'))))}{\sum_{a' \in \mathcal{A}} \pi_{0}(a') \cdot \exp(\lambda \cdot \hat{r}(a))} \\ &= \frac{\sum_{a \in \mathcal{A}} \pi_{0}(a) \cdot \exp(\lambda \cdot \hat{r}(a)) \cdot \lambda \cdot \hat{r}(a)}{\sum_{a' \in \mathcal{A}} \pi_{0}(a') \cdot \exp(\lambda \cdot \hat{r}(a'))} - \log\left(\sum_{a' \in \mathcal{A}} \pi_{0}(a') \cdot \exp(\lambda \cdot \hat{r}(a'))\right)\right), \\ \mathbb{E}_{a \sim \pi_{\lambda}}[r^{\star}(a)] &= \frac{\sum_{a \in \mathcal{A}} \pi_{0}(a) \cdot \exp(\lambda \cdot \hat{r}(a)) \cdot \lambda \cdot r^{\star}(a)}{\sum_{a' \in \mathcal{A}} \pi_{0}(a') \cdot \exp(\lambda \cdot \hat{r}(a'))}. \end{aligned}$$

The above equation provides a precise characterization of how the mismatch between \hat{r} and r^* leads to the overoptimization phenomenon, which can be validated from the experiments in Section B.3. To simplify the analysis and provide better intuition, we focus on the case when $\lambda \to \infty$, i.e., when the optimal policy selects the best empirical arm without considering the KL constraint. In this case, the final policy reduces to the empirical best arm, $\pi_{\infty}(a) = 1(a = \arg \max_{a'} \hat{r}(a'))$.

By definition, π_{∞} is the convergent policy when we keep loosening the KL divergence constraint in Equation (3.3). Thus the performance of π_{∞} is a good indicator of whether overoptimization exists during policy training. We thus define a notion fo sub-optimality to characterize the performance gap between the convergent policy and the optimal policy:

$$\mathsf{SubOpt}(\hat{\pi}) \coloneqq \max_{a} \mathbb{E}[r^{\star}(a) - r^{\star}(\hat{\pi})].$$

We know from Theorem 25 that, asymptotically, the MLE for reward \hat{r}_{MLE} converges to the ground truth reward r^* . As a direct result, when using the MLE as reward, the sub-optimality of the policy π_{∞} also converges to zero with an infinite number of samples.

However, as a corollary of Theorem 26 and a direct consequence of reward overfitting, π_{∞} may have large sub-optimality in the non-asymptotic regime when trained from \hat{r}_{MLE} .

Corollary 27 (Reward overoptimization of MLE in the non-asymptotic regime). Fix $r^*(a) = 1(a = 1)$. For any fixed n, there exists some 3-armed bandit problem such that with probability at least 0.09,

$$\mathsf{SubOpt}(\hat{\pi}_{\infty}) \geq 1.$$

The proof is deferred to Appendix A.2. This suggests that \hat{r}_{MLE} also leads to the reward overoptimization phenomenon in the non-asymptotic regime. In Section B.3, we conduct simulation in the exact same setting to verify the theoretical results.

3.3 Methods: Pessimistic MLE and Iterative Data Smoothing

The problem of overfitting and overoptimization calls for a design of better and practical reward learning algorithm that helps mitigate both issues. We first discuss the pessimistic MLE algorithm in [316], which is shown to converge to a policy with vanishing sub-optimality under good coverage assumption.

Pessimistic MLE

In the tabular case, the pessimistic MLE corrects the original MLE by subtracting a confidence interval. Precisely, we have

$$\hat{r}_{\mathsf{PE}}(a) = \hat{r}_{\mathsf{MLE}}(a) - \lambda \cdot \sqrt{\frac{1}{n}},\tag{3.5}$$

where n is the total number of samples and $\lambda = \|(L + \epsilon I)_j^{-1/2}\|_2$ is the norm of the *j*-th column of the matrix $(L + \epsilon I)^{-1/2}$, where L is the matrix that satisfies $L_{a,a} = n(a)/n$, $L_{a,a'} = -n(a,a')/n, \forall a \neq a'$, and ϵ is a small constant. Intuitively, for those arms that are compared fewer times, we are more uncertain about their ground-truth reward value. Pessimistic MLE penalizes these arms by directly subtracting the length of lower confidence interval of their reward, ensuring that the arms that are less often compared will be less likely to be chosen. It is shown in [316] that the sub-optimality of the policy optimizing \hat{r}_{PE} converges to zero under the following two conditions:

- The expected number of times that one compares optimal arm (or the expert arm to be compared with in the definition of sub-optimality) is lower bounded by some positive constant $\mu(a^*) \ge C$.
- The parameterized reward family lies in a bounded space $|\hat{r}(a)| \leq B, \forall a \in [K]$.

This indicates that pessimistic MLE can help mitigate the reward overoptimization phenomenon. However, for real-world reward training paradigm, the neural network parameterized reward family may not be bounded. Furthermore, estimating the exact confidence interval for a neural-network parameterized model can be hard and costly. This prevents the practical use of pessimistic MLE, and calls for new methods that can potentially go beyond these conditions and apply to neural networks.

Iterative Data Smoothing

We propose a new algorithm, Iterative Data Smoothing (IDS), that shares similar insights as pessimistic MLE. Intuitively, pessimistic MLE helps mitigate the reward overoptimization issue by reducing the estimated reward for less seen arms. In IDS, we achieve this by updating the label of the data we train on.

Algorithm 2 Iterative Data Smoothing $(\mathcal{D}, \theta_0, \alpha, \beta)$

Input: The pairwise comparison dataset $\mathcal{D} = \{a_i, a'_i, y_i\}_{i=1}^n$. A parameterized reward model family $\{r_{\theta} : \mathcal{A} \mapsto \mathbb{R} \mid \theta \in \Theta\}$ with initialization $\theta_0 \in \Theta$. Two step sizes α, β . An empirical loss function

$$\mathcal{L}_{\theta}(\{y_i\}, \mathcal{D}) = -\frac{1}{n} \sum_{i=1}^{n} y_i \cdot \log\left(\frac{\exp(r_{\theta}(a_i))}{\exp(r_{\theta}(a_i)) + \exp(r_{\theta}(a'_i))}\right) + (1 - y_i) \cdot \log\left(\frac{\exp(r_{\theta}(a_i)) + \exp(r_{\theta}(a'_i))}{\exp(r_{\theta}(a_i)) + \exp(r_{\theta}(a'_i))}\right)$$

Initialize t = 0 and $y_{i,0} = y_i, \forall i \in [n]$. while r_{θ_t} does not converge **do**

$$\theta_{t+1} \leftarrow \theta_t - \alpha \cdot \nabla \mathcal{L}_{\theta}(\{y_{i,t}\}, \mathcal{D})$$

$$y_{i,t+1} \leftarrow (1-\beta) \cdot y_{i,t} + \beta \cdot \frac{\exp(r_{\theta_{t+1}}(a_i))}{\exp(r_{\theta_{t+1}}(a_i)) + \exp(r_{\theta_{t+1}}(a'_i))}$$

$$t \leftarrow t+1$$

end while Return: r_{θ_t} As is shown in Algorithm 2, we initialize $y_{i,0}$ as the labels for the samples y_i . In the *t*-th epoch, we first update the model using the current comparison dataset with labels $\{y_{i,t}\}_{i=1}^{n}$. After the model is updated, we also update the data using the model by predicting the probability $\mathbb{P}(y_i = 1)$ for each comparison (a_i, a'_i) using the current reward estimate \hat{r}_{θ_t} . We update each label $y_{i,t}$ as a weighted combination of its previous value and the new predicted probability.

Intuitively, $y_{i,t}$ represents a proxy of the confidence level of labels predicted by interim model checkpoints. The idea is that as the model progresses through multiple epochs of training, it will bring larger change to rewards for frequently observed samples whose representation is covered well in the dataset. Meanwhile, for seldom-seen samples, the model will make minimal adjustments to the reward.

Benefit of one-step gradient descent

Before we analyze the IDS algorithm, we first discuss why training for one to two epochs in the traditional reward learning approach works well [191]. We provide the following analysis of the one-step gradient update for the reward model. The proof is deferred to Appendix B.7.

Theorem 28. Consider the same multi-armed bandit setting where the reward is initialized equally for all K arms. Then after one-step gradient descent, one has

$$\forall a, a' \in [K], \hat{r}(a) - \hat{r}(a') = \alpha \cdot (n_+(a) - n_-(a) - (n_+(a') - n_-(a'))),$$

where $n_+(a), n_-(a)$ refers to the total number of times that a is preferred and not preferred, respectively.

Remark 29. The result shows that why early stopping in the traditional reward learning works well in a simple setting. After one gradient step, the empirical best arm becomes the the arm whose absolute winning time is the largest. This can be viewed as another criterion besides pessimism that balances both the time of comparisons and the time of being chosen as the preferred arm. When the arm a is only compared few times, the difference $n_+(a) - n_-(a)$ will be bounded by the total number of comparisons, which will be smaller than those that have been compared much more times. Thus the reward model will penalize those arms seen less. After updating the label with the model prediction, the label of less seen samples will be closer to zero, thus getting implicitly penalized.

Benefit of iterative data smoothing

Due to under-optimization, the estimator from a one-step gradient update might still be far from the ground-truth reward. We provide an analysis here why IDS can be better. Consider any two arms a, a' with n(a, a') observations among n total observations. By computing the gradient, we can write the IDS algorithm as

$$\hat{r}_{t+1}(a) - \hat{r}_{t+1}(a') = \hat{r}_t(a) - \hat{r}_t(a') + \frac{\alpha \cdot n(a, a')}{n} \\ \cdot \left((\hat{\mu}(a \succ a') \cdot y_t + \hat{\mu}(a \prec a') \cdot (1 - y_t)) \cdot \frac{\exp(\hat{r}_t(a'))}{\exp(\hat{r}_t(a)) + \exp(\hat{r}_t(a'))} \right. \\ \left. - (\hat{\mu}(a \prec a') \cdot y_t + \hat{\mu}(a \succ a') \cdot (1 - y_t)) \cdot \frac{\exp(\hat{r}_t(a)) + \exp(\hat{r}_t(a'))}{\exp(\hat{r}_t(a)) + \exp(\hat{r}_t(a'))} \right) \\ y_{t+1} = (1 - \beta) \cdot y_t + \beta \cdot \frac{\exp(\hat{r}_{t+1}(a))}{\exp(\hat{r}_{t+1}(a)) + \exp(\hat{r}_{t+1}(a'))},$$

/ ...

where we define $\hat{\mu}(a \succ a') = n(a \succ a')/n(a, a')$. One can see that the effective step size for updating \hat{r} is $\alpha \cdot n(a, a')/n$, while the effective step size for updating y is β . Assume that we choose α, β, l, m such that

$$\alpha \cdot l/n \ll \beta \ll \alpha \cdot m/n.$$

Consider the following two scales:

- When there are sufficient observations, $n(a, a') \ge m$, we know that $\beta \ll \alpha \cdot n(a, a')/n$. In this case, the update step size of y_t is much slower than \hat{r}_t . One can approximately take $y_t \approx 0$ or 1 as unchanged during the update. Furthermore, since $n(a, a') \ge m$ is large enough, $\hat{\mu}$ concentrates around the ground truth μ . In this case, one can see that the reward converges to the ground truth reward $\hat{r}_t \to r^*$.
- When the number of observations is not large, i.e., $n(a, a') \leq l$, we know that $\alpha \cdot l/n \ll \beta$. In this case, the update of \hat{r} is much slower than y_t . When the \hat{r}_0 are initialized to be zero, y_t will first converge to 1/2, leading to $\hat{r}_t(a) \approx \hat{r}_t(a')$ when t is large.

To formalize the above argument, we consider the following differential equations:

$$\dot{d}(t) = \alpha n \cdot \left(\left(\mu \cdot y(t) + (1 - \mu) \cdot (1 - y(t)) \right) \cdot \frac{1}{1 + \exp(d(t))} - \left((1 - \mu) \cdot y(t) + \mu \cdot (1 - y(t)) \right) \cdot \frac{\exp(d(t))}{1 + \exp(d(t))} \right)$$
$$\dot{y}(t) = \beta \cdot \left(\frac{\exp(d(t))}{1 + \exp(d(t))} - y(t) \right).$$
(3.6)

Here d represents the difference of reward between two arms a, a', and μ represents the empirical frequency $\hat{\mu}(a \succ a')$. Let the initialization be d(0) = 0, y(0) = 1. We have the following theorem.

Theorem 30. The differential equations in Equation (3.6) have one unique stationary point $d(t) = 0, y(t) = \frac{1}{2}$. On the other hand, for any α, β, n, T with $\beta T \leq \epsilon \ll 1 \ll \alpha nT$, one has

$$\left|\frac{\exp(d(T))}{1+\exp(d(T))} - \mu\right| \le \max(2(1-\exp(-\epsilon)), \exp(-\mu(1-\mu)\alpha nT))$$
$$y(T) \ge \exp(-\epsilon).$$

The proof is deferred to Appendix B.8. The above argument only proves convergence to the empirical measure μ . One can combine standard concentration argument to prove the convergence to the ground truth probability. The result shows that when choosing α, β carefully, for the pair of arms with a large number of comparisons, the difference of reward will be close to the ground truth during the process of training. As a concrete example, by taking $\alpha = n^{-1/2}, \beta = n^{-1}T^{-2}, \epsilon = \beta T$, we have

$$\left|\frac{\exp(d(T))}{1+\exp(d(T))} - \mu\right| \le \max(2n^{-1}T^{-1}, \exp(-\mu(1-\mu)n^{1/2}T)).$$

For those pairs of comparisons with a large sample size n, the estimated probability is close to the ground truth probability. On the other hand, for those pairs that are compared less often, the difference d(t) is updated less frequently and remains close to the initialized values. Thus the algorithm implicitly penalizes the less frequently seen pairs, while still estimating the commonly seen pairs accurately. We also present an alternative formulation of IDS in Appendix B.2.

In summary, the IDS algorithm enjoys several benefits:

- For a sufficient number of observations, the estimated reward converges to the ground truth reward; while for an insufficient number of observations, the estimated reward remains largely unchanged at the initialization. Thus the reward model penalizes the less observed arms with higher uncertainty.
- It is easy to combine with neural networks, allowing arbitrary parametrization of the reward model.
- It utilizes the soft labels starting from the second epoch, which can be more effective than hard labels according to the literature on knowledge distillation [106, 307].

3.4 Experiments

We provide one experimental result in Figure 3.2, and leave all the details of experiments to Appendix B.3, where we conduct simulation study on multi-armed bandit environment, and real-world experiments with the human-labeled Helpfulness and Harmlessnes (HH) dataset

from [18] and TLDR dataset², along with comparison with more baseline algorithms including Laplace Smoothing [41]. In Figure 3.2, we train a reward model with 1 Billion parameters using HH dataset, and fine-tune the language model with the reward model. One can see that after 1 epoch of training, the test loss of MLE begins to increase, while IDS enables continuous decrease of test loss for more than 3 epochs. Furthermore, when tuning the language model with the proxy reward model, the ground-truth reward for IDS grows higher than that of MLE.



Figure 3.2. Comparisons of MLE and IDS when the reward is parameterized by a neural network.

3.5 Conclusions

We have presented analyses and methodology aimed at resolving the problems of overfitting and overoptimization for RLHF. We show that our proposed algorithm, IDS, helps mitigate these issues. While we identify the underlying source of reward overfitting and overoptimization as the variance of the human preference data, it is also possible that bias also contributes to these phenomena. In future work, it is interesting to pursue further theoretical analysis of the IDS algorithm, and explore potential applications beyond reward training in the generic domains of classification and prediction.

²https://huggingface.co/datasets/CarperAI/openai_summarize_comparisons

Chapter 4

Real World RLHF Experiments: Starling-7B

4.1 Introduction

This chapter presents Starling-7B, a strong 7B chat model undergoing human preference alignment, along with its training dataset Nectar, a high-quality preference dataset collected by prompting GPT-4 to rank responses on 182,954 chat prompts. Each prompt comprises 7 responses, distilled from a variety of models such as GPT-4, GPT-3.5-instruct, GPT-3.5-turbo, Mistral-7B-Instruct, and Llama2-7B-chat, ranked by GPT-4. This equates to a total of 3.8 million high-quality pairwise comparisons. We propose an internal pairwise rating technique, where the model considers all pairings before providing a ranking decision, leveraging the proven pairwise rating capability of LLMs without the cost of individual pairwise calls.

We also introduce novel techniques to reduce positional bias when prompting GPT-4 for rankings. This enables reinforcement learning from AI feedback (RLAIF), where we learn a proxy of human preferences from (potentially biased) GPT-4 ranking data. We compare the Nectar dataset with existing preference dataset [59, 18, 71] in Table 4.1.

Dataset	#Prompt	K	Focus Field	Ranking Source	Ranking Type	Response Source	# Pairwise Comparisons
Nectar	183K	7	Diverse	AI	Ordinal	Strong + Weak LLM	3.8M
Ultrafeedback	64K	4	Diverse	AI	Cardinal	Strong + Weak LLM	384K
Anthropic-HH	161K	2	Safety	Human	Ordinal	Weak LLM	161K
Stack Overflow	20M	varies	Coding	Human	Cardinal	Human	20M +
SHP	385K	2	Diverse (Reddit)	Human	Ordinal	Human	385K

 Table 4.1: Existing RLHF datasets compared to Nectar

Furthermore, we open source our reward model suites, Starling-RM-7B and Starling-RM-34B. The models are trained using the K-wise loss [316] on the Nectar dataset. We benchmark various reward model training pipelines across metrics such as human preference, truthfulness, and safety in Section 4.4.

To create the chat model Starling-LM-7B-alpha, we fine-tune Openchat-3.5 [268] with proximal policy optimization (PPO) [229] on the learned reward model Starling-RM-7B.

As a result, the MT-Bench score improves from 7.81 to 8.09, while the AlpacaEval score improves from 88.51% to 91.99%, and a human evaluation ELO increases from 1072 to 1087 on Chatbot Arena [48]. We also release a new version Starling-LM-7B-beta based on a larger reward model Starling-RM-34B and the initialized language model Openchat-3.5-0106, which increases the human evaluation ELO from 1089 to 1118 on Chatbot Arena. These metrics highlight the improvement of Starling-LM-7B in providing helpful responses.



Figure 4.1: MT Bench Evaluation of the Starling-LM-7B-alpha model.

To facilitate research and understanding of RLHF mechanisms, we open-source the Nectar dataset, the reward models, and the language models, along with their pipeline here: https://github.com/efrick2002/Starling. We hope the potential of these open-source contributions, including the dataset, reward model, and language model, enrich the understanding of RLHF mechanisms and fuel further alignment research.

4.2 Related Work

RLHF and RLAIF

Previous work has shown that human evaluators can effectively provide a reward signal to train agents in complex RL environments [54]. Specially, RLHF has shown immense effectiveness when applied to aligning LLMs to human preferences [191, 18, 6, 190, 262]. Since RLHF uses actual human preference data, data collection is a very expensive bottleneck. To mitigate this, recent efforts have tried RLAIF [17, 153, 6]. [174] utilizes the zero-shot performance of LLMs for ranking tasks while [199] explores LLMs' effectiveness as knowledge bases, which together may provide a lower cost solution to mimic human preference via careful rubric design and prompting.

Positional Bias and K-wise Rankings

Previous work has show that LLMs are not necessarily fair evaluators and exhibit significant positional bias [270]. As result, most LLMs human preference ranking strategies only use more reliable pairwise comparisons where ranking outputs are more stable [201]. Black-box pure K-wise ranking strategies have struggled to find success, citing problems with outputs missing choices, repeating choices, inconsistently ranking, or being irrelevant [201]. Some effort to generalize pairwise prompting to K-wise prompting has found success by utilizing multiple pairwise evaluations to sort a K-wise ranking. This includes aggregating all pairwise comparisons or utilizing pairwise comparisons for heapsort [201]. Using K choose 2 pairwise comparisons to create a final K-wise ranking is not favorable due to $O(K^2)$ evaluator queries, or $O(K \log K)$ for sorting algorithms. Ideally, a K-wise ranking could be completed zero-shot with minimal positional bias with just a single evaluator query. This chapter present a prompting framework that approaches this ideal.

4.3 Nectar Dataset

In this section, we discuss the creation of the Nectar dataset. The dataset is composed of a set of prompts, with 7 responses for each prompt distilled from existing models, along with GPT-4-based ranking for the responses. We provide details about the prompt collecting, response collection, and response ranking respectively.

Collecting Prompts

We aim to curate a dataset with diverse prompts from different sources, including adversarial jailbreaking prompts. The collected prompts are used for both reward learning and policy learning in RLHF. We collect data from a range of existing validated datasets, including ShareGPT, Anthropic-HH [18], UltraFeedback [59], and Lmsys-Chat-1M [311]. We first sub-sample prompts from each different category, and then combine the same prompts and their corresponding model-generated responses, resulting in a dataset with 182,954 prompts. The composition of the prompts is provided in Table 4.2.

Table 4.2: Number of Prompts From Each Source						
Source	Anthropic-HH	Lmsys-Chat-1M	UltraFeedback	ShareGPT		
Count	74097	43545	40411	25113		

Collecting Responses

For each prompt in the dataset, we distilled a response from GPT-4, GPT-4-0613, GPT-3.5-Turbo, GPT-3.5-Turbo-Instruct, Mistral-7B-instruct, and Llama-2-7B-chat. Each prompt also comes with a default response in the corresponding dataset. As result, we get a dataset

with 182,968 prompts and 7 responses per prompt. GPT-4-0613 was given a system prompt to help set the chatting context yielding better responses. GPT-3.5-Turbo-Instruct was chosen for the lack of extensive human feedback alignment, unlike many of the other models present, leading it to be more likely to comply with jailbreak prompts with high quality responses. We provide more detailed analysis on the distribution of the model responses in Appendix C.1.

Collecting Ranking Data

Given some ranking rubric and user prompt from the prompt dataset, we ask GPT-4 to rank the 7 responses from best to worst in accordance with rubric. The goal of the rating prompt was to capture a bimodal rubric strategy designed to balance helpfulness and harmlessness: Maximally answer good-natured conversations and harmlessly answer all conversations. We create the following rubric:

The helpfulness rubric is as follows:

- 1. Is the completion relevant? The completion should be the best possible assistant answer.
- 2. Is the completion accurate? The completion should accurately fulfill the prompt's request.
- 3. If a creative answer is expected, is the completion creative? If an analytical answer is expected, is the completion factual/objectively correct?
- 4. Is the completion detailed? The completion should at minimum satisfy the full level of detail required by the prompt.

The harmlessness rubric is as follows:

- 1. Does the completion adhere to moral and ethical standards?
- 2. Is the completion truthful?
- 3. Is the response polite and respectful?
- 4. If the request given is not good natured, does the completion politely decline to reply while giving a short explanation why they cannot comply with the request? Completions should always maintain harmlessness in the face of adversarial requests.

Our objective is to de-emphasize the role of helpfulness in the evaluation of responses to prompts of a negative nature, intending primarily, to maintain harmlessness. This approach is designed to prevent harmful outputs from obtaining higher scores than their harmless equivalents, simply because they adhere more closely to the instructions. Conversely, when dealing with standard prompts, we rank the responses purely based on the helpfulness.

As one of the major challenges in ranking 7 responses, we observe very serious positional bias in GPT-4 when ranking more than 2 responses. We discuss the findings in the section below.

Positional Bias

The notion of positional bias has been studied in a pairwise context, with results indicating that the first or the second index could be favored, contingent on the ranking model deployed [270]. We have similar observations in our K-wise rating settings.

For each prompt, we shuffle the 7 responses uniformly at random and prompt GPT-4 to provide a list-wise ranking. We observe that GPT-4 exhibited a strong preference for selecting responses seen at the beginning. Such responses won at a rate ten times greater than completions viewed at the sixth index (See Figure 4.2, when K = 7). These results are consistent with earlier observations, implying that GPT-4 has an inherent tendency to favor the first several responses [270]. Without effective strategies to address this bias, it becomes evident that earlier responses are overwhelmingly preferred over later ones. Moreover, our observations indicate that this positional bias worsens as K increases unless remedial actions are taken.



Figure 4.2. Winning Response Index Distribution with a naive prompt: positional bias for different values of K for K-wise comparisons. (n=200)

Figure 4.2 visualizes the increasing positional bias problem for larger K. While in pairwise rating situations, the bias is manageable, larger K-wise rating procedures are unusable. In the following sections, we explore various strategies to combat the positional bias problem for large K.

We introduce novel prompting techniques to mitigate the positional bias by first generating all pairwise rankings, and then summarizing the results into list-wise rankings. We leave the details to Appendix C.1. We also leave additional analysis, including model pairwise win rates and pairwise average ranking differential heatmaps to Appendix C.1.

4.4 Reward Learning

Formulation

In this section, we discuss the training of the reward model. The reward learning procedure in RLHF can be formulated in a contextual bandit environment. In a contextual bandit, for the *i*-th sample, a state (prompt) s^i is first sampled from some fixed distribution ρ . Given the state s^i , K actions $(a_0^i, a_1^i, \dots, a_{K-1}^i)$ are sampled from some joint distribution $\mathbb{P}(a_0, \dots, a_{K-1} \mid s^i)$. Let $\sigma^i : [K] \mapsto [K]$ be the output of the human labeller, which is a permutation function that denotes the ranking of the actions. Here $\sigma^i(0)$ represents the most preferred action. We use $a_0 \succ a_1$ to denote the event that the action a_0 is more preferred compared to a_1 . A common model on the distribution of σ under K-ary comparisons is a Plackett-Luce model [200, 173]. The Plackett-Luce model defines the probability of a state-action pair (s, a_i) being the largest among a given set $\{(s, a_i)\}_{i=0}^{K-1}$ as

$$\mathbb{P}(a_i \succ a_j, \forall j \neq i \mid s) = \frac{\exp(r_{\theta^\star}(s, a_i))}{\sum_{j=0}^{K-1} \exp(r_{\theta^\star}(s, a_j))}$$

Here r_{θ^*} is the ground truth reward function with parameter θ^* . Moreover, one can calculate the probability of observing the permutation σ as¹

$$\mathbb{P}(\sigma \mid s, \{a_i\}_{i=0}^{K-1}) = \prod_{i=0}^{K-1} \frac{\exp(r_{\theta^{\star}}(s, a_{\sigma(i)}))}{\sum_{j=i}^{K-1} \exp(r_{\theta^{\star}}(s, a_{\sigma(j)}))}$$

When K = 2, this reduces to the pairwise comparison considered in the Bradley-Terry-Luce (BTL) model [25]. In this case, the permutation σ can be reduced to a Bernoulli random variable, representing whether a_0 is preferred compared to a_1 . Concretely, for each queried state-actions pair (s, a_0, a_1) , we observe a sample y from a Bernoulli distribution with parameter $\frac{\exp(r_{\theta^*}(s,a_1))}{\exp(r_{\theta^*}(s,a_0)) + \exp(r_{\theta^*}(s,a_1))}$; i.e., for any $l \in \{0, 1\}$,

$$\mathbb{P}(y = l \mid s, a_0, a_1) = \frac{\exp(r_{\theta^{\star}}(s, a_l))}{\exp(r_{\theta^{\star}}(s, a_0)) + \exp(r_{\theta^{\star}}(s, a_1))}.$$

By observing the sampled dataset $\mathcal{D} = \{s^i, a_0^i, a_1^i, \cdots, a_{K-1}^i, \sigma^i\}$, we would like to estimate the ground truth reward r_{θ^*} with some proxy reward $r_{\hat{\theta}}$. We discuss in the following section some different choices of estimators we benchmark in the chapter.

Methods

Similar to the default methodology of training reward models [323, 191, 228], we remove the last layer of a pre-trained language model and concatenate with a linear layer for reward prediction. We take Llama-2-7B-Chat [262] as the base model for reward model. Observing the various performances of the 7B Llama-based reward models, we also train a 34B reward model utilizing Yi-34B-Chat as the base model [8]. The reward model takes in a prompt and response, and outputs a scalar representing whether the response is helpful and harmless given the prompt.

For learning the reward from K-wise comparisons, we consider three methods. The first method is the original method from [323, 191, 228], which decomposes each K-wise comparison into K(K-1)/2 pairs of pairwise comparisons, and then minimize the cross entropy loss. When there is only pairwise comparison, this is the maximum likelihood estimator under

¹In practice, one may introduce an extra temperature parameter γ and replace all r_{θ^*} with r_{θ^*}/γ . Here we take $\gamma = 1$.

Bradley-Terry-Luce model [25]. It has also been shown to converge to the ground truth when the reward model is linear and well-specified with $K \ge 2$ [316].

$$\hat{r}_{\mathsf{MLE}_{2}} \in \arg\min_{r} \ell_{\mathcal{D}}(r),$$

where $\ell_{\mathcal{D}}(r) = -\frac{1}{n} \sum_{i=1}^{n} \sum_{j=0}^{K-1} \sum_{k=j+1}^{K-1} \log\left(\frac{\exp(r(s^{i}, a^{i}_{\sigma_{i}(j)}))}{\exp(r(s^{i}, a^{i}_{\sigma_{i}(j)})) + \exp(r(s^{i}, a^{i}_{\sigma_{i}(k)}))}\right)$

The second method directly applies the maximum likelihood estimator under Plackett-Luce model, which is shown in [316] to be asymptotically more efficient than \hat{r}_{MLE_2} .

$$\hat{r}_{\mathsf{MLE}_{K}} \in \arg\min_{r} \ell_{\mathcal{D}}(r),$$

where $\ell_{\mathcal{D}}(r) = -\frac{1}{n} \sum_{i=1}^{n} \sum_{j=0}^{K-1} \log \left(\frac{\exp(r(s^{i}, a^{i}_{\sigma_{i}(j)}))}{\sum_{k=j}^{K-1} \exp(r(s^{i}, a^{i}_{\sigma_{i}(k)}))} \right).$

The third method is the iterative data smoothing algorithm proposed in [317], which iteratively updates the model with the data, and then updates the soft data label with the model after each epoch to prevent overfitting to the noisy data. We denote the final estimator as \hat{r}_{smooth} .

Reward Model Evaluation

In this section, we benchmark the results of the three methods detailed in 4.4 . Additionally, we report the result of training on a larger base model.

Model	Human Acc.	Truth Acc.	Safety Acc.	Verbose Acc.
\hat{r}_{MLE_K} (7B)	0.763	0.647	0.759	0.167
\hat{r}_{MLE_2} (7B)	0.751	0.636	0.729	0.133
\hat{r}_{smooth} (7B)	0.761	0.598	0.737	0.100
\hat{r}_{MLE_K} (34B)	0.807	0.712	0.782	0.367

Table 4.3: Human, truth, safety, and verbosity accuracy for each reward model.

Model	Human Loss	Truth Loss	Safety Loss	Verbose Loss
\hat{r}_{MLE_K} (7B)	0.552	1.881	0.508	2.365
\hat{r}_{MLE_2} (7B)	0.704	1.933	0.620	3.327
\hat{r}_{smooth} (7B)	0.537	2.043	0.562	1.956
\hat{r}_{MLE_K} $(\bar{3}\bar{4}\bar{\mathrm{B}})$	0.468	1.766	0.48	1.461

Table 4.4: Human, truth, safety, and verbosity loss for each reward model.

We evaluate our trained reward models across four different categories: Human, truth, safety, and verbosity. Evaluation accuracy is measured as $\frac{1}{N} \sum 1(\hat{r}(\text{Best}) > \hat{r}(\text{Other}))$.

Evaluation loss is measured with log loss over the softmax probabilities arriving from the reward values: $\frac{1}{N} \sum \log(\frac{e^{\hat{r}(\text{Best})}}{e^{\hat{r}(\text{Best})} + e^{\hat{r}(\text{Other})}})$. In *K*-wise case where K > 2, the *K*-wise loss is sum of the pairwise loss for each $\binom{K}{2}$ pairings. The accuracy and loss results for each trained reward model are shown in Table 4.3 and Table 4.4, respectively. Additional details on these evaluation metrics and what they represent can be found in Appendix C.2. Note, the models chosen for release as Starling-RM-7B and Starling-RM-34B are the 7B and 34B models trained under the \hat{r}_{MLE_K} formulation, respectively, as we find \hat{r}_{MLE_K} to produces the strongest reward models in accordance to the evaluations. Additionally, likely as result of the significantly stronger base model, Starling-RM-34B (\hat{r}_{MLE_K} (34B)) exceeds all 7B counterparts in performance for every evaluation metric in Table 4.3 and Table 4.4.

Model	Chat	Chat Hard	Safety	Reasoning	Average
Starling-RM-34B	96.9	57.2	88.2	88.5	82.7
Tulu-2-DPO-70B	97.5	60.5	83.9	74.1	79.0
Mixtral-8x7B-Instruct-v0.1	95.0	64.0	73.4	78.7	77.8
Nous-Hermes-2-Mistral-7B-DPO	92.2	60.5	82.3	73.8	77.2
Zephyr-7B-alpha	91.6	62.5	74.3	75.1	75.9
${f Starling}-{f RM}-7{f B}-{f alpha}$	98.0	45.8	85.8	57.4	71.8
oasst-rm-2.1-pythia-1.4b	88.5	48.5	65.3	78.0	70.1
UltraRM-13B	96.1	58.6	54.3	65.4	68.6
Beaver-7B-v1.0-Cost	60.9	45.0	81.5	46.7	58.5

Other Benchmarks

 Table 4.5: Reward Bench scores for various models.

On Reward Bench, Starling-RM-34B achieves state-of-the art performance [150]. Most notably in the "Reasoning" category on Reward Bench, Starling-RM-34B has a large improvement over Starling-RM-7B, suggesting that larger and more capable base models may help the reward model's performance in downstream complex reasoning tasks. On the contrary, the "Chat" category performance seems to saturate quickly. Additionally, both Starling models surpass all other models in the "Safety" category, including models specifically trained for safety such as [60]'s Beaver-7B-v1.0-Cost. These results indicate that the inclusion of prompts of good and bad nature along with a judging rubric that enforces harmlessness despite helpfulness yields safer models on par with models more solely focused on harmlessness— all while maintaining exceptional performance on helpfulness categories.

4.5 Policy Learning

In policy learning stage, we use proximal policy optimization (PPO) [229] to fine-tune the language model based on the learned reward model. To create Starling-LM-7B-alpha, we fine-tune Openchat-3.5 [268] with the learned reward model Starling-RM-7B-alpha.

As a result, the MT-Bench score improves from 7.81 to 8.09, while the AlpacaEval score improves from 88.51% to 91.99%, and a human evaluation ELO increases from 1072 to 1087 on Chatbot Arena [48]. We also release a new version Starling-LM-7B-beta by fine-tuning Openchat-3.5-0106 on the larger reward model Starling-RM-34B, which increases the human evaluation ELO from 1089 to 1118 on Chatbot Arena.

Implementation Details

We implement the PPO algorithm using the TRLX library [100], incorporating all standard tricks of PPO as documented in [112]. In our implementation, we decouple the actor (the LM) and the critic, allowing for independent gradient updates. We find out that PPO is highly unstable. For example, after exposure to only 1000-2000 prompts, the model rapidly learns to generate excessively verbose outputs. To mitigate this instability and improve training robustness, we propose the following techniques:

Shifting the reward mean for length control: During our experiments, we observe that the reward model may assign highly negative rewards to the initial actor's outputs. Although the absolute magnitude of the reward does not convey preference information, and only the relative difference in reward between two responses indicates preference [280], the overall reward magnitude significantly affects the generated response length. We find that a very negative reward can cause the model to become excessively verbose after just a few gradient updates. To address this issue, we propose adding a constant to the reward to make it slightly positive. The intuition behind this approach is as follows: during the initial gradient steps, we should increase the likelihood of the end-of-sequence (EOS) token. which is equivalent to making the response shorter. We can estimate the advantage of the EOS token as $(r_{eos} + V_{eos}) - r_{prev}$, where r_{prev} is the reward assigned to the token before the EOS token (usually the negative KL penalty), r_{eos} is the reward scored by the reward model, and V_{eos} is the critic's value estimate for the EOS token. During initialization, we know that r_{prev} and V_{eos} are approximately zero, so a slightly positive r_{eos} will penalize the length. We observe that after applying the reward mean shift, the model initially generates slightly shorter responses and then gradually increases the response length as training progresses, driven by the verbosity preference in the reward.

Pretraining the critic model: During the initial stages of training, we observe that a randomly initialized critic (initialized from the language model with a linear layer and Gaussian-initialized head) can negatively impact the early performance of the language model. This is evident from a slight decrease in the MT-Bench score during the first 30% of the training process. To minimize early performance degradation, we propose pretraining the critic. We begin by conducting an RL run with a randomly initialized critic and the supervised fine-tuned (SFT) actor. We then use the final critic model from this run as the initialization for a new run with the same SFT actor. With this modification, we observe that the actor optimizes the reward more rapidly compared to the case without critic model pretraining. This approach helps to stabilize the early stages of training and allows the actor to more effectively optimize the reward signal. Full parameter tuning yields the best results: In our early experiments, we initialize both the actor and critic from the supervised fine-tuned (SFT) model, while unfreezing only the top 4 layers of the actor and critic during the RL stage. Although this approach improves the MT-Bench score from 7.81 to 8.09, the improvement is significantly smaller compared to initializing the actor and critic separately and performing full parameter tuning. By initializing the actor and critic separately and tuning all parameters, we can improve the score from 7.81 to 8.33, starting from the same SFT model. While MT-Bench serves as a proxy for model performance, our internal human evaluation also confirms that the full parameter tuned model is more preferred by human raters. This finding highlights the importance of full parameter tuning in achieving the best possible performance gains during the RL stage.

Hyperparameter Tuning

We perform extensive hyper-parameter searching to find the best configuration. We mainly focus on tuning the learning rate, training batch size, and KL penalty. These hyperparameter choices aim to strike a balance between model performance, and computational efficiency while minimizing undesirable behaviors such as over-optimization and excessive response length.

Our findings suggest that a slightly larger learning rate can be beneficial in reducing the over-optimization issue. When using a small learning rate and a longer training duration, the exploration in online RL may generate unusual responses that still receive high rewards due to the discrepancy between the ground truth reward and the proxy reward.

Additionally, we observe that a slightly larger KL penalty helps to stabilize the generated response length, although it can only mitigate the increase in length and not completely prevent it. The final run uses an KL penalty 0.01 and learning rate 10^{-6} . We also find that extremely large batch sizes do not yield observable improvements in human evaluation. Therefore, we utilize a medium batch size to facilitate faster iteration and experimentation, sampling 512 responses to form the replay buffer. Furthermore, the policy and critic updates are performed using a micro-batch size of 32.

Checkpoint Selection

We adopt a multi-layer checkpoint filtering strategy, where a checkpoint must pass all the previous filters to be considered as a candidate for final release. Our criteria are as follows:

- **KL-Reward Trade-off:** We require the model to achieve a predefined reward threshold while maintaining a KL-divergence lower than a specified KL threshold. This ensures that the model improves its performance without deviating too far from the base model's output distribution.
- Verbosity Monitoring: We observe that the reward model often prefers longer responses, which can lead the language model to generate unnecessarily verbose ex-

planations or even repeated outputs. To mitigate this issue, we monitor the average generated response length on a held-out validation set of prompts during training and filter out checkpoints that exceed a predefined length threshold.

- Measuring Instruction-Following Capability: We directly measure the model's ability to follow instructions using IFEval [315]. IFEval focuses on 25 types of "verifiable instructions", such as "write in more than 400 words" and "mention the keyword 'AI' at least 3 times." We find that steering the language model toward human preferences with RLHF can sometimes compromise its ability to precisely follow instructions. To address this, we predefine a threshold based on the base model's score and ensure that the checkpoint's score does not fall below this threshold.
- MT-Bench Evaluation: For checkpoints that pass all the previous tests, we evaluate the model using GPT-4 as a proxy. MT-Bench is a well-understood proxy for human preference; the evaluation process involves generating responses to a fixed set of prompts spanning a diverse range of topics and using GPT-4 to score the responses. We find that PPO can significantly increase the MT-Bench score. We set a reasonably high MT-Bench threshold, as we observe that scores higher than this threshold do not necessarily indicate a true human-perceivable performance difference.
- Internal Human Evaluation: We create a fixed set of 10-20 challenging prompts that assess the model's reasoning, instruction-following, and verbosity. We then judge the quality of the generated responses through internal human evaluation.

4.6 Conclusion

In this chapter, we produce the open-source RLHF pipeline that improves the helpfulness and harmlessness of the chat model. We release Nectar, a first-of-its-kind open-source high-quality 7-wise ranking dataset, and encourage future work to further explore the potentials of this dataset. Furthermore, we explore the effects of various reward training formulations on reward model performance. Finally, we perform extensive hyperparameter tuning on PPO to maximize performance gain in the RL stage. Ultimately, we hope to promote open source research and democratize the access of strong LLMs by releasing the dataset, methods, and the models trained from our RLHF procedure. We show that our data and training pipelines are capable of creating state-of-the-art reward and language models, pushing the limits of open-source LLMs.

There are also limitations with respect to the data and methodology, which may be important as future work. The Nectar dataset is collected purely from GPT-4-based preference, which may contain significant bias that is inconsistent with true human preferences. It would be worth exploring methods that mitigate biases from synthetic data. In addition, Nectar focuses more the helpfulness and harmlessness of the responses and less on the instruction following property. As such, according to our observations, the resulting language model can



Figure 4.3. (Left) Impact of reward shifting constant on response length. With no reward shifting (green), the initial actor starts with a reward around -7.3. Three different reward shifting parameters are tested: 0 (no shift), 7.2 (starting with a slightly negative reward), and 7.5 (starting with a slightly positive reward). The results show that starting with a slightly positive reward can better control the response length while achieving the same final reward. (Right) Optimal run with base model Openchat-3.5-0106. We evaluate the model's performance using two metrics: (Orange) Using Starling-RM-34B to score the responses generated from a fixed validation prompt set at every step. A step consists of creating a replay buffer of length 512 and performing 16 gradient updates with a micro-batch size of 32; (Blue) Evaluating the language model on MT-Bench every 5 steps. Our findings show that while the reward measured by the reward model increases throughout the training process, the MT-Bench score, which is considered a better proxy for human preference, starts to decrease after step 25.

be less capable of following exact instructions. There may also be improved reward training and policy learning algorithms leading to better reward models and language models. We hope future work can explore these areas in greater depth.

Chapter 5

Efficient Serving of Large Language Models

5.1 Introduction

The recent emergence of Large Language Models (LLMs) and foundation models has significantly increased the capabilities of AI systems [29, 188, 323, 191, 6, 21, 53, 273, 91]. This progress comes at a cost, however, of increased resource consumption and latency during both training and inference, presenting challenges not only in real-world deployment but also in terms of environmental impact and energy usage [238, 197, 24]. For instance, LLM-based chatbots typically consist of large transformer-based networks with parameter counts ranging from one to several hundred billion [314]. Moreover, the auto-regressive nature of LLMs exacerbates the issue of latency and resource consumption because the model can only generate one token at a time. Thus, compared to traditional AI-powered services, language model inference costs are much higher and the latency is significantly longer, making it nearly impossible to process each query using LLMs in high-throughput query systems such as search engines.

In this chapter, we explore two simple yet effective strategies to mitigate this problem: (1) employing a caching system to store previous queries, and (2) developing a model multiplexer to choose the most appropriate model from a set of models for processing the queries. The general workflow of our proposed LLM-based inference system is shown in Figure 5.1: upon receiving a query or prompt, we initially check if it can be retrieved from the cache. If the query is not found in the cache, we employ the model multiplexer to determine which model should be used for processing it first, based on the estimated cost for both models.

The choice of cost function and models can vary based on the goal. One measure of cost, for example, could be floating point operations (FLOPs). Other alternatives could include the number of API calls as a measure of resource consumption, latency as a measure of time consumption, or a score provided by a user as a measure of user satisfaction. The cost could also be a weighted sum of multiple factors. For the models, a natural choice would be to

have a small and a large model, where the small model costs less and is also less accurate, and the large model has a higher cost and also provides higher accuracy. Another alternative would be to have models with expertise in different areas, i.e., each model has high accuracy in its own area of expertise. We provide more discussion in Appendix D.1.



Figure 5.1: A workflow for LLM-based inference with caching and model multiplexing.

There is a long history of existing literature on caching algorithms, with prominent applications including computer architecture and web retrieval [243, 269, 146]. Existing caching algorithms deal with queries with different frequencies and cost, and must also provide guidelines for choosing the cache size. In addition to these well-known difficulties, the use of caching for LLMs raises new challenges, including:

- The need for fuzzy search. Since the prompt lies in a discrete space that is exponentially large with respect to the token size, it is impossible to match and save all distinct queries. Thus, to be at all useful, approximate matching and grouping is required when retrieving queries saved in the cache.
- The randomness of the cost. The cost for processing each query is a random variable that depends on the query and has a large variance due to the auto-regressive generation procedure and the difference in the length and quality of generated responses. When combined with the long-tailed distribution of the query frequency, the estimation of the cost requires a non-trivial algorithm design.
- The effect of model multiplexing. When the cache system is combined with the model multiplexer, the estimation of cost must change accordingly to take into consideration the different costs induced by various models.

For the fuzzy search problem, semantic search or vector-embedding-based ideas provide a systematic solution that includes embedding extraction and matching algorithms [20, 34, 134]. To simplify the problem, we assume that there exists some semantic search oracle that can group the prompts with the same semantic meaning and that the total cache size is limited by the number of queries, ignoring the difference in cache size between each individual query and response.

The remainder of this chapter is organized as follows. In Section 5.2, we formally define the pipeline of caching and model multiplexing. In Section 5.3, we study the optimality of the Least Expected Cost (LEC) caching strategy, which estimates the frequency and cost of processing each query, and evicts the one with the least estimated expected cost when there is only one model to call. In section 5.4, we consider the case when we have access to two models, and jointly design optimal caching and model multiplexer. In both sections, we start by assuming there are infinite samples and then analyze the offline and online learning cases where the cost and frequency need to be learned from data. The experimental results are presented in Section 5.5. We discuss the potential choices of cost, model, and output in the real world in Appendix D.1. We provide a brief discussion of the generalization to variable cache sizes in Appendix D.2 and of the generalization to multi-model multiplexing in Appendix D.3.

Related work

Cache replacement algorithms Traditional cache replacement algorithms investigate optimal ways to cache queries with different frequencies, costs, and cache sizes. To address varying frequencies, a standard approach is to use a Least Frequently Used (LFU) or Least Recently Used (LRU) cache eviction strategy [152]. These have been proven to be optimal for both adversarial and stochastic queries [251, 30]. Caching has also been combined with machine learning advice and online learning analysis in the literature [35, 242, 128, 101, 182, 73]. When varying costs and varying frequencies exist simultaneously, [130, 13] propose and study the Greedy Dual-Size with Frequency (GDSF) replacement algorithm, which takes both frequency and cost into consideration. [15] proposes the Least Expected Cost (LEC) algorithm, which is similar to GDSF, except that it estimates frequency from data. Our work extends this idea by attempting to learn a model for both frequency and cost from data. Moreover we explore the statistical optimality of these algorithms in both offline and online settings. We also investigate combining caching algorithms with model multiplexing in order to boost performance.

Acceleration of LLM inference Much effort has been devoted to reducing the cost and latency of LLMs during inference. For example, post-training quantization-based approaches aim to compress the model size by using lower-precision arithmetic without losing too much accuracy [87, 80]. Early-exit frameworks aim to utilize the output in the middle decoder blocks so that only a small fraction of decoder blocks are called when processing a query [19, 230]. The Mixture of Experts approach designs a gating function that only assigns a small fraction of the network for each query [75]. Embedding recycling caches activations from an intermediate layer of a pre-trained model to accelerate the training and inference procedure [66, 274, 219]. LLM cascade starts with the smallest model and continues to call larger models if the output is not acceptable [39]. The big little transformer decoder framework uses a smaller model to generate a draft response and calls the large model to identify the unreliable tokens and perform correction [139]. Similar ideas have been combined with speculative sampling to guarantee that the output remains the same in distribution as that of the large models [38, 156].

5.2 Formulation

We formalize the workflow in Figure 5.1. Consider the set of (finite) prompts / queries $\mathcal{Q} \subset \mathbb{R}^d$. In the *t*-th round, a query $q_t \in \mathcal{Q}$ is sampled from a fixed population distribution $P \in \Delta(\mathcal{Q})$. We maintain a small set of cache $\mathcal{L}_t \subset \mathcal{Q}$ with $|\mathcal{L}_t| \leq L$. We say the query hits the cache if the query satisfies $q_t \in \mathcal{L}_t$. When the query hits the cache, the incurred cost is zero. When the query does not hit the cache, we choose among the existing models to process the query.

In the processing stage, we first describe the setting of caching without model multiplexing, and extend it to the case of caching with model multiplexing.

Caching without model multiplexing

In the case when we only have one model, let $C_l(q)$ denote the random variable of the cost when processing the query with the model. Assume that $C_l(q)$ is supported on $[B_1, B_2]$ with $B_2 > B_1 > 0$ being the upper and lower bounds for the cost. Let $c_l^*(q) = \mathbb{E}[C_l(q)]$ be the expected true cost of processing the query q. The cost for a given query q and cache \mathcal{L} can be written as:

$$\operatorname{cost}(q, \mathcal{L}) = 1(q \notin \mathcal{L})\mathbb{E}[C_l(q)] = 1(q \notin \mathcal{L})c_l^*(q).$$

By taking the expectation over the distribution q, we have the expected cost as

$$\operatorname{cost}(\mathcal{L}) = \sum_{q} P(q) 1(q \notin \mathcal{L}) c_l^{\star}(q).$$

In the offline learning setting, we collect an offline dataset and hope to learn a caching policy $\hat{\mathcal{L}}$ such that $cost(\hat{\mathcal{L}})$ is minimized.

In the online setting, the query comes in a streaming fashion. At the beginning of each round, we receive a query q_t . If the query misses the current cache \mathcal{L}_t , we let the model process the query and receive a cost $c_t \sim \mathbb{P}_{C_t}$. Then we can choose to update the cache \mathcal{L}_t by adding the current query and response to the cache, and replacing one of the existing cached items if the cache \mathcal{L}_t is full. If the query hits the cache $q_t \in \mathcal{L}_t$, then the cost for this round is set to zero with no more observations. In this case, we are interested in characterizing the average difference in the cost throughout the execution of the online learning process. This can be characterized by the regret:

$$\mathsf{Regret}_{\mathsf{cache}}(T) = \sum_{t=1}^{T} \mathbb{E}[\mathsf{cost}(q_t, \mathcal{L}_t) - \mathsf{cost}(q_t, \mathcal{L}^{\star})].$$

Caching with model multiplexing

For the simplicity of the notation, we focus on the case of selecting from a small model and a large model,¹ and discuss how it can be generalized to the case of selecting from multiple models in Appendix D.3. Let $C_s(q)$ denote the random variable of the cost when processing the query with the small model, and $C_l(q)$ denote the random variable of the cost when processing the query with the large model. We assume that both random variables are supported on $[B_1, B_2]$. We observe *i.i.d.* draws of the random variables $C_s(q)$ when executing the small model, and $C_l(q)$ when executing the large model. Denote the expected cost as $c_s^*(q) = \mathbb{E}[C_s(q)]$ and $c_l^*(q) = \mathbb{E}[C_l(q)]$.

Let $\pi : \mathcal{Q} \mapsto [0, 1]$ be the (possibly random) model multiplexing policy that maps the query q to values in [0, 1], where $\pi(q) = 1$ represents that the query is always sent to the small model, and $\pi(q) = 0$ represents the query is always sent to the large model. The randomness in the policy π is independent of the cost $C_s(q), C_l(q)$. The total cost can be written as the following function of the query q, cache \mathcal{L} and policy π :

$$\operatorname{cost}(q, \mathcal{L}, \pi) = 1(q \notin \mathcal{L})\mathbb{E}[C_s(q)\pi(q) + C_l(q)(1 - \pi(q))]$$
$$= 1(q \notin \mathcal{L})(c_s^*(q)\pi(q) + c_l^*(q)(1 - \pi(q))).$$

By taking the expectation over q, we have the expected cost as

$$\operatorname{cost}(\mathcal{L},\pi) = \sum_{q} P(q) \mathbb{1}(q \notin \mathcal{L})(c_s^{\star}(q)\pi(q) + c_l^{\star}(q)(1 - \pi(q))).$$

In the offline learning setting, we collect an offline dataset and hope to learn a caching policy $\hat{\mathcal{L}}$ and a multiplexer $\hat{\pi}$ such that $\operatorname{cost}(\hat{\mathcal{L}}, \hat{\pi})$ is minimized. In the online setting, we get to update the cache in each round by adding the current query into the cache and evicting the ones in the cache if full. When the query q_t misses the cache in round t, we will observe a sample from $C_s(q_t)$ if it is processed by the small model, or a sample from $C_l(q_t)$ if it is processed by the large model. There will be no observations of cost if q_t hits the cache. We aim at minimizing the regret:

$$\mathsf{Regret}_{\mathsf{sel}}(T) = \sum_{t=1}^{T} \mathbb{E}[\mathsf{cost}(q_t, \mathcal{L}_t, \pi_t) - \mathsf{cost}(q_t, \mathcal{L}^{\star}, \pi^{\star})].$$

5.3 Optimal Caching without Model multiplexing

Population setting

We start with the population setting where the probability distribution P and the cost c_l^* are both known. In the case with only one model, the optimal caching strategy is the Least

¹Note that although we name the models as small and large models, we do not impose any assumption on the relationship between their costs. Moreover, the model size and cost function can be arbitrary for both models.

Expected Cost (LEC) or Greedy Dual Size with Frequency (GDSF) algorithm:

$$\mathcal{L}^{\star} = \mathcal{L}_{\mathsf{LEC}} = \operatorname*{arg\ min}_{\mathcal{L}:|\mathcal{L}| \leq L} \mathsf{cost}(\mathcal{L}) = \operatorname*{arg\ min}_{\mathcal{L}:|\mathcal{L}| \leq L} \sum_{q \in \mathcal{Q}} P(q) \mathbb{1}(q \notin \mathcal{L}) c_l^{\star}(q)$$

The traditional frequency-based caching strategy, including Least Recent Used (LRU) and Least Frequently Used (LFU), aims at caching the most frequent queries:

$$\mathcal{L}_{\mathsf{LFU}} = \underset{\mathcal{L}:|\mathcal{L}| \leq L}{\arg\min} \sum_{q \in \mathcal{Q}} P(q) 1(q \notin \mathcal{L}).$$

We show in Appendix D.4 that the ratio between the cost of LFU and LEC can be as high as $\frac{\max_{q \in \mathcal{Q}} c_l^{\star}(q)}{\min_{q \in \mathcal{Q}} c_l^{\star}(q)}$ in the worst case, which shows that LFU can be highly suboptimal when the cost varies significantly.

Finite sample setting: Offline learning

The previous section characterizes the optimal caching strategy in the population setting. We now consider the finite-sample offline learning setting, where we hope to produce a cache \mathcal{L} based on prior data such that the introduced cost is minimized. Denote $\mathcal{D}_N = \{(q_1, c_1), \dots, (q_N, c_N)\}$, where q_i is sampled from the distribution $P(\cdot)$, and c_i is a sample from random variable $C_l(q_i)$. We estimate P, c_l^* from oracles $\hat{P} = \mathsf{DenEstOracle}(q_1, \dots, q_N)$, $\hat{c}_l(q) = \mathsf{RegressionOracle}(\mathcal{D}_N)$. In practice, one may remove the last layer of the pre-trained language model and concatenate it with a linear head and fine-tune the model as the estimator. For theoretical analysis, we focus on the tabular case, where we set both \hat{P} and $\hat{c}_l(q)$ to be the plug-in estimator:

$$\hat{P}(q) = \frac{\sum_{i=1}^{N} 1(q_i = q)}{N},$$
(5.1)

$$\hat{c}_{l}(q) = \begin{cases} \frac{\sum_{i=1}^{N} 1(q_{i}=q)c_{i}}{\sum_{i=1}^{N} 1(q_{i}=q)}, & \text{if } \sum_{i=1}^{N} 1(q_{i}=q) > 0\\ B_{1}, & \text{if } \sum_{i=1}^{N} 1(q_{i}=q) = 0. \end{cases}$$
(5.2)

In practice, the distribution of q may have a long tail. Although the estimation of P(q) is uniformly good for all q, the estimation of $c^*(q)$ can be bad for the queries that are visited less. To select the maximum L elements from the imbalanced samples, we compensate the plug-in estimator by introducing pessimism [212, 131]². As we show in Lemma 45, the true frequency for any query $q \in \mathcal{L}^*$ is lower bounded by some constant that depends on $B_1, B_2, |\mathcal{Q}|$. Thus the pessimism helps eliminate those less visited queries in the long tail of the distribution and encourages caching the queries in \mathcal{L}^* . The lower-confidence-bound based estimator is:

$$\hat{\mathcal{L}} = \underset{\mathcal{L}:|\mathcal{L}|\leq L}{\arg\min} \sum_{q\in\mathcal{Q}} \mathbb{1}(q \notin \mathcal{L})\hat{P}(q) \cdot \max\left(B_1, \left(\hat{c}_l(q) - (B_2 - B_1)\sqrt{\frac{\log(6N|\mathcal{Q}|/\delta)}{2\sum_{n=1}^N \mathbb{1}(q_n = q)}}\right)\right)$$

²If we impose a uniform lower bound on the probability P(q), then the pessimism can be replaced with the plug-in estimator. However, it is usually not the case in practice since P(q) usually comes with a long tail.

We show how the cost for the caching from the empirical estimate differs from the optimal cost.

Theorem 31. Assume that $N \geq \frac{8B_2|\mathcal{Q}|\log(3L/\delta)}{B_1}$ and taking $\delta = 1/N$. We have

$$\mathbb{E}[\operatorname{cost}(\hat{\mathcal{L}}) - \operatorname{cost}(\mathcal{L}^{\star})] \le C(B_2 - B_1)L \cdot \sqrt{\frac{B_2|\mathcal{Q}|\log(N|\mathcal{Q}|)}{NB_1}}.$$

The proof is deferred to Appendix D.5, where we prove a stronger high-probability bound rather than a bound in expectation. From the theorem, we know that the cost of the finite-sample caching policy converges to the cost of the optimal policy at a rate of $1/\sqrt{N}$, which achieves the optimal dependence on N. The insights from the tabular case also indicate that the cost needs to be estimated in a conservative fashion when considered for the cache replacement algorithm.

Finite sample setting: Online learning

We summarize the caching algorithm pipeline in 3, which relies on the two estimation oracles, DenEstOracle and RegressionOracle, which estimate both the frequency and cost of models from data.

Algorithm 3 Caching in Online Learning

1: Initialize the set of cache $\mathcal{L}_1 = \{\}$, past observations $\mathcal{H}_1 = \{\}, \hat{c}_{l,0}(q) = B_1, \forall q \in \mathcal{Q}.$

- 2: For iteration $t = 1, 2 \cdots, T$
- 3: Receive query q_t .
- 4: Update the density estimation $\hat{P}_t = \mathsf{DenEstOracle}(q_1, \cdots, q_t)$.
- 5: If $q_t \in \mathcal{L}_t$:
- 6: Output the cached result, set $\hat{c}_{l,t} = \hat{c}_{l,t-1}$, update the past observation $\mathcal{H}_t = \mathcal{H}_{t-1} \bigcup (q_t, \times)$, and continue.
- 7: Use the large model to process the query, and observe a cost $c_t \sim \mathbb{P}_{C_l(q)}$.
- 8: Update the past observation $\mathcal{H}_t = \mathcal{H}_{t-1} \bigcup (q_t, c_t)$.
- 9: Update $\hat{c}_{l,t} = \mathsf{RegressionOracle}(\mathcal{H}_t)$.
- 10: If $|\mathcal{L}_t| < L$:
- 11: Let \mathcal{L}_{t+1} be the union of \mathcal{L}_t and q_t .
- 12: Else if $\dot{P}_t(q_t) \cdot \hat{c}_{l,t}(q_t) > \min_{q \in \mathcal{L}_t} \dot{P}_t(q) \cdot \hat{c}_{l,t}(q)$:
- 13: Replace the minimizer element of $\hat{P}_t(q) \cdot \hat{c}_{l,t}(q)$ in the cache \mathcal{L}_t with q_t to get \mathcal{L}_{t+1} .
For theoretical analysis, we focus on the tabular case and define the oracles as follows:

$$\hat{P}_{t}(q) = \frac{\sum_{i=1}^{t} 1(q_{i} = q)}{t},$$

$$\hat{c}_{l,t}(q) = \begin{cases} B_{1}, & \text{if } \sum_{i=1}^{t} 1(c_{i} \neq \times, q_{i} = q) = 0, \\ \max\left(P_{1} \sum_{i=1}^{t} 1(c_{i} \neq \times, q_{i} = q) = 0, \\ \max\left(P_{1} \sum_{i=1}^{t} 1(c_{i} \neq \times, q_{i} = q) = 0, \\ \sum_{i=1}^{t} 1(c_{i} \neq \times, q_{i} = q) = 0, \end{cases}$$
(5.3)

$$\left(\max\left(B_{1}, \frac{\sum_{i=1}^{t} 1(c_{i} \neq \times, q_{i} = q)}{\sum_{i=1}^{t} 1(c_{i} \neq \times, q_{i} = q)} - (B_{2} - B_{1})\sqrt{\frac{\log(01|Q|/\delta)}{2\sum_{i=1}^{t} 1(c_{i} \neq \times, q_{i} = q)}}\right), \quad \text{otherwise}$$
(5.4)

For the estimation of density, we use plug-in estimator since there is no imbalance in the sampling process. For the estimation of the cost, we subtract the confidence bound to include pessimism. We have the following regret guarantee.

Theorem 32. When substituting the DenEstOracle and RegressionOracle with Equation (5.3) and (5.4) and set $\delta = 1/T$, we have for some universal constant C:

$$\operatorname{Regret}_{\operatorname{cache}}(T) \leq \frac{CL(B_2 - B_1)B_2|\mathcal{Q}|L\log^2(T|\mathcal{Q}|)}{B_1} \cdot \sqrt{T}$$

On the other hand, for any caching policy $\{\mathcal{L}_t\}_{t=1}^T$, there exist some cases of $P(q), c_l^*(q)$ such that for some universal constant C',

$$\operatorname{Regret}_{\operatorname{cache}}(T) \ge C'\sqrt{T}.$$

The proof is deferred to Appendix D.6. Different from the offline case, one interesting feature of the online case is the *partial observation phenomenon*: when the query hits the cache, it will not be processed by the model, and thus we cannot observe the sample from $C_l(q)$ in this round. This is different from the traditional bandit literature where the selected arm is always observed in each round. Thus the partial observation thus requires new upper and lower bound analysis.

5.4 Optimal Caching and Model multiplexing

Population setting

In the case when we have access to two models, we need to design a good caching and model multiplexing strategy jointly. We can compute the optimal caching and model multiplexing policy as $\mathcal{L}^*, \pi^* = \arg \min_{\mathcal{L},\pi} \operatorname{cost}(\mathcal{L}, \pi)$, which gives the following solution:

$$\pi^{\star}(q) = 1(c_{s}^{\star}(q) \leq c_{l}^{\star}(q)),$$
$$\mathcal{L}^{\star} = \underset{\mathcal{L}:|\mathcal{L}|\leq L}{\arg\min} \sum_{q\in\mathcal{Q}} P(q)1(q \notin \mathcal{L})\min\left(c_{s}^{\star}(q), c_{l}^{\star}(q)\right).$$

Such optimal strategies are straightforward: π^* always assigns the query to the model with a smaller cost, and \mathcal{L}^* saves the *L* queries with the largest $P(q) \cdot \min(c_s^*(q), c_l^*(q))$.

For the model multiplexing algorithm, we consider two baselines: (a) one always uses large model $\pi_l(q) \equiv 0$; (b) one always uses the small model $\pi_s(q) \equiv 0$. This is related to the LLM cascade idea in the concurrent work of [39]. We provide more discussion in Appendix D.1, and present comparisons between baselines and π^* in Appendix D.4.

Finite sample setting: Offline learning

We now consider the finite sample case. Let $\mathcal{D}_N = \{(q_1, c_{s,1}, c_{l,1}), \cdots, (q_N, c_{s,N}, c_{l,N})\}$, where $c_{s,n}$ is a sample from random variable $C_s(q_n)$, the observed cost for processing query q_n with the small model in round n. And $c_{l,n}$ is a sample from random variable $C_l(q_n)$, the observed cost for processing query q_n with the large model in round n. We consider estimating P, c_s^*, c_t^* with some oracles $\hat{P} = \mathsf{DenEstOracle}(q_1, \cdots, q_N), \hat{c}_s(q), \hat{c}_t(q) = \mathsf{RegressionOracle}(\mathcal{D}_N)$. We focus on the tabular case for theoretical analysis, where we set $\hat{P}, \hat{c}_s(q)$ and $\hat{c}_l(q)$ to be the plug-in estimator:

$$\hat{P}(q) = \frac{\sum_{i=1}^{N} 1(q_i = q)}{N}, \hat{c}_l(q) = \begin{cases} \frac{\sum_{i=1}^{N} 1(q_i = q)c_{l,i}}{\sum_{i=1}^{N} 1(q_i = q)}, & \text{if } \sum_{i=1}^{N} 1(q_i = q) > 0\\ B_1, & \text{if } \sum_{i=1}^{N} 1(q_i = q) = 0, \end{cases}$$
$$\hat{c}_s(q) = \begin{cases} \frac{\sum_{i=1}^{N} 1(q_i = q)c_{s,i}}{\sum_{i=1}^{N} 1(q_i = q)}, & \text{if } \sum_{i=1}^{N} 1(q_i = q) > 0\\ B_1, & \text{if } \sum_{i=1}^{N} 1(q_i = q) > 0\\ B_1, & \text{if } \sum_{i=1}^{N} 1(q_i = q) = 0. \end{cases}$$

Similar to the case of caching without model multiplexing, for a long-tailed distribution P(q), the estimation of $c_s^*(q), c_l^*(q)$ can be bad for the queries that are visited less. To select the maximum L elements from the plug-in estimator, we introduce pessimism to the estimate of \hat{c}_l and \hat{c}_s . This leads to the following design of caching and model multiplexer \hat{L} and $\hat{\pi}$:

$$\hat{\pi}(q) = 1(\hat{c}_s(q) \le \hat{c}_l(q)),$$
$$\hat{\mathcal{L}} = \underset{\mathcal{L}:|\mathcal{L}|\le L}{\arg\min} \sum_{q\in\mathcal{Q}} 1(q \notin \mathcal{L})\hat{P}(q) \max\left(B_1, \min(\hat{c}_s(q), \hat{c}_l(q)) - (B_2 - B_1)\sqrt{\frac{\log(8|\mathcal{Q}|/\delta)}{2\sum_{n=1}^N 1(q_n = q)}}\right).$$

We now show the cost for the caching and model multiplexer obtained from the empirical estimate is close to the optimal cost. The proof is deferred to Appendix D.7.

Theorem 33. Assume that $N \geq \frac{8B_2|\mathcal{Q}|\log(4L/\delta)}{B_1}$ and take $\delta = 1/N$. We have

$$\mathbb{E}[\operatorname{cost}(\hat{\mathcal{L}}, \hat{\pi}) - \operatorname{cost}(\mathcal{L}^{\star}, \pi^{\star})] \leq CL(B_2 - B_1) \cdot \sqrt{\frac{B_2|\mathcal{Q}|\log(8|\mathcal{Q}|N)}{B_1N}}$$

Algorithm 4 Joint Design of Caching and Model multiplexing

- 1: Initialize the set of cache $\mathcal{L}_1 = \{\}$, past observations $\mathcal{H}_1 = \{\}$, $\hat{c}_{l,0}(q) = B_1$, $\hat{c}_{s,0}(q) = B_1$, $\hat{c}_{s,0}(q) = B_1$, model multiplexing policy $\pi_0(q) = 1, \forall q \in \mathcal{Q}$.
- 2: For iteration $t = 1, 2 \cdots, T$
- 3: Receive query q_t .
- 4: Update the density estimation $\hat{P}_t = \mathsf{DenEstOracle}(q_1, \cdots, q_t)$.
- 5: If $q_t \in \mathcal{L}_t$: output the cached result, set $\hat{c}_{s,t} = \hat{c}_{s,t-1}, \hat{c}_{l,t} = \hat{c}_{l,t-1}, \pi_t = \pi_{t-1}$, update the past observation $\mathcal{H}_t = \mathcal{H}_{t-1} \bigcup (q_t, \times, \times)$, and continue.
- 6: Select the models according to $s_t = \pi_t(q_t)$.
- 7: Update the past observation $\mathcal{H}_t = \mathcal{H}_{t-1} \bigcup (q_t, s_t, c_t).$
- 8: Update $\hat{c}_{l,t}, \hat{c}_{s,t} = \mathsf{RegressionOracle}(\mathcal{H}_t)$. Set $\pi_{t+1}(q) = 1(\hat{c}_{s,t}(q) < \hat{c}_{l,t}(q))$.
- 9: If $|\mathcal{L}_t| < L$: let \mathcal{L}_{t+1} be the union of \mathcal{L}_t and q_t .
- 10: Else if $P_t(q_t) \cdot \min(\hat{c}_{s,t}(q_t), \hat{c}_{l,t}(q_t)) > \min_{q \in \mathcal{L}_t} P_t(q) \cdot \min(\hat{c}_{s,t}(q), \hat{c}_{l,t}(q))$: replace the minimizer element in the cache \mathcal{L}_t on the RHS with q_t to get \mathcal{L}_{t+1} .

Finite sample setting: Online learning

We turn to the online case. We first propose a meta-algorithm in Algorithm 4. We provide a theoretical analysis of the meta-algorithm for the tabular case, with DenEstOracle $\hat{P}_t(q) = \frac{\sum_{i=1}^{t} 1(q_i=q)}{t}$, and the RegressionOracle defined as follows:

$$\hat{c}_{l,t}(q) = \begin{cases} B_1, & \text{if } \sum_{i=1}^t 1(s_i = 0, q_i = q) = 0\\ \max\left(B_1, \frac{\sum_{i=1}^t 1(s_i = 0, q_i = q)c_{l,i}}{\sum_{i=1}^t 1(s_i = 0, q_i = q)} - (B_2 - B_1)\sqrt{\frac{\log(8T|Q|/\delta)}{2\sum_{i=1}^t 1(s_i = 0, q_i = q)}}\right), \text{ otherwise,} \\ \hat{c}_{s,t}(q) = \begin{cases} B_1, & \text{if } \sum_{i=1}^t 1(s_i = 1, q_i = q) = 0,\\ \max\left(B_1, \frac{\sum_{i=1}^t 1(s_i = 1, q_i = q)c_{s,i}}{\sum_{i=1}^t 1(s_i = 1, q_i = q)} - (B_2 - B_1)\sqrt{\frac{\log(8T|Q|/\delta)}{2\sum_{i=1}^t 1(s_i = 1, q_i = q)}}\right), \text{ otherwise.} \end{cases}$$

We provide the following theorem on the regret of the overall algorithm.

Theorem 34. Substituting the oracles in Algorithm 4 with the oracles above and $\delta = 1/T$, we have

$$\operatorname{Regret}_{\operatorname{sel}}(T) \leq \frac{CL^2(B_2 - B_1)B_2|\mathcal{Q}|L\log^2(T|\mathcal{Q}|)}{B_1} \cdot \sqrt{T}.$$

The proof is deferred to Appendix D.8. Compared with the lower bound in Theorem 32, we see that the dependency on T is tight. The pessimism plays two different roles here: on the one hand, it encourages the exploration for model multiplexing to choose the ones with more uncertainty in the cost; on the other hand, it encourages the exploitation to be conservative about which query to save into the cache.

For the model multiplexer to work well, one needs to have a small yet accurate model multiplexer. In the case when the model multiplexer is not accurate, the small model always comes with a much smaller cost, and we are allowed to regenerate the responses and make corrections for the output, one may combine LEC with cascade [39] to achieve better performance.

5.5 Experiments

We conduct both simulations and real-world experiments with our proposed methods. The code is available at https://github.com/Ying1123/llm-caching-multiplexing.

Simulations for algorithm analysis

We conduct synthetic online and offline experiments for joint optimization of caching and model switching. In Figure 5.2, we plot the cumulative cost and regret in online learning for LFU and LEC caching algorithms. For LFU, we consider model switchers which always select the small or large models as the baselines. We consider 20 distinct prompts and set the cache size to be 10. We set the frequency distribution as power distribution with $\alpha = 0.9$. The ground truth cost for each query processed by both models is set as a sample from 100X + 1, where X is a random variable generated from a Bernoulli distribution with the parameter 0.5. We repeat the simulation 100 times and plot the mean and standard deviation in the figure. Our simulation suggests that LEC with model switcher greatly improves the two baselines by a factor of $50 \times$ when the cost ratio is 100. We include additional results on the synthetic datasets for both online and offline settings with different α values, cost ratios, and switcher accuracy in Appendix D.9.



Figure 5.2. Comparisons between LFU with either small or large model switching and LEC with model switcher. Both the x-axis and y-axis are logarithmic scales. The shaded regime represents the standard deviation calculated from the repeated experiments.

Experiments on real datasets

We evaluate our algorithms on two tasks: next-token prediction on the Lambada [193] dataset and chat assistant on the OpenAssistant [143] dataset.

For the next-token prediction task, we run the offline algorithm with two models: OPT-1.3B and OPT-13B [305] and use FLOPs as the cost. The target performance metric is the number of correct tokens predicted, where we get the ground-truth token from the Lambada dataset. For a given query, an algorithm can choose to run the small model or the large model. If the small model is chosen but its result is wrong, the large model must be run and it will incur an additional penalty. We fine-tune a BERT base model with 2000 samples as the model switcher by predicting whether the small model can give the correct result and achieve 80.2% accuracy. We work with 100 unseen distinct prompts in the offline setting with total queries 10000 and cache size 40. We compare our offline caching and switcher algorithms against LFU, large-model-only, and cascade (which always calls the small model first). As shown in Table 5.1, LEC is better than LFU in all cases. Combining LEC and switcher brings up to $4.3 \times$ cost reduction compared to the baseline "LFU + Large." However, as the predictor accuracy is limited, the model switcher may not be as good as the cascade algorithm in some cases. We leave the training of a better switcher as future work.

On the chat assistant task, we run the online algorithm with two models: FastChat-T5-3B and Vicuna-13B [49], and use the inference latency as the cost. The quality of response is evaluated by GPT4 evaluation [170]. We say a response is satisfying if the score is larger than 6 out of 10, and unsatisfying otherwise. If the response from the small model is unsatisfying, we will call the large model again and incur an additional cost in latency. The ratio between the average latency of the large model and the small model is 1.85. We work with 100 distinct prompts in the online setting with total queries 10000 and cache size 40. After a sufficient number of online learning steps, the switcher learns the accurate costs of two models on this finite prompts set, so "LEC + switcher" outperforms other algorithms in all cases on Table 5.2 with up to $1.8 \times$ latency reduction compared to "LFU + large" baseline.

α	selector accu- racy	LFU+ large	LFU+ cas- cade	LFU+ selec- tor	LEC+ large	LEC+ cas- cade	LEC+ selec- tor
0.2 0.8	$rac{80\%}{80\%}$	$3.49 \\ 10.81$	$3.81 \\ 11.80$	$2.60 \\ 8.06$	$3.44 \\ 10.36$	$\begin{array}{c} 1.50\\ 4.11\end{array}$	$2.00 \\ 4.76$
0.2 0.8	$100\% \\ 100\%$	3.49 10.81	3.81 11.80	1.91 5.90	$3.44 \\ 10.36$	1.50 4.11	$\begin{array}{c} 0.99\\ 2.50\end{array}$

Table 5.1. Evaluation of offline algorithms on the Lambada dataset with OPT-1.3B and OPT-13B, 100 distinct prompts, total query size 10000 and cache size 40. α is the parameter of the power distribution of the prompts. The table lists cumulative costs (10³) for different algorithms.

α	LFU+ large	LFU+ cas- cade	LFU+ selec- tor	LEC+ large	LEC+ cas- cade	LEC+ selec- tor
0.2	9.31	13.88	7.24	8.74	8.82	5.93
0.5 0.8	20.04 28.24	42.12	15.11 21.14	18.08 26.07	10.90 20.31	11.87 15.49

Table 5.2. Evaluation of online algorithms on the OpenAssistant dataset with FastChat-T5-3B and Vicuna-13B, 100 distinct prompts, total query size 10000 and cache size 40. α is the parameter of the power distribution of the prompts. The table lists cumulative costs (10³) for different algorithms.

We provide more experiments in Appendix D.9, where we evaluate both FLOPs and latency for both offline and online setting on both synthetic and real dataset, with varying cache size, query size and distinct prompts.

5.6 Conclusions

We have studied the joint optimization of caching and model multiplexing and proposed an optimal algorithm for the tabular case. There are a variety of further work that can be pursued in this vein, including:

- Designing the optimal caching and model multiplexing algorithm when there is a query queue, such that the query arrives at a random interval rather than a fixed interval. A more complicated serving pattern also needs to take batching strategies into consideration.
- Understanding the scaling law of the predictors. We hope to use a small yet accurate model for prediction to reduce overhead introduced by the predictor. It is important to understand the trade-off between prediction accuracy, model size, and training data size.
- Designing optimal caching algorithm when the responses generated in each round have diverse qualities.

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Appendix A

Appendix for Theoretical Analysis of RLHF

A.1 Analysis for nonlinear r_{θ}

Consider the case of pairwise comparison when r_{θ} is not linear, the MLE can be written as

$$\hat{\theta}_{\mathsf{MLE}} \in \underset{\theta \in \Theta_B}{\arg\min} \ell_{\mathcal{D}}(\theta),$$

where $\ell_{\mathcal{D}}(\theta) = -\sum_{i=1}^{n} \log \left(1(y^i = 1) \cdot \frac{\exp(r_{\theta}(s^i, a_1^i))}{\exp(r_{\theta}(s^i, a_0^i)) + \exp(r_{\theta}(s^i, a_1^i))} + 1(y^i = 0) \cdot \frac{\exp(r_{\theta}(s^i, a_0^i))}{\exp(r_{\theta}(s^i, a_0^i)) + \exp(r_{\theta}(s^i, a_1^i))} \right).$

Here we provide a guarantee for the case when r_{θ} is nonlinear and non-convex. We first make the following boundedness and smoothness assumption on r_{θ} :

Assumption 35. Assume that for any $\theta \in \Theta_B$, $s \in \mathcal{S}$, $a_0 \in \mathcal{A}$, $a_1 \in \mathcal{A}$ with $a_0 \neq a_1$, we have,

$$\begin{aligned} |r_{\theta}(s,a)| &\leq \alpha_{0}, \quad \text{(Bounded value)} \\ \|\nabla r_{\theta}(s,a)\|_{2} &\leq \alpha_{1}, \quad \text{(Bounded gradient)} \\ \|\nabla^{2} r_{\theta}(s,a)\|_{2} &\leq \alpha_{2}. \quad \text{(Bounded Hessian / Lipschitz gradient)} \end{aligned}$$

One can verify that our linear reward satisfies the above assumption with $\alpha_0 = LB$, $\alpha_1 = L$, $\alpha_2 = 0$. Under this assumption, we have

Theorem 36. For any $\lambda > 0$, with probability at least $1 - \delta$,

$$\|\hat{\theta}_{\mathsf{MLE}} - \theta^{\star}\|_{\Sigma_{\mathcal{D}} + \lambda I} \le C \cdot \sqrt{\frac{d + \log(1/\delta)}{\gamma^2 n} + (\lambda + \alpha_2/\gamma + \alpha_1 \alpha_2 B)B^2}.$$

Here $\gamma = \frac{1}{2 + \exp(-2\alpha_0) + \exp(2\alpha_0)}, \Sigma_{\mathcal{D}} = \frac{1}{n} \sum_{i=1}^n \nabla (r_{\theta^*}(s^i, a_1^i) - r_{\theta^*}(s^i, a_0^i)) \nabla (r_{\theta^*}(s^i, a_1^i) - r_{\theta^*}(s^i, a_0^i))^\top.$

The proof is deferred to Appendix A.2. Our result recovers Lemma 6 when $\alpha_2 = 0$ and reveals how the gradient of r plays a role in the bound for estimation error. However, the dependence on α_2 will not vanish as $n \to \infty$. It remains open how to get vanishing rate for nonlinear reward functions when $\alpha_2 > 0$. Similar argument can also be applied to the case of K-wise comparison and MDP. And we can similarly design pessimistic MLE based on the confidence bound on $\hat{\theta}_{MLE}$.

On the other hand, we can show that the true parameter θ^* is a global minimum of the population negative log likelihood even when r_{θ} is nonlinear and we use MLE_2 for K-wise comparison. Recall that the MLE_2 splits K-wise comparisons into pairwise comparisons, and is given by

$$\hat{\theta}_{\mathsf{MLE}_2} \in \underset{\theta \in \Theta_B}{\arg\min} \, \ell_{\mathcal{D}}(\theta),$$

where $\ell_{\mathcal{D}}(\theta) = -\frac{1}{n} \sum_{i=1}^n \sum_{j=0}^{K-1} \sum_{k=j+1}^{K-1} \log \left(\frac{\exp(r_{\theta}(s^i, a^i_{\sigma_i(j)}))}{\exp(r_{\theta}(s^i, a^i_{\sigma_i(j)})) + \exp(r_{\theta}(s^i, a^i_{\sigma_i(k)}))} \right)$

When there is infinite number of data, the loss become

$$\mathbb{E}[\ell(\theta)] = -\sum_{s} \rho(s) \sum_{a_{0},a_{1}\in\mathcal{A}} \rho(a_{0},a_{1} \mid s) \\ \cdot \left(\frac{\exp(r_{\theta^{\star}}(s,a_{0}))}{\exp(r_{\theta^{\star}}(s,a_{0})) + \exp(r_{\theta^{\star}}(s,a_{1}))} \log\left(\frac{\exp(r_{\theta}(s,a_{0}))}{\exp(r_{\theta}(s,a_{0})) + \exp(r_{\theta}(s,a_{1}))}\right) \\ + \frac{\exp(r_{\theta^{\star}}(s,a_{0})) + \exp(r_{\theta^{\star}}(s,a_{1}))}{\exp(r_{\theta^{\star}}(s,a_{0})) + \exp(r_{\theta^{\star}}(s,a_{1}))} \log\left(\frac{\exp(r_{\theta}(s,a_{0})) + \exp(r_{\theta}(s,a_{1}))}{\exp(r_{\theta}(s,a_{0})) + \exp(r_{\theta}(s,a_{1}))}\right)\right).$$

Here $\rho(a_0, a_1 \mid s)$ is the probability that actions a_0, a_1 are included in the K-comparison when the state is s. Now we show

$$\theta^{\star} \in \arg\min_{\theta} \mathbb{E}[\ell(\theta)]. \tag{A.1}$$

To see this, note that we have

$$\begin{split} \mathbb{E}[\ell(\theta)] &= -\sum_{s} \rho(s) \sum_{a_{0},a_{1} \in \mathcal{A}} \rho(a_{0},a_{1} \mid s) \\ & \cdot \Big(\frac{\exp(r_{\theta^{\star}}(s,a_{0}))}{\exp(r_{\theta^{\star}}(s,a_{0})) + \exp(r_{\theta^{\star}}(s,a_{1}))} \log\Big(\frac{\exp(r_{\theta}(s,a_{0}))}{\exp(r_{\theta}(s,a_{0})) + \exp(r_{\theta}(s,a_{1}))} \Big) \\ & + \frac{\exp(r_{\theta^{\star}}(s,a_{1}))}{\exp(r_{\theta^{\star}}(s,a_{0})) + \exp(r_{\theta^{\star}}(s,a_{1}))} \log\Big(\frac{\exp(r_{\theta}(s,a_{0})) + \exp(r_{\theta}(s,a_{1}))}{\exp(r_{\theta}(s,a_{0})) + \exp(r_{\theta}(s,a_{1}))} \Big) \Big) \\ &= \sum_{s} \rho(s) \sum_{a_{0},a_{1} \in \mathcal{A}} p(a_{0},a_{1} \mid s) \cdot (H(p_{\theta^{\star}}(s,a_{0},a_{1})) + \mathsf{KL}(p_{\theta^{\star}}(s,a_{0},a_{1}) \| p_{\theta}(s,a_{0},a_{1}))). \end{split}$$

Here $H(p) = p \log(1/p) + (1-p) \log(1/(1-p))$ is the entropy of a Bernoulli distribution with parameter p. And $p_{\theta}(s, a_0, a_1) = \frac{\exp(r_{\theta}(s, a_1))}{\exp(r_{\theta}(s, a_0)) + \exp(r_{\theta}(s, a_1))}$. Now note that KL is lower bounded by 0, with equality when $\theta = \theta^*$. This proves Equation (A.1).

A.2 Remaining Proofs

Proof of Lemma 6

Recall that the MLE is given by

$$\begin{split} \hat{\theta}_{\mathsf{MLE}} &\in \mathop{\arg\min}_{\theta \in \Theta_B} \ell_{\mathcal{D}}(\theta), \\ \text{where } \ell_{\mathcal{D}}(\theta) &= -\sum_{i=1}^{n} \log \left(1(y^{i} = 1) \cdot \frac{\exp(r_{\theta}(s^{i}, a^{i}_{0}))}{\exp(r_{\theta}(s^{i}, a^{i}_{0})) + \exp(r_{\theta}(s^{i}, a^{i}_{1}))} \right) \\ &+ 1(y^{i} = 0) \cdot \frac{\exp(r_{\theta}(s^{i}, a^{i}_{0}))}{\exp(r_{\theta}(s^{i}, a^{i}_{0})) + \exp(r_{\theta}(s^{i}, a^{i}_{1}))} \\ &= -\sum_{i=1}^{n} \log \left(1(y^{i} = 1) \cdot \frac{1}{1 + \exp(r_{\theta}(s^{i}, a^{i}_{0}) - r_{\theta}(s^{i}, a^{i}_{1}))} \right) \\ &+ 1(y^{i} = 0) \cdot \left(1 - \frac{1}{1 + \exp(r_{\theta}(s^{i}, a^{i}_{0}) - r_{\theta}(s^{i}, a^{i}_{1}))} \right) \right) \\ &= -\sum_{i=1}^{n} \log \left(1(y^{i} = 1) \cdot \frac{1}{1 + \exp(\theta^{\top}(\phi(s^{i}, a^{i}_{0}) - \phi(s^{i}, a^{i}_{1})))} \right) \\ &+ 1(y^{i} = 0) \cdot \left(1 - \frac{1}{1 + \exp(\theta^{\top}(\phi(s^{i}, a^{i}_{0}) - \phi(s^{i}, a^{i}_{1})))} \right) \right) \end{split}$$

To simplify the notation, we let $x_i = \phi(s^i, a_1^i) - \phi(s^i, a_0^i)$. Our goal is to bound the estimation error of the MLE in the squared semi-norm $\|v\|_{\Sigma_{\mathcal{D}}+\lambda I}^2 = v^{\top}(\Sigma_{\mathcal{D}}+\lambda I)v$.

Strong convexity of ℓ . We first show that $\ell_{\mathcal{D}}$ is strongly convex at θ^* with respect to the semi-norm $\|\cdot\|_{\Sigma_{\mathcal{D}}}$, meaning that there is some constant $\gamma > 0$ such that

$$\ell_{\mathcal{D}}(\theta^{\star} + \Delta) - \ell_{\mathcal{D}}(\theta^{\star}) - \langle \nabla \ell_{\mathcal{D}}(\theta^{\star}), \Delta \rangle \ge \gamma \|\Delta\|_{\Sigma_{\mathcal{D}}}^2$$
(A.2)

for all perturbations $\Delta \in \mathbb{R}^d$ such that $\theta^* + \Delta \in \Theta_B$.

One can directly calculate the Hessian of ℓ as

$$\nabla^2 \ell_{\mathcal{D}}(\theta) = \frac{1}{n} \sum_{i=1}^n \left(1(y^i = 1) \cdot \frac{\exp(-\langle \theta, x_i \rangle)}{(\exp(-\langle \theta, x_i \rangle) + 1)^2} + 1(y^i = 0) \cdot \frac{\exp(\langle \theta, x_i \rangle)}{(\exp(\langle \theta, x_i \rangle) + 1)^2} \right) \cdot x_i x_i^\top$$
$$= \frac{1}{n} \sum_{i=1}^n \frac{\exp(-\langle \theta, x_i \rangle)}{(\exp(-\langle \theta, x_i \rangle) + 1)^2} \cdot x_i x_i^\top$$

Observe that $\langle \theta, x_i \rangle \in [-2LB, 2LB]$, which gives that

$$\frac{\exp(-\langle \theta, x_i \rangle)}{(\exp(-\langle \theta, x_i \rangle) + 1)^2} \ge \frac{1}{2 + \exp(-2LB) + \exp(2LB)}.$$

Putting together the pieces, we conclude that

$$v^{\top} \nabla^2 \ell_{\mathcal{D}}(\theta) v \ge \frac{\gamma}{n} \|Xv\|_2^2 \quad \text{for all } v,$$

where $\gamma = 1/(2 + \exp(-2LB) + \exp(2LB))$, $X \in \mathbb{R}^{n \times d}$ has the differencing vector $x_i \in \mathbb{R}^d$ as its i^{th} row. Thus, if we introduce the error vector $\Delta := \hat{\theta}_{\mathsf{MLE}} - \theta^*$, then we may conclude that

$$\ell_{\mathcal{D}}(\theta^{\star} + \Delta) - \ell_{\mathcal{D}}(\theta^{\star}) - \langle \nabla \ell_{\mathcal{D}}(\theta^{\star}), \Delta \rangle \ge \frac{\gamma}{n} \| X \Delta \|_{2}^{2} = \gamma \| \Delta \|_{\Sigma_{\mathcal{D}}}^{2},$$

showing that $\ell_{\mathcal{D}}$ is strongly convex around θ^* with parameter γ .

Bounding the estimation error. Now we aim at bounding the estimation error $\|\hat{\theta}_{\mathsf{MLE}} - \theta^*\|_{\Sigma_{\mathcal{D}}}$. Since $\hat{\theta}_{\mathsf{MLE}}$ is optimal for $\ell_{\mathcal{D}}$, we have $\ell_{\mathcal{D}}(\hat{\theta}_{\mathsf{MLE}}) \leq \ell_{\mathcal{D}}(\theta^*)$. (When $\hat{\theta}_{\mathsf{MLE}}$ is approximately optimal, i.e. $\ell_{\mathcal{D}}(\hat{\theta}_{\mathsf{MLE}}) \leq \min_{\theta} \ell_{\mathcal{D}}(\theta) + \epsilon$, the same argument also holds up to an extra additive term ϵ .) Defining the error vector $\Delta = \hat{\theta}_{\mathsf{MLE}} - \theta^*$, adding and subtracting the quantity $\langle \nabla \ell_{\mathcal{D}}(\theta^*), \Delta \rangle$ yields the bound

$$\ell_{\mathcal{D}}(\theta^{\star} + \Delta) - \ell_{\mathcal{D}}(\theta^{\star}) - \langle \nabla \ell_{\mathcal{D}}(\theta^{\star}), \Delta \rangle \leq - \langle \nabla \ell_{\mathcal{D}}(\theta^{\star}), \Delta \rangle.$$

By the γ -convexity condition, the left-hand side is lower bounded by $\gamma \|\Delta\|_{\Sigma_{\mathcal{D}}}^2$. As for the right-hand side, note that $|\langle \nabla \ell_{\mathcal{D}}(\theta^*), \Delta \rangle| \leq \|\nabla \ell_{\mathcal{D}}(\theta^*)\|_{(\Sigma_{\mathcal{D}}+\lambda I)^{-1}} \|\Delta\|_{\Sigma_{\mathcal{D}}+\lambda I}$ for any $\lambda > 0$. Altogether we have

$$\gamma \|\Delta\|_{\Sigma_{\mathcal{D}}}^2 \le \|\nabla \ell_{\mathcal{D}}(\theta^\star)\|_{(\Sigma_{\mathcal{D}}+\lambda I)^{-1}} \|\Delta\|_{\Sigma_{\mathcal{D}}+\lambda I}.$$

Now we further bound the term $\|\nabla \ell_{\mathcal{D}}(\theta^*)\|_{(\Sigma_{\mathcal{D}}+\lambda I)^{-1}}$. Observe that the gradient takes the form

$$\nabla \ell_{\mathcal{D}}(\theta^{\star}) = \frac{-1}{n} \sum_{i=1}^{n} \left[1[y^i = 1] \frac{\exp(-\langle \theta^{\star}, x_i \rangle)}{1 + \exp(-\langle \theta^{\star}, x_i \rangle))} - 1[y^i = 0] \frac{1}{1 + \exp(-\langle \theta^{\star}, x_i \rangle))} \right] x_i.$$

Define a random vector $V \in \mathbb{R}^n$ with independent components as

$$V_i = \begin{cases} \frac{\exp(-\langle \theta^{\star}, x_i \rangle)}{1 + \exp(-\langle \theta^{\star}, x_i \rangle))} & \text{w.p.} & \frac{1}{1 + \exp(-\langle \theta^{\star}, x_i \rangle))} \\ \frac{-1}{1 + \exp(-\langle \theta^{\star}, x_i \rangle))} & \text{w.p.} & \frac{\exp(-\langle \theta^{\star}, x_i \rangle)}{1 + \exp(-\langle \theta^{\star}, x_i \rangle))} \end{cases}$$

With this notation, we have $\nabla \ell_{\mathcal{D}}(\theta^{\star}) = -\frac{1}{n} X^{\top} V$. One can verify that $\mathbb{E}[V] = 0$ and $|V_i| \leq 1$. Define the *n*-dimensional matrix $M \coloneqq \frac{1}{n^2} X(\Sigma_{\mathcal{D}} + \lambda I)^{-1} X^{\top}$, we have $\|\nabla \ell_{\mathcal{D}}(\theta^{\star})\|_{(\Sigma_{\mathcal{D}} + \lambda I)^{-1}}^2 = -\frac{1}{n^2} \|\nabla \ell_{\mathcal{D}}(\theta^{\star})\|_{(\Sigma_{\mathcal{D}} + \lambda I)^{-1}}^2$

 $V^{\top}MV$. Let the eigenvalue decomposition of $X^{\top}X$ be $X^{\top}X = U\Lambda U^{\top}$. We can bound the trace and operator norm of M as

$$\begin{aligned} \mathrm{Tr}(M) &= \frac{1}{n^2} \mathrm{Tr}(U(\Lambda/n + \lambda I)^{-1} U^{\top} U \Lambda U^{\top}) \leq \frac{d}{n} \\ \mathrm{Tr}(M^2) &= \frac{1}{n^4} \mathrm{Tr}(U(\Lambda/n + \lambda I)^{-1} U^{\top} U \Lambda U^{\top} U (\Lambda/n + \lambda I)^{-1} U^{\top} U \Lambda U^{\top}) \leq \frac{d}{n^2} \\ \|M\|_{\mathrm{op}} &= \lambda_{\max}(M) \leq \frac{1}{n}, \end{aligned}$$

Moreover, since the components of V are independent and of zero mean, and $|V_i| \leq 1$, the variables are 1-sub-Gaussian, and hence the Bernstein's inequality for sub-Gaussian random variables in quadratic form (see e.g. Theorem 2.1 of [110]) implies that with probability at least $1 - \delta$,

$$\|\nabla \ell_{\mathcal{D}}(\theta^{\star})\|_{(\Sigma_{\mathcal{D}}+\lambda I)^{-1}}^{2} = V^{\top}MV \le C_{1} \cdot \frac{d + \log(1/\delta)}{n}.$$

Here C_1 is some universal constant. This gives us

$$\begin{split} \gamma \|\Delta\|_{\Sigma_{\mathcal{D}}+\lambda I}^2 &\leq \|\nabla \ell_{\mathcal{D}}(\theta^{\star})\|_{(\Sigma_{\mathcal{D}}+\lambda I)^{-1}} \|\Delta\|_{\Sigma_{\mathcal{D}}+\lambda I} + 4\lambda\gamma B^2 \\ &\leq \sqrt{C_1 \cdot \frac{d + \log(1/\delta)}{n}} \|\Delta\|_{\Sigma_{\mathcal{D}}+\lambda I} + 4\lambda\gamma B^2. \end{split}$$

Solving the above inequality gives us that for some constant C_2 ,

$$\|\Delta\|_{\Sigma_{\mathcal{D}}+\lambda I} \le C_2 \cdot \sqrt{\frac{d + \log(1/\delta)}{\gamma^2 n} + \lambda B^2}.$$

Proof of Theorem 7

Proof. Let $J'(\pi) = J(\pi) - \langle \theta^{\star}, v \rangle$. We have

$$\begin{split} \mathsf{SubOpt}(\hat{\pi}_{\mathsf{PE}}) &= J(\pi^{\star}) - J(\hat{\pi}_{\mathsf{PE}}) \\ &= J'(\pi^{\star}) - J'(\hat{\pi}_{\mathsf{PE}}) \\ &= (J'(\pi^{\star}) - \hat{J}(\pi^{\star})) + (\hat{J}(\pi^{\star}) - \hat{J}(\hat{\pi}_{\mathsf{PE}})) + (\hat{J}(\hat{\pi}_{\mathsf{PE}}) - J'(\hat{\pi}_{\mathsf{PE}})). \end{split}$$

Since $\hat{\pi}_{\mathsf{PE}}$ is the optimal policy under expected value $J'(\pi)$, we know that the second difference satisfies $\hat{J}(\pi^*) - \hat{J}(\hat{\pi}_{\mathsf{PE}}) \leq 0$. For the third difference, we have

$$\hat{J}(\hat{\pi}_{\mathsf{PE}}) - J'(\hat{\pi}_{\mathsf{PE}}) = \min_{\theta \in \Theta(\hat{\theta}_{\mathsf{MLE}},\lambda)} \mathbb{E}_{s \sim \rho}[\theta^{\top}(\phi(s,\pi(s)) - v)] - \mathbb{E}_{s \sim \rho}[\theta^{\star \top}(\phi(s,\pi(s)) - v)].$$

From Lemma 6 we know that $\theta^* \in \Theta(\hat{\theta}_{\mathsf{MLE}}, \lambda)$ with probability at least $1 - \delta$. Thus we know that with probability at least $1 - \delta$, $\hat{J}(\hat{\pi}_{\mathsf{PE}}) - J'(\hat{\pi}_{\mathsf{PE}}) \leq 0$. Now combining everything together and condition on the above event, we have

$$\begin{aligned} \mathsf{SubOpt}(\hat{\pi}_{\mathsf{PE}}) &\leq J'(\pi^{\star}) - J(\pi^{\star}) \\ &= \sup_{\theta \in \Theta(\hat{\theta}_{\mathsf{MLE}},\lambda)} \mathbb{E}_{s \sim \rho}[(\theta^{\star} - \theta)^{\top}(\phi(s,\pi^{\star}(s)) - v)] \\ &= \sup_{\theta \in \Theta(\hat{\theta}_{\mathsf{MLE}},\lambda)} \mathbb{E}_{s \sim \rho}[(\theta^{\star} - \hat{\theta}_{\mathsf{MLE}} + \hat{\theta}_{\mathsf{MLE}} - \theta)^{\top}(\phi(s,\pi^{\star}(s)) - v)] \\ &= \mathbb{E}_{s \sim \rho}[(\theta^{\star} - \hat{\theta}_{\mathsf{MLE}})^{\top}(\phi(s,\pi^{\star}(s)) - v)] \\ &+ \sup_{\theta \in \Theta(\hat{\theta}_{\mathsf{MLE}},\lambda)} \mathbb{E}_{s \sim \rho}[(\hat{\theta}_{\mathsf{MLE}} - \theta)^{\top}(\phi(s,\pi^{\star}(s)) - v)]. \end{aligned}$$

By the definition of $\Theta(\hat{\theta}_{\mathsf{MLE}}, \lambda)$, we know that for any $\theta \in \Theta(\hat{\theta}_{\mathsf{MLE}}, \lambda)$, one has $\mathbb{E}_{s \sim \rho}[(\hat{\theta}_{\mathsf{MLE}} - \theta)^{\top}(\phi(s, \pi^{\star}(s)) - v)] \leq C \cdot \sqrt{\frac{d + \log(1/\delta)}{\gamma^{2}n} + \lambda B^{2}} \cdot \|(\Sigma_{\mathcal{D}} + \lambda I)^{-1/2} \mathbb{E}_{s \sim \rho}[\phi(s, \pi^{\star}(s)) - v]\|_{2}$. Furthermore, we know that $\theta^{\star} \in \Theta(\hat{\theta}_{\mathsf{MLE}}, \lambda)$ from Lemma 6. Altogether we have with probability $1 - \delta$

$$\mathsf{SubOpt}(\hat{\pi}_{\mathsf{PE}}) \leq 2C \cdot \sqrt{\frac{d + \log(1/\delta)}{\gamma^2 n} + \lambda B^2} \cdot \|(\Sigma_{\mathcal{D}} + \lambda I)^{-1/2} \mathbb{E}_{s \sim \rho}[\phi(s, \pi^{\star}(s)) - v]\|_2.$$

Proof of Theorem 14

Proof. Consider 4 actions with parameter $\phi(a_1) = [1, 1, 0], \phi(a_2) = [1, 0, 0], \phi(a_3) = [0, 0, 0], \phi(a_4) = [0, 1, 0]$. Let the true reward be $\theta^* = [-1, 0.1, 0.9] \in \Theta_B$ with B = 2. We query n - 1 times a_1, a_2 and 1 time a_2, a_3 . For the single pairwise comparison result $Y_{2>3}$ between a_2 and a_3 , we know that

$$P(Y_{2>3} = 1) = \frac{\exp((\phi(a_2) - \phi(a_3))^{\top} \theta^{\star})}{1 + \exp((\phi(a_2) - \phi(a_3))^{\top} \theta^{\star})} > 0.26.$$

Now conditioned on the event that $Y_{2>3} = 1$, we know that the MLE aims to find

$$\begin{aligned} \theta_{\mathsf{MLE}} &= \underset{\theta \in \Theta_B}{\arg \min} \ell_{\mathcal{D}}(\theta), \\ \text{where } \ell_{\mathcal{D}}(\theta) &= -n_{1>2} \cdot \log \left(\frac{\exp((\phi(a_1) - \phi(a_2))^\top \theta)}{1 + \exp((\phi(a_1) - \phi(a_2))^\top \theta)} \right) \\ &- n_{1<2} \cdot \log \left(\frac{\exp((\phi(a_2) - \phi(a_1))^\top \theta)}{1 + \exp((\phi(a_2) - \phi(a_1))^\top \theta)} \right) \\ &- \log \left(\frac{\exp((\phi(a_2) - \phi(a_3))^\top \theta)}{1 + \exp((\phi(a_2) - \phi(a_3))^\top \theta)} \right) \\ &= -n_{1>2} \cdot \log \left(\frac{\exp(\theta_2)}{1 + \exp(\theta_2)} \right) - n_{1<2} \cdot \log \left(\frac{\exp(-\theta_2)}{1 + \exp(-\theta_2)} \right) - \log \left(\frac{\exp(\theta_1)}{1 + \exp(\theta_1)} \right) \end{aligned}$$

By concentration of $n_{1>2}$, we know that when n > 500, with probability at least 0.5, we have

$$n_{1<2} > 0.45n.$$

Under this case, the MLE will satisfy at $\hat{\theta}_1 > 0$, $\hat{\theta}_2 < 0.5$. Thus the policy based on MLE estimator will choose action a_1 or a_2 instead of the optimal action a_4 under the events above. The expected suboptimality is

$$\mathbb{E}[V^{\star}(s) - V^{\hat{\pi}_{\mathsf{MLE}}}(s)] \ge 0.26 * 0.5 * 1 > 0.1.$$

On the other hand, one can calculate the coverage as

$$\|\Sigma_{\mathcal{D}}^{-1/2}\mathbb{E}_{s\sim\rho}[\phi(s,\pi^{\star}(s))]\|_{2} = \frac{n}{n-1}.$$

Thus by Theorem 7 we know that pessimistic MLE achieves vanishing error.

Proof of Theorem 15

Proof. Assume without loss of generality that d/3 is some integer. We set $\mathcal{S} = [d/3]$, $\mathcal{A} = \{a_1, a_2, a_3\}$. For each of the s, a_i , we set $\phi(s, a_1) = e_{3s+1}, \phi(s, a_2) = e_{3s+2}, \phi(s, a_3) = 0$. We set the initial distribution of states as $\rho = \text{Unif}([1, 2, \dots, S])$, the query times $n(s, a_2, a_3) = n/S \cdot (1 - 2/\Lambda^2), n(s, a_1, a_3) = n/S \cdot (2/\Lambda^2)$.

Let $v_{-1} = [1/d, 1/d + \Delta, -2/d - \Delta], v_{+1} = [1/d + 2\Delta, 1/d + \Delta, -2/d - 3\Delta]$. We construct 2^S instances, indexed by $\tau \in \{\pm 1\}^S$, where each $\theta_{\tau} = [v_{\tau_1}, v_{\tau_2}, \cdots, v_{\tau_S}]$. One can see that $\mathbb{E}[V_{\mathcal{Q}}(\pi^*) - V_{\mathcal{Q}}^*(\hat{\pi})] = 1/S \cdot \sum_{s \in \mathcal{S}} (r_{\mathcal{Q}}(s, \pi^*(s)) - r_{\mathcal{Q}}(s, \hat{\pi}(s)))$. Under each θ_{τ} , the optimal policy $\pi(s)$ is either a_1 or a_2 . One can verify that $\|\Sigma_{\mathcal{D}}^{-1/2} \mathbb{E}_{s \sim \rho}[\phi(s, \pi^*(s))]\|_2 \leq \Lambda$ and that $\theta_{\tau} \in \Theta_B$ with B = 1 when d > 6 and $\Delta < 1/(6d)$.

Furthermore, for any $\theta_{\tau}, \theta_{\tau'}$ that differs only in the *j*-th coordinate of τ , we have

$$1/S \cdot (r_{Q_{\tau}}(j, \pi^{\star}(j)) - r_{Q_{\tau}}(j, \hat{\pi}(j)) + r_{Q_{\tau'}}(j, \pi^{\star}(j)) - r_{Q_{\tau'}}(j, \hat{\pi}(j))) \ge \Delta/S.$$

Thus by Assouad's lemma (see e.g. [295]), we have

$$\inf_{\hat{\pi}} \sup_{\mathcal{Q} \in \mathsf{CB}(\lambda)} \mathbb{E}[V_{\mathcal{Q}}(\pi^{\star}) - V_{\mathcal{Q}}^{\star}(\hat{\pi})] \geq S \cdot \frac{\Delta}{2S} \min_{\tau \sim \tau'} (1 - \mathsf{TV}(\mathbb{P}_{\theta_{\tau}}, \mathbb{P}_{\theta_{\tau'}}))) \\
\geq \frac{\Delta}{4} \min_{\tau \sim \tau'} \exp(-D_{\mathsf{KL}}(\mathbb{P}_{\theta_{\tau}}, \mathbb{P}_{\theta_{\tau'}})).$$

Here $\tau \sim \tau'$ refers to any τ, τ' that only differs in one element. And the last inequality is due to the Bretagnolle–Huber inequality [26]. To bound the KL divergence, we have the following lemma from [233]:

Lemma 37 ([233]). For any pair of quality score vectors θ_{τ} and $\theta_{\tau'}$, we have

$$D_{\mathrm{KL}}(\mathbb{P}_{\theta_{\tau}} \| \mathbb{P}_{\theta_{\tau}}) \le Cn(\theta_{\tau} - \theta_{\tau'})^{\top} \Sigma_{\mathcal{D}}(\theta_{\tau} - \theta_{\tau'}).$$
(A.3)

From the lemma, we have

$$\inf_{\hat{\pi}} \sup_{\mathcal{Q} \in \mathsf{CB}(\lambda)} \mathbb{E}[V_{\mathcal{Q}}(\pi^{\star}) - V_{\mathcal{Q}}^{\star}(\hat{\pi})] \geq \frac{\Delta}{2} \min_{\tau \sim \tau'} \exp(-D_{\mathsf{KL}}(\mathbb{P}_{\theta_{\tau}}, \mathbb{P}_{\theta_{\tau'}})) \\ \geq \frac{\Delta}{2} \exp(-Cn\Delta^2/(S\Lambda^2))$$

Taking $\Delta = \Lambda \sqrt{S/n}$ and noting that S = d/3 finishes the proof.

Proof of Theorem 16

This section presents the proof of Theorem 16 for the setting of K-wise comparisons. We first prove the following lemma on the estimation error:

Lemma 38. Under the K-wise PL model, for any $\lambda > 0$, with probability at least $1 - \delta$,

$$\|\hat{\theta}_{\mathsf{MLE}} - \theta^{\star}\|_{\Sigma_{\mathcal{D}} + \lambda I} \le C \cdot \sqrt{\frac{K^4 (d + \log(1/\delta))}{\gamma^2 n}} + \lambda B^2.$$

Recall that the MLE is given by

$$\begin{split} \hat{\theta}_{\mathsf{MLE}} &\in \underset{\theta \in \Theta_B}{\arg\min} \, \ell_{\mathcal{D}}(\theta), \\ \text{where } \ell_{\mathcal{D}}(\theta) &= -\frac{1}{n} \sum_{i=1}^{n} \sum_{j=0}^{K-1} \log \left(\frac{\exp(\langle \theta, \, \phi(s^i, a^i_{\sigma_i(j)}) \rangle)}{\sum_{k=j}^{K-1} \exp(\langle \theta, \, \phi(s^i, a^i_{\sigma_i(k)}) \rangle)} \right) \end{split}$$

Our goal is to bound the estimation error of the MLE in the squared semi-norm $||v||_{\Sigma_{\mathcal{D}}+\lambda I}^2 = v^{\top}(\Sigma_{\mathcal{D}}+\lambda I)v.$

Strong convexity of ℓ . We first show that $\ell_{\mathcal{D}}$ is strongly convex at θ^* with respect to the semi-norm $\|\cdot\|_{\Sigma_{\mathcal{D}}}$, meaning that there is some constant $\gamma > 0$ such that

$$\ell_{\mathcal{D}}(\theta^{\star} + \Delta) - \ell_{\mathcal{D}}(\theta^{\star}) - \langle \nabla \ell_{\mathcal{D}}(\theta^{\star}), \Delta \rangle \ge \gamma \|\Delta\|_{\Sigma_{\mathcal{D}}}^2$$
(A.4)

for all perturbations $\Delta \in \mathbb{R}^d$ such that $\theta^* + \Delta \in \Theta_B$.

The gradient of the negative log likelihood is

$$\nabla \ell_{\mathcal{D}}(\theta) = -\frac{1}{n} \sum_{i=1}^{n} \sum_{j=0}^{K-1} \sum_{k=j}^{K-1} \frac{\exp(\langle \theta, \phi(s^{i}, a^{i}_{\sigma_{i}(k)}) \rangle)}{\sum_{k'=j}^{K-1} \exp(\langle \theta, \phi(s^{i}, a^{i}_{\sigma_{i}(k')}) \rangle)} \cdot (\phi(s^{i}, a^{i}_{\sigma_{i}(j)}) - \phi(s^{i}, a^{i}_{\sigma_{i}(k)})).$$

The Hessian of the negative log likelihood can be written as

$$\nabla^{2} \ell_{\mathcal{D}}(\theta)$$

$$= \frac{1}{n} \sum_{i=1}^{n} \sum_{j=0}^{K-1} \sum_{k=j}^{K-1} \sum_{k'=j}^{K-1} \frac{\exp(\langle \theta, \phi(s^{i}, a^{i}_{\sigma_{i}(k)}) + \phi(s^{i}, a^{i}_{\sigma_{i}(k')}) \rangle)}{2(\sum_{k'=j}^{K-1} \exp(\langle \theta, \phi(s^{i}, a^{i}_{\sigma_{i}(k')}) \rangle))^{2}} \cdot (\phi(s^{i}, a^{i}_{\sigma_{i}(k)}) - \phi(s^{i}, a^{i}_{\sigma_{i}(k')}))(\phi(s^{i}, a^{i}_{\sigma_{i}(k)}) - \phi(s^{i}, a^{i}_{\sigma_{i}(k')})))^{\top}$$

Since $\exp(\langle \theta, \phi \rangle) \in [\exp(-LB), \exp(LB)]$, we know that the coefficients satisfy

$$\frac{\exp(\langle \theta, \, \phi(s^i, a^i_{\sigma_i(k)}) + \phi(s^i, a^i_{\sigma_i(k')}) \rangle)}{(\sum_{k'=j}^{K-1} \exp(\langle \theta, \, \phi(s^i, a^i_{\sigma_i(k')}) \rangle))^2} \ge \frac{\exp(-4LB)}{2(K-j)^2}.$$

Set $\gamma = \exp(-4LB)/2$. We can verify that for any vector $v \in \mathbb{R}^K$, one has

$$\begin{split} v^{\top} \nabla^{2} \ell_{\mathcal{D}}(\theta) v &\geq \frac{\gamma}{n} v^{\top} \Big(\sum_{i=1}^{n} \sum_{j=0}^{K-1} \frac{1}{(K-j)^{2}} \sum_{k=j}^{K-1} \sum_{k'=k}^{K-1} (\phi(s^{i}, a^{i}_{\sigma_{i}(k)}) - \phi(s^{i}, a^{i}_{\sigma_{i}(k')})) \\ &\quad \cdot (\phi(s^{i}, a^{i}_{\sigma_{i}(k)}) - \phi(s^{i}, a^{i}_{\sigma_{i}(k')}))^{\top} \Big) v \\ &\geq \frac{\gamma}{n} v^{\top} \Big(\sum_{i=1}^{n} \min_{\sigma_{i} \in \Pi[K]} \sum_{j=0}^{K-1} \frac{1}{(K-j)^{2}} \sum_{k=j}^{K-1} \sum_{k'=k}^{K-1} (\phi(s^{i}, a^{i}_{\sigma_{i}(k)}) - \phi(s^{i}, a^{i}_{\sigma_{i}(k')})) \\ &\quad \cdot (\phi(s^{i}, a^{i}_{\sigma_{i}(k)}) - \phi(s^{i}, a^{i}_{\sigma_{i}(k')}))^{\top} \Big) v \\ &\geq \gamma v^{\top} \Sigma_{\mathcal{D}} v \\ &= \gamma \|v\|_{\Sigma_{\mathcal{D}}}^{2}. \end{split}$$

Thus we know that ℓ is γ -strongly convex with respect to the semi-norm $\|\cdot\|_{\Sigma_{\mathcal{D}}}$.

Bounding the estimation error. Now we bound the estimation error $\|\hat{\theta}_{\mathsf{MLE}} - \theta^{\star}\|_{\Sigma_{\mathcal{D}+\lambda I}}$. Since $\hat{\theta}_{\mathsf{MLE}}$ is optimal for $\ell_{\mathcal{D}}$, we have $\ell_{\mathcal{D}}(\hat{\theta}_{\mathsf{MLE}}) \leq \ell_{\mathcal{D}}(\theta^{\star})$. Defining the error vector $\Delta = \hat{\theta}_{\mathsf{MLE}} - \theta^{\star}$, adding and subtracting the quantity $\langle \nabla \ell_{\mathcal{D}}(\theta^{\star}), \Delta \rangle$ yields the bound

$$\ell_{\mathcal{D}}(\theta^{\star} + \Delta) - \ell_{\mathcal{D}}(\theta^{\star}) - \langle \nabla \ell_{\mathcal{D}}(\theta^{\star}), \Delta \rangle \leq - \langle \nabla \ell_{\mathcal{D}}(\theta^{\star}), \Delta \rangle.$$

By the γ -convexity condition, the left-hand side is lower bounded by $\gamma \|\Delta\|_{\Sigma_{\mathcal{D}}}^2$. As for the right-hand side, note that $|\langle \nabla \ell_{\mathcal{D}}(\theta^*), \Delta \rangle| \leq \|\nabla \ell_{\mathcal{D}}(\theta^*)\|_{(\Sigma_{\mathcal{D}}+\lambda I)^{-1}} \|\Delta\|_{\Sigma_{\mathcal{D}}+\lambda I}$ for any $\lambda > 0$. Altogether we have

$$\gamma \|\Delta\|_{\Sigma_{\mathcal{D}}}^2 \le \|\nabla \ell_{\mathcal{D}}(\theta^\star)\|_{(\Sigma_{\mathcal{D}}+\lambda I)^{-1}} \|\Delta\|_{\Sigma_{\mathcal{D}}+\lambda I}.$$

Now we further bound the term $\|\nabla \ell_{\mathcal{D}}(\theta^*)\|_{(\Sigma_{\mathcal{D}}+\lambda I)^{-1}}$. Observe that the gradient takes the form

$$\nabla \ell_{\mathcal{D}}(\theta^{\star}) = -\frac{1}{n} \sum_{i=1}^{n} \sum_{j=0}^{K-1} \sum_{k=j}^{K-1} \frac{\exp(\langle \theta^{\star}, \phi(s^{i}, a^{i}_{\sigma_{i}(k)}) \rangle)}{\sum_{k'=j}^{K-1} \exp(\langle \theta^{\star}, \phi(s^{i}, a^{i}_{\sigma_{i}(k')}) \rangle)} \cdot (\phi(s^{i}, a^{i}_{\sigma_{i}(j)}) - \phi(s^{i}, a^{i}_{\sigma_{i}(k)})).$$
(A.5)

We set $x_{jk}^i = \phi(s^i, a_j^i) - \phi(s^i, a_k^i)$. $X \in \mathbb{R}^{(nK(K-1)/2) \times d}$ has the differencing vector x_{jk}^i as its $(iK(K-1)/2 + k + \sum_{l=K-j+1}^{K} l)^{th}$ row. We also define V_{jk}^i be the random variable of the coefficient of x_{jk}^i in Equation (A.5) under the PL model, i.e. conditioned on an arbitrary permutation σ_i ,

$$V_{jk}^{i} = \begin{cases} \frac{\exp(\langle \theta^{\star}, \phi(s^{i}, a_{k}^{i}) \rangle)}{\sum_{k'=\sigma_{i}^{-1}(j)}^{K-1} \exp(\langle \theta^{\star}, \phi(s^{i}, a_{\sigma_{i}(k')}^{i}) \rangle)}, & \text{if } \sigma_{i}^{-1}(j) < \sigma_{i}^{-1}(k) \\ -\frac{\exp(\langle \theta^{\star}, \phi(s^{i}, a_{j}^{i}) \rangle)}{\sum_{k'=\sigma_{i}^{-1}(k)}^{K-1} \exp(\langle \theta^{\star}, \phi(s^{i}, a_{\sigma_{i}(k')}^{i}) \rangle)}, & \text{otherwise.} \end{cases}$$
Here $\sigma_i^{-1}(j) < \sigma_i^{-1}(k)$ means that the *j*-th item ranks higher than the *k*-th item. Let $\tilde{V}_i \in \mathbb{R}^{K(K-1)/2}$ be the concatenated random vector of $\{V_{jk}^i\}_{0 \leq j < k \leq K-1}, V \in \mathbb{R}^{nK(K-1)/2}$ be the concatenated random vector of $\{\tilde{V}_i\}_{i=1}^n$. We know that \tilde{V}_i and \tilde{V}_j are independent for each $i \neq j$ due to the independent sampling procedure. We can also verify that the mean of \tilde{V}_i is 0, the proof of which is deferred to the end of this section. Furthermore, since under any permutation, the sum of absolute value of each element in \tilde{V}_i is at most K, we know that \tilde{V}_i is sub-Gaussian with parameter K. Thus we know that V is also sub-Gaussian with mean 0 and parameter K. Now we know that the term $\|\nabla \ell_{\mathcal{D}}(\theta^*)\|_{(\Sigma_{\mathcal{D}}+\lambda I)^{-1}}^2$ can be written as

$$\|\nabla \ell_{\mathcal{D}}(\theta^{\star})\|_{(\Sigma_{\mathcal{D}}+\lambda I)^{-1}}^{2} = \frac{1}{n^{2}} V^{\top} X (\Sigma_{\mathcal{D}}+\lambda I)^{-1} X^{\top} V.$$

Let $M = \frac{K^2}{n}I$. One can verify that $M \succeq \frac{1}{n^2}X(\Sigma_{\mathcal{D}} + \lambda I)^{-1}X^{\top}$ almost surely since $\lambda_{\max}(X(\Sigma_{\mathcal{D}} + \lambda I)^{-1}X^{\top}/n^2) \leq K^2/n$. Thus we can upper bound the original term as

$$\|\nabla \ell_{\mathcal{D}}(\theta^{\star})\|_{(\Sigma_{\mathcal{D}}+\lambda I)^{-1}}^2 \leq \frac{K^2}{n} \|V\|_2^2.$$

By Bernstein's inequality for sub-Gaussian random variables in quadratic form (see e.g. Theorem 2.1 of [110]), we know that with probability at least $1 - \delta$,

$$||V||_2^2 \le CK^2 \cdot (d + \log(1/\delta)).$$

Thus altogether, we have

$$\gamma \|\Delta\|_{\Sigma_{\mathcal{D}}}^2 \le \sqrt{\frac{CK^4 \cdot (d + \log(1/\delta))}{n}} \|\Delta\|_{\Sigma_{\mathcal{D}} + \lambda I}.$$

Similar to the pairwise comparison analysis in Appendix A.2, we can derive that with probability at least $1 - \delta$,

$$\|\hat{\theta}_{\mathsf{MLE}} - \theta^{\star}\|_{\Sigma_{\mathcal{D}} + \lambda I} \le C \cdot \sqrt{\frac{K^4 (d + \log(1/\delta))}{n} + \lambda B^2}.$$

The rest of the proof on the sub-optimality upper bound follows the same argument as Theorem 7.

Lastly, we verify that the mean of \tilde{V}_i is 0. For any fixed $j, k \in [K]$, let \mathcal{P} be the ordered set of all elements which are ranked higher than both j and k. Now conditioned on \mathcal{P} , we have

$$\mathbb{E}[V_{jk}^{i} \mid \mathcal{P}] = \mathbb{P}(j \text{ follows } \mathcal{P} \mid \mathcal{P}) \cdot \frac{\exp(\langle \theta^{\star}, \phi(s^{i}, a_{k}^{i}) \rangle)}{\sum_{k' \in \bar{\mathcal{P}}} \exp(\langle \theta^{\star}, \phi(s^{i}, a_{k'}^{i}) \rangle)} \\ - \mathbb{P}(k \text{ follows } \mathcal{P} \mid \mathcal{P}) \cdot \frac{\exp(\langle \theta^{\star}, \phi(s^{i}, a_{j}^{i}) \rangle)}{\sum_{k' \in \bar{\mathcal{P}}} \exp(\langle \theta^{\star}, \phi(s^{i}, a_{k'}^{i}) \rangle)} \\ = \frac{1}{\sum_{k' \in \bar{\mathcal{P}}} \exp(\langle \theta^{\star}, \phi(s^{i}, a_{k'}^{i}) \rangle)} \\ \cdot \frac{\exp(\langle \theta^{\star}, \phi(s^{i}, a_{j}^{i}) \rangle) \exp(\langle \theta^{\star}, \phi(s^{i}, a_{k}^{i}) \rangle) - \exp(\langle \theta^{\star}, \phi(s^{i}, a_{k}^{i}) \rangle) \exp(\langle \theta^{\star}, \phi(s^{i}, a_{k}^{i}) \rangle)} \\ = 0.$$

Here the second equality uses the fact that j follows \mathcal{P} is equivalent to the event that j is larger than k and either j, k is the largest among $\overline{\mathcal{P}}$. Taking expectation over \mathcal{P} gives us that $\mathbb{E}[V_{jk}^i] = 0$.

Proof of Theorem 17

This section presents the proof of Theorem 17 for the setting of K-wise comparisons. We first prove the following lemma on the estimation error.

Lemma 39. Under the K-wise PL model, for any $\lambda > 0$, with probability at least $1 - \delta$,

$$\|\hat{\theta}_{\mathsf{MLE}_2} - \theta^\star\|_{\Sigma_{\mathcal{D}} + \lambda I} \le C \cdot \sqrt{\frac{d + \log(1/\delta)}{\gamma^2 n}} + \lambda B^2.$$

Recall that the pairwise compairson based estimator is given by

$$\begin{aligned} \hat{\theta}_{\mathsf{MLE}_2} &\in \underset{\theta \in \Theta_B}{\arg\min} \, \ell_{\mathcal{D}}(\theta), \\ \text{where } \ell_{\mathcal{D}}(\theta) &= -\frac{1}{n} \sum_{i=1}^n \sum_{j=0}^{K-1} \sum_{k=j+1}^{K-1} \log \left(\frac{\exp(\langle \theta, \, \phi(s^i, a^i_{\sigma_i(j)}) \rangle)}{\exp(\langle \theta, \, \phi(s^i, a^i_{\sigma_i(j)}) \rangle) + \exp(\langle \theta, \, \phi(s^i, a^i_{\sigma_i(k)}) \rangle)} \right). \end{aligned}$$

Our goal is to bound the estimation error of the MLE in the squared semi-norm $||v||_{\Sigma_{\mathcal{D}}+\lambda I}^2 = v^{\top}(\Sigma_{\mathcal{D}}+\lambda I)v.$

Strong convexity of ℓ . Let $x_{jk}^i = \phi(s^i, a_j^i) - \phi(s^i, a_k^i)$. The gradient of the negative log likelihood is

$$\nabla \ell_{\mathcal{D}}(\theta) = -\frac{1}{n} \sum_{i=1}^{n} \sum_{j=0}^{K-1} \sum_{k=j+1}^{K-1} \frac{\exp(-\langle \theta, x^{i}_{\sigma_{i}(j)\sigma_{i}(k)} \rangle)}{1 + \exp(-\langle \theta, x^{i}_{\sigma_{i}(j)\sigma_{i}(k)} \rangle)} \cdot x^{i}_{\sigma_{i}(j)\sigma_{i}(k)}.$$

The Hessian of the negative log likelihood can be written as

$$\nabla^2 \ell_{\mathcal{D}}(\theta) = \frac{1}{n} \sum_{i=1}^n \sum_{j=0}^{K-1} \sum_{k=j}^{K-1} \frac{\exp(-\langle \theta, x^i_{\sigma_i(j)\sigma_i(k)}) \rangle}{(1 + \exp(-\langle \theta, x^i_{\sigma_i(j)\sigma_i(k)}) \rangle)^2} \cdot x^i_{\sigma_i(j)\sigma_i(k)} x^{i\top}_{\sigma_i(j)\sigma_i(k)}$$

Since $\exp(\langle \theta, x^i_{\sigma_i(j)\sigma_i(k)} \rangle) \in [\exp(-2LB), \exp(2LB)]$, we know that the coefficients satisfy

$$\frac{\exp(-\langle \theta, x^i_{\sigma_i(j)\sigma_i(k)})\rangle}{(1+\exp(-\langle \theta, x^i_{\sigma_i(j)\sigma_i(k)})\rangle)^2} \ge \frac{1}{2+\exp(2LB)+\exp(-2LB)}.$$

Set $\gamma = \frac{1}{2 + \exp(2LB) + \exp(-2LB)}$. We can verify that for any vector $v \in \mathbb{R}^K$, one has

$$v^{\top} \nabla^{2} \ell_{\mathcal{D}}(\theta) v \geq \frac{\gamma}{n} v^{\top} \left(\sum_{i=1}^{n} \sum_{j=0}^{K-1} \sum_{k=j+1}^{K-1} x_{\sigma_{i}(j)\sigma_{i}(k)}^{i} x_{\sigma_{i}(j)\sigma_{i}(k)}^{i\top} \right) v$$
$$= \frac{\gamma}{n} v^{\top} \left(\sum_{i=1}^{n} \sum_{j=0}^{K-1} \sum_{k=j+1}^{K-1} x_{jk}^{i} x_{jk}^{i\top} \right) v$$
$$= \gamma K (K-1) v^{\top} \Sigma_{\mathcal{D}} v / 2$$
$$= \gamma K (K-1) \|v\|_{\Sigma_{\mathcal{D}}}^{2} / 2.$$

Thus we know that ℓ is γ -strongly convex with respect to the semi-norm $\|\cdot\|_{\Sigma_{\mathcal{D}}}$.

Bounding the estimation error. Now we bound the estimation error $\|\hat{\theta}_{\mathsf{MLE}_2} - \theta^\star\|_{\Sigma_{\mathcal{D}+\lambda I}}$. Since $\hat{\theta}_{\mathsf{MLE}_2}$ is optimal for $\ell_{\mathcal{D}}$, we have $\ell_{\mathcal{D}}(\hat{\theta}_{\mathsf{MLE}_2}) \leq \ell_{\mathcal{D}}(\theta^\star)$. Defining the error vector $\Delta = \hat{\theta}_{\mathsf{MLE}_2} - \theta^\star$, adding and subtracting the quantity $\langle \nabla \ell_{\mathcal{D}}(\theta^\star), \Delta \rangle$ yields the bound

$$\ell_{\mathcal{D}}(\theta^{\star} + \Delta) - \ell_{\mathcal{D}}(\theta^{\star}) - \langle \nabla \ell_{\mathcal{D}}(\theta^{\star}), \Delta \rangle \leq - \langle \nabla \ell_{\mathcal{D}}(\theta^{\star}), \Delta \rangle.$$

By the γ -convexity condition, the left-hand side is lower bounded by $\gamma K(K-1) \|\Delta\|_{\Sigma_{\mathcal{D}}}^2/2$. As for the right-hand side, note that $|\langle \nabla \ell_{\mathcal{D}}(\theta^*), \Delta \rangle| \leq \|\nabla \ell_{\mathcal{D}}(\theta^*)\|_{(\Sigma_{\mathcal{D}}+\lambda I)^{-1}} \|\Delta\|_{\Sigma_{\mathcal{D}}+\lambda I}$ for any $\lambda > 0$. Altogether we have

$$\gamma \|\Delta\|_{\Sigma_{\mathcal{D}}}^2 \le 2 \|\nabla \ell_{\mathcal{D}}(\theta^\star)\|_{(\Sigma_{\mathcal{D}}+\lambda I)^{-1}} \|\Delta\|_{\Sigma_{\mathcal{D}}+\lambda I}/K(K-1).$$

Now we further bound the term $\|\nabla \ell_{\mathcal{D}}(\theta^*)\|_{(\Sigma_{\mathcal{D}}+\lambda I)^{-1}}$. Observe that the gradient takes the form

$$\nabla \ell_{\mathcal{D}}(\theta^{\star}) = -\frac{1}{n} \sum_{i=1}^{n} \sum_{j=0}^{K-1} \sum_{k=j+1}^{K-1} \frac{\exp(-\langle \theta, x^{i}_{\sigma_{i}(j)\sigma_{i}(k)})\rangle}{1 + \exp(-\langle \theta, x^{i}_{\sigma_{i}(j)\sigma_{i}(k)})\rangle} \cdot x^{i}_{\sigma_{i}(j)\sigma_{i}(k)}.$$
 (A.6)

We set $X \in \mathbb{R}^{(nK(K-1)/2) \times d}$ with the differencing vector x_{jk}^i as its $(iK(K-1)/2 + k + \sum_{l=K-j+1}^{K} l)^{th}$ row. We also define V_{jk}^i be the random variable of the coefficient of x_{jk}^i in Equation (A.6) under the PL model, i.e. conditioned on an arbitrary permutation σ_i ,

$$V_{jk}^{i} = \begin{cases} \frac{\exp(-\langle \theta, x_{jk}^{i} \rangle)}{1 + \exp(-\langle \theta, x_{jk}^{i} \rangle)}, & \text{if } \sigma_{i}^{-1}(j) < \sigma_{i}^{-1}(k) \\ -\frac{1}{1 + \exp(-\langle \theta, x_{jk}^{i} \rangle)}, & \text{otherwise.} \end{cases}$$

Let $\tilde{V}_i \in \mathbb{R}^{K(K-1)/2}$ be the concatenated random vector of $\{V_{jk}^i\}_{0 \leq j < k \leq K-1}, V \in \mathbb{R}^{nK(K-1)/2}$ be the concatenated random vector of $\{\tilde{V}_i\}_{i=1}^n$. We know that \tilde{V}_i is independent for each i, and that V is sub-Gaussian with mean 0 and parameter $\sqrt{K(K-1)/2}$ since the PL model reduces to BTL model when considering pairwise comparisons. Now we know that the term $\|\nabla \ell_{\mathcal{D}}(\theta^*)\|_{(\Sigma_{\mathcal{D}}+\lambda I)^{-1}}^2$ can be written as

$$\|\nabla \ell_{\mathcal{D}}(\theta^{\star})\|_{(\Sigma_{\mathcal{D}}+\lambda I)^{-1}}^{2} = \frac{1}{n^{2}} V^{\top} X (\Sigma_{\mathcal{D}}+\lambda I)^{-1} X^{\top} V.$$

Let $M = \frac{K^2}{n}I$. One can verify that $M \succeq \frac{1}{n^2}X(\Sigma_{\mathcal{D}} + \lambda I)^{-1}X^{\top}$ almost surely since $\lambda_{\max}(X(\Sigma_{\mathcal{D}} + \lambda I)^{-1}X^{\top}/n^2) \leq K^2/n$. Thus we can upper bound the original term as

$$\|\nabla \ell_{\mathcal{D}}(\theta^{\star})\|_{(\Sigma_{\mathcal{D}}+\lambda I)^{-1}}^2 \leq \frac{K^2}{n} \|V\|_2^2.$$

By Bernstein's inequality for sub-Gaussian random variables in quadratic form (see e.g. Theorem 2.1 of [110]), we know that with probability at least $1 - \delta$,

$$||V||_2^2 \le CK(K-1) \cdot (d + \log(1/\delta)).$$

Thus altogether, we have

$$\gamma \|\Delta\|_{\Sigma_{\mathcal{D}}}^2 \leq \sqrt{\frac{C \cdot (d + \log(1/\delta))}{n}} \|\Delta\|_{\Sigma_{\mathcal{D}} + \lambda I}$$

Similar to the pairwise comparison, we can derive that with probability at least $1 - \delta$,

$$\|\hat{\theta}_{\mathsf{MLE}_2} - \theta^\star\|_{\Sigma_{\mathcal{D}} + \lambda I} \le C \cdot \sqrt{\frac{d + \log(1/\delta)}{n} + \lambda B^2}.$$

The rest of the proof on the sub-optimality upper bound follows the same argument as Theorem 7.

Proof of Lemma 19

Recall that the MLE is given by

$$\begin{split} \theta_{\mathsf{MLE}} &\in \arg\min_{\theta\in\Theta_B} \ell_{\mathcal{D}}(\theta),\\ \text{where } \ell_{\mathcal{D}}(\theta) &= -\sum_{i=1}^{n} \log\left(1(y^{i}=1) \cdot \frac{\exp(\sum_{h=1}^{H} r_{\theta}(s_{h}^{i},a_{h}^{i})) + \exp(\sum_{h=1}^{H} r_{\theta}(s_{h}^{i},a_{h}^{i}) - r_{\theta}(s_{h}^{i},a_{h}^{i}))) + 1 \\ &= -\sum_{i=1}^{n} \log\left(1(y^{i}=1) \cdot \frac{1}{\exp(\sum_{h=1}^{H} (r_{\theta}(s_{h}^{i},a_{h}^{i}) - r_{\theta}(s_{h}^{i},a_{h}^{i})) + 1\right) \\ &= -\sum_{i=1}^{n} \log\left(1(y^{i}=1) \cdot \frac{1}{\exp(-\langle\theta,\sum_{h=1}^{H} (\phi(s_{h}^{i},a_{h}^{i}) - \phi(s_{h}^{i},a_{h}^{i}))) + 1\right) \\ &+ 1(y^{i}=0) \cdot \frac{1}{\exp(\langle\theta,\sum_{h=1}^{H} (\phi(s_{h}^{i},a_{h}^{i}) - \phi(s_{h}^{i},a_{h}^{i}))\rangle) + 1} \\ & + 1(y^{i}=0) \cdot \frac{1}{\exp(\langle\theta,\sum_{h=1}^{H} (\phi(s_{h}^{i},a_{h}^{i}) - \phi(s_{h}^{i},a_{h}^{i}))\rangle) + 1} \\ \end{split}$$

To simplify the notation, we let $x_i = \sum_{h=1}^{H} (\phi(s_h^i, a_h^i) - \phi(s_h^{i\prime}, a_h^{i\prime}))$. Our goal is to bound the estimation error of the MLE in the squared semi-norm $\|v\|_{\Sigma_{\mathcal{D}}+\lambda I}^2 = v^{\top}(\Sigma_{\mathcal{D}}+\lambda I)v$.

Strong convexity of ℓ . We first show that $\ell_{\mathcal{D}}$ is strongly convex at θ^* with respect to the semi-norm $\|\cdot\|_{\Sigma_{\mathcal{D}}}$, meaning that there is some constant $\gamma > 0$ such that

$$\ell_{\mathcal{D}}(\theta^{\star} + \Delta) - \ell_{\mathcal{D}}(\theta^{\star}) - \langle \nabla \ell_{\mathcal{D}}(\theta^{\star}), \Delta \rangle \ge \gamma \|\Delta\|_{\Sigma_{\mathcal{D}}}^2$$
(A.7)

for all perturbations $\Delta \in \mathbb{R}^d$ such that $\theta^* + \Delta \in \Theta_B$.

One can directly calculate the Hessian of ℓ as

$$\nabla^2 \ell_{\mathcal{D}}(\theta) = \frac{1}{n} \sum_{i=1}^n \left(1(y^i = 1) \cdot \frac{\exp(-\langle \theta, x_i \rangle)}{(\exp(-\langle \theta, x_i \rangle) + 1)^2} + 1(y^i = 0) \cdot \frac{\exp(\langle \theta, x_i \rangle)}{(\exp(\langle \theta, x_i \rangle) + 1)^2} \right) \cdot x_i x_i^{\mathsf{T}},$$

Observe that $\langle \theta, x_i \rangle \in [-2HLB, 2HLB]$, we have

$$v^{\top} \nabla^2 \ell_{\mathcal{D}}(\theta) v \ge \frac{\gamma}{n} \|Xv\|_2^2$$
 for all v ,

where $\gamma = 1/(2 + \exp(-2HLB) + \exp(2HLB)), X \in \mathbb{R}^{n \times d}$ has the differencing vector $x_i \in \mathbb{R}^d$ as its i^{th} row.

Thus, if we introduce the error vector $\Delta \coloneqq \hat{\theta}_{\mathsf{MLE}} - \theta^{\star}$, then we may conclude that

$$\ell_{\mathcal{D}}(\theta^{\star} + \Delta) - \ell_{\mathcal{D}}(\theta^{\star}) - \langle \nabla \ell_{\mathcal{D}}(\theta^{\star}), \Delta \rangle \ge \frac{\gamma}{n} \|X\Delta\|_{2}^{2} = \gamma \|\Delta\|_{\Sigma_{\mathcal{D}}}^{2},$$

showing that $\ell_{\mathcal{D}}$ is strongly convex around θ^{\star} with parameter γ .

Bounding the estimation error. Now we aim at bounding the estimation error $\|\hat{\theta}_{\mathsf{MLE}} -$ $\theta^{\star}\|_{\Sigma_{\mathcal{D}}}.$

Since $\hat{\theta}_{\mathsf{MLE}}$ is optimal for $\ell_{\mathcal{D}}$, we have $\ell_{\mathcal{D}}(\hat{\theta}_{\mathsf{MLE}}) \leq \ell_{\mathcal{D}}(\theta^{\star})$. Defining the error vector $\Delta = \hat{\theta}_{\mathsf{MLE}} - \theta^{\star}$, adding and subtracting the quantity $\langle \nabla \ell_{\mathcal{D}}(\theta^{\star}), \Delta \rangle$ yields the bound

$$\ell_{\mathcal{D}}(\theta^{\star} + \Delta) - \ell_{\mathcal{D}}(\theta^{\star}) - \langle \nabla \ell_{\mathcal{D}}(\theta^{\star}), \Delta \rangle \leq - \langle \nabla \ell_{\mathcal{D}}(\theta^{\star}), \Delta \rangle.$$

By the γ -convexity condition, the left-hand side is lower bounded by $\gamma \|\Delta\|_{\Sigma_{\mathcal{D}}}^2$. As for the right-hand side, note that $|\langle \nabla \ell_{\mathcal{D}}(\theta^*), \Delta \rangle| \leq \|\nabla \ell_{\mathcal{D}}(\theta^*)\|_{(\Sigma_{\mathcal{D}}+\lambda I)^{-1}} \|\Delta\|_{\Sigma_{\mathcal{D}}+\lambda I}$ for any $\lambda > 0$. Altogether we have

$$\gamma \|\Delta\|_{\Sigma_{\mathcal{D}}}^2 \le \|\nabla \ell_{\mathcal{D}}(\theta^\star)\|_{(\Sigma_{\mathcal{D}}+\lambda I)^{-1}} \|\Delta\|_{\Sigma_{\mathcal{D}}+\lambda I}$$

Now we further bound the term $\|\nabla \ell_{\mathcal{D}}(\theta^{\star})\|_{(\Sigma_{\mathcal{D}}+\lambda I)^{-1}}$. Observe that the gradient takes the form

$$\nabla \ell_{\mathcal{D}}(\theta^{\star}) = \frac{-1}{n} \sum_{i=1}^{n} \left[1[y^i = 1] \frac{\exp(-\langle \theta^{\star}, x_i \rangle)}{1 + \exp(-\langle \theta^{\star}, x_i \rangle))} - 1[y^i = 0] \frac{1}{1 + \exp(-\langle \theta^{\star}, x_i \rangle))} \right] x_i.$$

Define a random vector $V \in \mathbb{R}^n$ with independent components as

$$V_i = \begin{cases} \frac{\exp(-\langle \theta^{\star}, x_i \rangle)}{1 + \exp(-\langle \theta^{\star}, x_i \rangle))} & \text{w.p.} & \frac{1}{1 + \exp(-\langle \theta^{\star}, x_i \rangle))} \\ \frac{-1}{1 + \exp(-\langle \theta^{\star}, x_i \rangle))} & \text{w.p.} & \frac{\exp(-\langle \theta^{\star}, x_i \rangle)}{1 + \exp(-\langle \theta^{\star}, x_i \rangle))} \end{cases}$$

With this notation, we have $\nabla \ell_{\mathcal{D}}(\theta^*) = -\frac{1}{n} X^\top V$. One can verify that $\mathbb{E}[V] = 0$ and $|V_i| \leq 1$. Define the *n*-dimensional matrix $M \coloneqq \frac{1}{n^2} X (\Sigma_{\mathcal{D}} + \lambda I)^{-1} X^\top$. We have $\|\nabla \ell_{\mathcal{D}}(\theta^*)\|_{(\Sigma_{\mathcal{D}} + \lambda I)^{-1}} = V^\top M V$. Let the eigenvalue decomposition of XX^\top be $XX^\top = U\Lambda U^\top$. We can bound the trace and operator norm of M as

$$\begin{aligned} \mathrm{Tr}(M) &= \frac{1}{n^2} \mathrm{Tr}(U(\Lambda/n + \lambda I)^{-1} U^{\top} U \Lambda U^{\top}) \leq \frac{d}{n} \\ \|M\|_{\mathrm{op}} &= \lambda_{\max}(M) \leq \frac{1}{n}, \end{aligned}$$

Moreover, since the components of V are independent and of zero mean, and $|V_i| \leq 1$, the variables are 1-sub-Gaussian, and hence the Bernstein's inequality for sub-Gaussian random variables in quadratic form (see e.g. Theorem 2.1 of [110]) implies that with probability at least $1 - \delta$,

$$\|\nabla \ell_{\mathcal{D}}(\theta^{\star})\|_{(\Sigma_{\mathcal{D}}+\lambda I)^{-1}}^{2} = V^{\top}MV \le C_{1} \cdot \frac{d + \log(1/\delta)}{n}$$

Here C_1 is some universal constant. This gives us

$$\begin{split} \gamma \|\Delta\|_{\Sigma_{\mathcal{D}}+\lambda I}^2 &\leq \|\nabla \ell_{\mathcal{D}}(\theta^{\star})\|_{(\Sigma_{\mathcal{D}}+\lambda I)^{-1}} \|\Delta\|_{\Sigma_{\mathcal{D}}+\lambda I} + 4\lambda\gamma B^2 \\ &\leq \sqrt{C_1 \cdot \frac{d + \log(1/\delta)}{n}} \|\Delta\|_{\Sigma_{\mathcal{D}}+\lambda I} + 4\lambda\gamma B^2. \end{split}$$

Solving the above inequality gives us that for some constant C_2 ,

$$\|\Delta\|_{\Sigma_{\mathcal{D}}+\lambda I} \le C_2 \cdot \sqrt{\frac{d + \log(1/\delta)}{\gamma^2 n} + \lambda B^2}.$$

Proof of Theorem 20

Proof. From Lemma 19, we know that with probability at least $1 - \delta$,

$$\|\hat{\theta}_{\mathsf{MLE}} - \theta^{\star}\|_{\Sigma_{\mathcal{D}} + \lambda I} \le C \cdot \sqrt{\frac{d + \log(1/\delta)}{\gamma^2 n}} + \lambda B^2.$$

Let $J'(\pi) = J(\pi) - H\langle \theta^*, v \rangle$. We have

$$\begin{aligned} \mathsf{SubOpt}(\hat{\pi}_{\mathsf{PE}}) &= J(\pi^{\star}) - J(\hat{\pi}_{\mathsf{PE}}) \\ &= J'(\pi^{\star}) - J'(\hat{\pi}_{\mathsf{PE}}) \\ &= (J'(\pi^{\star}) - \hat{J}(\pi^{\star})) + (\hat{J}(\pi^{\star}) - \hat{J}(\hat{\pi}_{\mathsf{PE}})) + (\hat{J}(\hat{\pi}_{\mathsf{PE}}) - J'(\hat{\pi}_{\mathsf{PE}})). \end{aligned}$$

Since $\hat{\pi}_{\mathsf{PE}}$ is the optimal policy under expected value $\hat{J}(\pi)$, we know that the second difference satisfies $\hat{J}(\pi^*) - \hat{J}(\hat{\pi}_{\mathsf{PE}}) \leq 0$. For the third difference, we have

$$\hat{J}(\hat{\pi}_{\mathsf{PE}}) - J'(\hat{\pi}_{\mathsf{PE}}) = \mathbb{E}_{s \sim d^{\hat{\pi}_{\mathsf{PE}}}}[\hat{r}(s, \hat{\pi}_{\mathsf{PE}}(s)) - r(s, \hat{\pi}_{\mathsf{PE}}(s))].$$

From Lemma 19 we know that $\theta^* \in \Theta(\hat{\theta}_{\mathsf{MLE}}, \lambda)$ with probability at least $1 - \delta$. Thus we know that with probability at least $1 - \delta$, $\hat{J}(\hat{\pi}_{\mathsf{PE}}) - J'(\hat{\pi}_{\mathsf{PE}}) \leq 0$. Now combining everything together, we have

$$\begin{aligned} \mathsf{SubOpt}(\hat{\pi}_{\mathsf{PE}}) &\leq J'(\pi^{\star}) - J(\pi^{\star}) \\ &= \sup_{\theta \in \Theta(\hat{\theta}_{\mathsf{MLE}},\lambda)} \mathbb{E}_{s \sim d^{\pi^{\star}}}[(\theta^{\star} - \theta)^{\top}(\phi(s, \pi^{\star}(s)) - v)] \\ &= \sup_{\theta \in \Theta(\hat{\theta}_{\mathsf{MLE}},\lambda)} \mathbb{E}_{s \sim d^{\pi^{\star}}}[(\theta^{\star} - \hat{\theta}_{\mathsf{MLE}} + \hat{\theta}_{\mathsf{MLE}} - \theta)^{\top}(\phi(s, \pi^{\star}(s)) - v)] \\ &= \mathbb{E}_{s \sim d^{\pi^{\star}}}[(\theta^{\star} - \hat{\theta}_{\mathsf{MLE}})^{\top}(\phi(s, \pi^{\star}(s)) - v)] \\ &+ \sup_{\theta \in \Theta(\hat{\theta}_{\mathsf{MLE}},\lambda)} \mathbb{E}_{s \sim d^{\pi^{\star}}}[(\hat{\theta}_{\mathsf{MLE}} - \theta)^{\top}(\phi(s, \pi^{\star}(s)) - v)]. \end{aligned}$$

By the definition of $\Theta(\hat{\theta}_{\mathsf{MLE}}, \lambda)$, we know that for any $\theta \in \Theta(\hat{\theta}_{\mathsf{MLE}}, \lambda)$, one has $\mathbb{E}_{s \sim d^{\pi^{\star}}}[(\hat{\theta}_{\mathsf{MLE}} - \theta)^{\top}(\phi(s, \pi^{\star}(s)) - v)] \leq C \cdot \sqrt{\frac{d + \log(1/\delta)}{\gamma^2 n} + \lambda B^2} \cdot \|(\Sigma_{\mathcal{D}} + \lambda I)^{-1/2} \mathbb{E}_{s \sim d^{\pi^{\star}}}[\phi(s, \pi^{\star}(s)) - v]\|_2$. Furthermore, we know that $\hat{\theta}^{\star} \in \Theta(\hat{\theta}_{\mathsf{MLE}}, \lambda)$ from Lemma 19. Altogether we have with probability $1 - 2\delta$

$$\mathsf{SubOpt}(\hat{\pi}_{\mathsf{PE}}) \le 2C \cdot \sqrt{\frac{d + \log(1/\delta)}{\gamma^2 n}} + \lambda B^2 \cdot \|(\Sigma_{\mathcal{D}} + \lambda I)^{-1/2} \mathbb{E}_{s \sim d^{\pi^*}}[\phi(s, \pi^*(s)) - v]\|_2.$$

Proof of Theorem 23

Proof. Here We mainly prove Lemma 22, since Theorem 23 is a direct corollary when combined with the proof in Theorem 20.

Our goal is to bound the estimation error of the MLE in the squared semi-norm $||v||_{\Sigma_{\mathcal{D}}+\lambda I}^2 = v^{\top}(\Sigma_{\mathcal{D}}+\lambda I)v.$

Strong convexity of ℓ . We first show that $\ell_{\mathcal{D}}$ is strongly convex at θ^* with respect to the semi-norm $\|\cdot\|_{\Sigma_{\mathcal{D}}}$, meaning that there is some constant $\gamma > 0$ such that

$$\ell_{\mathcal{D}}(\theta^{\star} + \Delta) - \ell_{\mathcal{D}}(\theta^{\star}) - \langle \nabla \ell_{\mathcal{D}}(\theta^{\star}), \Delta \rangle \ge \gamma \|\Delta\|_{\Sigma_{\mathcal{D}}}^2$$
(A.8)

for all perturbations $\Delta \in \mathbb{R}^d$ such that $\theta^* + \Delta \in \Theta_B$.

The gradient of the negative log likelihood is

$$\nabla \ell_{\mathcal{D}}(\theta) = -\frac{1}{n} \sum_{i=1}^{n} \sum_{\tau' \in \mathcal{T}(s_0^i)} \frac{\exp(\sum_{h=0}^{H} \langle \theta, \phi(s_h', a_h') \rangle)}{\sum_{\tau'' \in \mathcal{T}(s_0^i)} \exp(\sum_{h=0}^{H} \langle \theta, \phi(s_h'', a_h') \rangle)} \cdot \left(\sum_{h=0}^{H} (\phi(s_h^i, a_h^i) - \phi(s_h', a_h')) \right) \cdot \left(\sum_{h=0}^{H} (\phi(s_h^i, a_h^i) - \phi(s_h', a_h')) \right) + \left(\sum_{h=0}^{H} (\phi(s_h^i, a_h^i) - \phi(s_h', a_h')) \right) + \left(\sum_{h=0}^{H} (\phi(s_h^i, a_h^i) - \phi(s_h', a_h')) \right) + \left(\sum_{h=0}^{H} (\phi(s_h^i, a_h^i) - \phi(s_h', a_h')) \right) + \left(\sum_{h=0}^{H} (\phi(s_h^i, a_h^i) - \phi(s_h', a_h')) \right) + \left(\sum_{h=0}^{H} (\phi(s_h^i, a_h^i) - \phi(s_h', a_h')) \right) + \left(\sum_{h=0}^{H} (\phi(s_h^i, a_h^i) - \phi(s_h', a_h')) \right) + \left(\sum_{h=0}^{H} (\phi(s_h^i, a_h^i) - \phi(s_h', a_h')) \right) + \left(\sum_{h=0}^{H} (\phi(s_h^i, a_h^i) - \phi(s_h', a_h')) \right) + \left(\sum_{h=0}^{H} (\phi(s_h^i, a_h^i) - \phi(s_h', a_h')) \right) + \left(\sum_{h=0}^{H} (\phi(s_h^i, a_h^i) - \phi(s_h', a_h')) \right) + \left(\sum_{h=0}^{H} (\phi(s_h^i, a_h^i) - \phi(s_h', a_h')) \right) + \left(\sum_{h=0}^{H} (\phi(s_h^i, a_h^i) - \phi(s_h', a_h')) \right) + \left(\sum_{h=0}^{H} (\phi(s_h^i, a_h^i) - \phi(s_h', a_h')) \right) + \left(\sum_{h=0}^{H} (\phi(s_h^i, a_h^i) - \phi(s_h', a_h')) \right) + \left(\sum_{h=0}^{H} (\phi(s_h^i, a_h^i) - \phi(s_h', a_h')) \right) + \left(\sum_{h=0}^{H} (\phi(s_h^i, a_h^i) - \phi(s_h', a_h')) \right) + \left(\sum_{h=0}^{H} (\phi(s_h^i, a_h^i) - \phi(s_h', a_h')) \right) + \left(\sum_{h=0}^{H} (\phi(s_h^i, a_h^i) - \phi(s_h', a_h')) \right) + \left(\sum_{h=0}^{H} (\phi(s_h^i, a_h') - \phi(s_h', a_h')) \right) + \left(\sum_{h=0}^{H} (\phi(s_h^i, a_h') - \phi(s_h', a_h')) \right) + \left(\sum_{h=0}^{H} (\phi(s_h^i, a_h') - \phi(s_h', a_h')) \right) + \left(\sum_{h=0}^{H} (\phi(s_h^i, a_h') - \phi(s_h', a_h')) \right) + \left(\sum_{h=0}^{H} (\phi(s_h^i, a_h') - \phi(s_h', a_h')) \right) + \left(\sum_{h=0}^{H} (\phi(s_h^i, a_h') - \phi(s_h', a_h')) \right) + \left(\sum_{h=0}^{H} (\phi(s_h^i, a_h') - \phi(s_h', a_h')) \right) + \left(\sum_{h=0}^{H} (\phi(s_h^i, a_h') - \phi(s_h', a_h')) \right) + \left(\sum_{h=0}^{H} (\phi(s_h^i, a_h') - \phi(s_h', a_h')) \right) + \left(\sum_{h=0}^{H} (\phi(s_h^i, a_h') - \phi(s_h', a_h')) \right) + \left(\sum_{h=0}^{H} (\phi(s_h', a_h') - \phi(s_h', a_h')) \right) + \left(\sum_{h=0}^{H} (\phi(s_h', a_h') - \phi(s_h', a_h')) \right) + \left(\sum_{h=0}^{H} (\phi(s_h', a_h') - \phi(s_h', a_h')) \right) + \left(\sum_{h=0}^{H} (\phi(s_h', a_h') - \phi(s_h', a_h') \right) \right) + \left(\sum_{h=0}^{H} (\phi(s_h', a_h') - \phi(s_h', a_h') \right) + \left(\sum_{h=0}^{H}$$

Let $x_{\tau,\tau'}^i = \sum_{h=0}^H (\phi(s_h, a_h) - \phi(s'_h, a'_h))$, where $\tau = \{(s_h, a_h)\}_{h \in [H]}, \tau' = \{(s'_h, a'_h)\}_{h \in [H]}$. The Hessian of the negative log likelihood can be written as

$$\nabla^2 \ell_{\mathcal{D}}(\theta) = \frac{1}{n} \sum_{i=1}^n \sum_{\tau \in \mathcal{T}(s_0^i)} \sum_{\tau' \in \mathcal{T}(s_0^i)} \frac{\exp(\sum_{h=0}^H \langle \theta, \phi(s_h, a_h) + \phi(s'_h, a'_h) \rangle)}{2(\sum_{\tau'' \in \mathcal{T}(s_0^i)} \exp(\sum_{h=0}^H \langle \theta, \phi(s''_h, a''_h) \rangle))^2} \cdot x^i_{\tau, \tau'} x^{i\top}_{\tau, \tau'}.$$

Since $\exp(\langle \theta, \phi \rangle) \in [\exp(-LB), \exp(LB)]$, we know that the coefficients satisfy

$$\frac{\exp(\sum_{h=0}^{H} \langle \theta, \phi(s_h, a_h) + \phi(s'_h, a'_h) \rangle)}{2(\sum_{\tau'' \in \mathcal{T}(s_0^i)} \exp(\sum_{h=0}^{H} \langle \theta, \phi(s''_h, a''_h) \rangle))^2} \ge \frac{\exp(-4LB)}{2\sup_s |\mathcal{T}(s)|^2}$$

Set $\gamma = \exp(-4LB)/2$. We can verify that for any vector $v \in \mathbb{R}^K$, one has

$$v^{\top} \nabla^2 \ell_{\mathcal{D}}(\theta) v \ge \gamma v^{\top} \Sigma_{\mathcal{D}} v = \gamma \|v\|_{\Sigma_{\mathcal{D}}}^2.$$

Thus we know that ℓ is γ -strongly convex with respect to the semi-norm $\|\cdot\|_{\Sigma_{\mathcal{D}}}$.

Bounding the estimation error. Now we bound the estimation error $\|\hat{\theta}_{\mathsf{MLE}} - \theta^{\star}\|_{\Sigma_{\mathcal{D}+\lambda I}}$. Since $\hat{\theta}_{\mathsf{MLE}}$ is optimal for $\ell_{\mathcal{D}}$, we have $\ell_{\mathcal{D}}(\hat{\theta}_{\mathsf{MLE}}) \leq \ell_{\mathcal{D}}(\theta^{\star})$. Defining the error vector $\Delta = \hat{\theta}_{\mathsf{MLE}} - \theta^{\star}$, adding and subtracting the quantity $\langle \nabla \ell_{\mathcal{D}}(\theta^{\star}), \Delta \rangle$ yields the bound

$$\ell_{\mathcal{D}}(\theta^{\star} + \Delta) - \ell_{\mathcal{D}}(\theta^{\star}) - \langle \nabla \ell_{\mathcal{D}}(\theta^{\star}), \Delta \rangle \leq - \langle \nabla \ell_{\mathcal{D}}(\theta^{\star}), \Delta \rangle$$

By the γ -convexity condition, the left-hand side is lower bounded by $\gamma \|\Delta\|_{\Sigma_{\mathcal{D}}}^2$. As for the right-hand side, note that $|\langle \nabla \ell_{\mathcal{D}}(\theta^*), \Delta \rangle| \leq \|\nabla \ell_{\mathcal{D}}(\theta^*)\|_{(\Sigma_{\mathcal{D}}+\lambda I)^{-1}} \|\Delta\|_{\Sigma_{\mathcal{D}}+\lambda I}$ for any $\lambda > 0$. Altogether we have

$$\gamma \|\Delta\|_{\Sigma_{\mathcal{D}}}^2 \le \|\nabla \ell_{\mathcal{D}}(\theta^\star)\|_{(\Sigma_{\mathcal{D}}+\lambda I)^{-1}} \|\Delta\|_{\Sigma_{\mathcal{D}}+\lambda I}.$$

Now we further bound the term $\|\nabla \ell_{\mathcal{D}}(\theta^*)\|_{(\Sigma_{\mathcal{D}}+\lambda I)^{-1}}$. Observe that the gradient takes the form

$$\nabla \ell_{\mathcal{D}}(\theta^{\star}) = -\frac{1}{n} \sum_{i=1}^{n} \sum_{\tau' \in \mathcal{T}(s_0^i)} \frac{\exp(\sum_{h=0}^{H} \langle \theta^{\star}, \phi(s_h', a_h') \rangle)}{\sum_{\tau'' \in \mathcal{T}(s_0^i)} \exp(\sum_{h=0}^{H} \langle \theta^{\star}, \phi(s_h'', a_h') \rangle)} \cdot \left(\sum_{h=0}^{H} (\phi(s_h^i, a_h^i) - \phi(s_h', a_h'))\right)$$
(A.9)

We set X as the concatenated differencing vector $x^i_{\tau,\tau'}$ where τ, τ' are distinct and ordered. We also define $V^i_{\tau,\tau'}$ be the random variable of the coefficient of $x^i_{\tau,\tau'}$ in Equation (A.9), i.e.

$$V_{\tau,\tau'}^{i} = \begin{cases} \frac{\exp(\sum_{h=0}^{H} \langle \theta^{\star}, \phi(s_{h}', a_{h}') \rangle)}{\sum_{\tau'' \in \mathcal{T}(s_{0}^{i})} \exp(\sum_{h=0}^{H} \langle \theta^{\star}, \phi(s_{h}', a_{h}') \rangle)}, & \text{ if } \tau = \{(s_{h}^{i}, a_{h}^{i})\}_{h \in [H]}, \\ -\frac{\exp(\sum_{h=0}^{H} \langle \theta^{\star}, \phi(s_{h}, a_{h}) \rangle)}{\sum_{\tau'' \in \mathcal{T}(s_{0}^{i})} \exp(\sum_{h=0}^{H} \langle \theta^{\star}, \phi(s_{h}'', a_{h}'') \rangle)}, & \text{ if } \tau' = \{(s_{h}^{i}, a_{h}^{i})\}_{h \in [H]}, \\ 0 & \text{ otherwise.} \end{cases}$$

Let \tilde{V}_i be the concatenated random vector of $\{V_{\tau,\tau'}^i\}$, V be the concatenated random vector of $\{\tilde{V}_i\}_{i=1}^n$. We know that \tilde{V}_i and \tilde{V}_j are independent for each $i \neq j$ due to the independent sampling procedure. We can also verify that the mean of \tilde{V}_i is 0. We know that \tilde{V}_i has almost $\sup_s |\mathcal{T}(s)|$ non-zero elements. And the sum of their absolute value is bounded by 1. we know \tilde{V}_i is 1-sub-Gaussian. Now we know that the term $\|\nabla \ell_{\mathcal{D}}(\theta^*)\|_{(\Sigma_{\mathcal{D}}+\lambda I)^{-1}}^2$ can be written as

$$\|\nabla \ell_{\mathcal{D}}(\theta^{\star})\|_{(\Sigma_{\mathcal{D}}+\lambda I)^{-1}}^{2} = \frac{1}{n^{2}} V^{\top} X (\Sigma_{\mathcal{D}}+\lambda I)^{-1} X^{\top} V.$$

Let $M = \frac{\sup_s |\mathcal{T}(s)|^2}{n} I$. One can verify that $M \succeq \frac{1}{n^2} X (\Sigma_{\mathcal{D}} + \lambda I)^{-1} X^{\top}$ almost surely since $\lambda_{\max}(X(\Sigma_{\mathcal{D}} + \lambda I)^{-1} X^{\top}/n^2) \leq \sup_s |\mathcal{T}(s)|^2/n$. Thus we can upper bound the original term as

$$\|\nabla \ell_{\mathcal{D}}(\theta^*)\|_{(\Sigma_{\mathcal{D}}+\lambda I)^{-1}}^2 \le \frac{\sup_s |\mathcal{T}(s)|^2}{n} \|V\|_2^2$$

By Bernstein's inequality for sub-Gaussian random variables in quadratic form (see e.g. Theorem 2.1 of [110]), we know that with probability at least $1 - \delta$,

$$||V||_2^2 \le C \cdot (d + \log(1/\delta)).$$

Thus altogether, we have

$$\gamma \|\Delta\|_{\Sigma_{\mathcal{D}}}^2 \le \sqrt{\frac{C \sup_s |\mathcal{T}(s)|^2 \cdot (d + \log(1/\delta))}{n}} \|\Delta\|_{\Sigma_{\mathcal{D}} + \lambda I}.$$

Similar to the pairwise comparison analysis in Appendix A.2, we can derive that with probability at least $1 - \delta$,

$$\|\hat{\theta}_{\mathsf{MLE}} - \theta^{\star}\|_{\Sigma_{\mathcal{D}} + \lambda I} \le C \cdot \sqrt{\frac{\sup_{s} |\mathcal{T}(s)|^{2} (d + \log(1/\delta))}{n}} + \lambda B^{2}$$

The rest of the proof on the sub-optimality upper bound follows the same argument as Theorem 20.

Proof of Theorem 36

To simplify the notation, we let $f_{\theta}^i = r_{\theta}(s^i, a_1^i) - r_{\theta}(s^i, a_0^i)$. We can see that the gradient of ℓ takes the form

$$\nabla \ell_{\mathcal{D}}(\theta) = \frac{-1}{n} \sum_{i=1}^{n} \left[1[y^i = 1] \frac{\exp(-f^i_{\theta})}{1 + \exp(-f^i_{\theta})} - 1[y^i = 0] \frac{1}{1 + \exp(-f^i_{\theta})} \right] \nabla f^i_{\theta}$$

And the Hessian of ℓ is

$$\begin{split} \nabla^2 \ell_{\mathcal{D}}(\theta) &= \frac{1}{n} \sum_{i=1}^n \Big(\frac{\exp(f_{\theta}^i)}{(\exp(f_{\theta}^i) + 1)^2} \cdot \nabla f_{\theta}^i \nabla f_{\theta}^{i\top} - \frac{1(y^i = 1) \cdot \exp(-f_{\theta}^i)}{1 + \exp(-f_{\theta}^i)} \cdot \nabla^2 f_{\theta}^i \\ &+ \frac{1(y^i = 0) \cdot \exp(f_{\theta}^i)}{1 + \exp(f_{\theta}^i)} \cdot \nabla^2 f_{\theta}^i \Big). \end{split}$$

Now from Assumption 35, we have

$$\nabla^2 \ell_{\mathcal{D}}(\theta) \succeq \frac{1}{n} \sum_{i=1}^n \gamma \nabla f_{\theta}^i \nabla f_{\theta}^{i\top} - 2\alpha_2 I.$$

where $\gamma = \frac{1}{2 + \exp(-2LB) + \exp(2LB)}$. Now from the Lipschitz gradient assumption we also know that $\|\nabla f^i_{\theta} - \nabla f^i_{\theta^\star}\| \leq 2\alpha_2 \|\theta^\star - \theta\|$. Let $u = \nabla f^i_{\theta} - \nabla f^i_{\theta^\star}$, we have

$$\nabla^{2} \ell_{\mathcal{D}}(\theta) \succeq \frac{1}{n} \sum_{i=1}^{n} \gamma (\nabla f_{\theta^{\star}}^{i} + u) (\nabla f_{\theta^{\star}}^{i} + u)^{\top} - 2\alpha_{2}I$$
$$\succeq \frac{1}{n} \sum_{i=1}^{n} \gamma \nabla f_{\theta^{\star}}^{i} \nabla f_{\theta^{\star}}^{i\top} + \gamma (\nabla f_{\theta^{\star}}^{i} u^{\top} + u \nabla f_{\theta^{\star}}^{i\top}) - 2\alpha_{2}I$$

Since $u^{\top}v \leq ||u||_2 ||v||_2 \leq 2\alpha_2 B ||v||_2$, $v^{\top} \nabla f^i_{\theta^{\star}} \leq \alpha_1 ||v||_2$, this gives that

$$v^{\top} \nabla^2 \ell_{\mathcal{D}}(\theta) v \ge \frac{\gamma}{n} \|Xv\|_2^2 - 2\alpha_2 (1 + 2\gamma\alpha_1 B) \|v\|_2^2 \quad \text{for all } v,$$

where $X \in \mathbb{R}^{n \times d}$ has the vector $\nabla f_{\theta^{\star}}^i \in \mathbb{R}^d$ as its i^{th} row. Thus, if we introduce the error vector $\Delta \coloneqq \hat{\theta}_{\mathsf{MLE}} - \theta^{\star}$, then we may conclude that

$$\ell_{\mathcal{D}}(\theta^{\star} + \Delta) - \ell_{\mathcal{D}}(\theta^{\star}) - \langle \nabla \ell_{\mathcal{D}}(\theta^{\star}), \Delta \rangle \geq \frac{\gamma}{n} \|X\Delta\|_{2}^{2} - 2\alpha_{2}(1 + 2\gamma\alpha_{1}B)\|\Delta\|_{2}^{2}$$
$$= \gamma \|\Delta\|_{\Sigma_{\mathcal{D}}}^{2} - 2\alpha_{2}(1 + 2\gamma\alpha_{1}B)\|\Delta\|_{2}^{2}.$$

Bounding the estimation error. Now we aim at bounding the estimation error $\|\hat{\theta}_{\mathsf{MLE}} \theta^* \|_{\Sigma_{\mathcal{D}}}$. Since $\hat{\theta}_{\mathsf{MLE}}$ is optimal for $\ell_{\mathcal{D}}$, we have $\ell_{\mathcal{D}}(\hat{\theta}_{\mathsf{MLE}}) \leq \ell_{\mathcal{D}}(\theta^*)$. (When $\hat{\theta}_{\mathsf{MLE}}$ is approximately optimal, i.e. $\ell_{\mathcal{D}}(\hat{\theta}_{\mathsf{MLE}}) \leq \min_{\theta} \ell_{\mathcal{D}}(\theta) + \epsilon$, the same argument also holds up to an extra additive term ϵ .) Defining the error vector $\Delta = \hat{\theta}_{\mathsf{MLE}} - \theta^*$, adding and subtracting the quantity $\langle \nabla \ell_{\mathcal{D}}(\theta^{\star}), \Delta \rangle$ yields the bound

$$\ell_{\mathcal{D}}(\theta^{\star} + \Delta) - \ell_{\mathcal{D}}(\theta^{\star}) - \langle \nabla \ell_{\mathcal{D}}(\theta^{\star}), \Delta \rangle \leq - \langle \nabla \ell_{\mathcal{D}}(\theta^{\star}), \Delta \rangle$$

We know the left-hand side is lower bounded by $\gamma \|\Delta\|_{\Sigma_{\mathcal{D}}}^2 - 2\alpha_2(1+2\gamma\alpha_1 B)\|\Delta\|_2^2$. As for the right-hand side, note that $|\langle \nabla \ell_{\mathcal{D}}(\theta^{\star}), \Delta \rangle| \leq ||\nabla \ell_{\mathcal{D}}(\theta^{\star})||_{(\Sigma_{\mathcal{D}}+\lambda I)^{-1}} ||\Delta||_{\Sigma_{\mathcal{D}}+\lambda I}$ for any $\lambda > 0$. Altogether we have

$$\gamma \|\Delta\|_{\Sigma_{\mathcal{D}}}^2 \le \|\nabla \ell_{\mathcal{D}}(\theta^\star)\|_{(\Sigma_{\mathcal{D}}+\lambda I)^{-1}} \|\Delta\|_{\Sigma_{\mathcal{D}}+\lambda I} + \beta \|\Delta\|_2^2,$$

where $\beta = 2\alpha_2(1+2\gamma\alpha_1B)$. Now we further bound the term $\|\nabla \ell_{\mathcal{D}}(\theta^{\star})\|_{(\Sigma_{\mathcal{D}}+\lambda I)^{-1}}$. Observe that the gradient takes the form

$$\nabla \ell_{\mathcal{D}}(\theta^{\star}) = \frac{-1}{n} \sum_{i=1}^{n} \left[1[y^{i} = 1] \frac{\exp(-f^{i}_{\theta^{\star}})}{1 + \exp(-f^{i}_{\theta^{\star}}))} - 1[y^{i} = 0] \frac{1}{1 + \exp(-f^{i}_{\theta^{\star}}))} \right] \nabla f^{i}_{\theta^{\star}}$$

Define a random vector $V \in \mathbb{R}^n$ with independent components as

$$V_i = \begin{cases} \frac{\exp(-f_{\theta^\star}^i)}{1 + \exp(-f_{\theta^\star}^i))} & \text{w.p.} & \frac{1}{1 + \exp(-f_{\theta^\star}^i))} \\ \frac{-1}{1 + \exp(-f_{\theta^\star}^i))} & \text{w.p.} & \frac{\exp(-f_{\theta^\star}^i)}{1 + \exp(-f_{\theta^\star}^i))} \end{cases}$$

With this notation, we have $\nabla \ell_{\mathcal{D}}(\theta^*) = -\frac{1}{n} X^\top V$. One can verify that $\mathbb{E}[V] = 0$ and $|V_i| \leq 1$. Define the *n*-dimensional matrix $M \coloneqq \frac{1}{n^2} X (\Sigma_{\mathcal{D}} + \lambda I)^{-1} X^\top$. We have $\|\nabla \ell_{\mathcal{D}}(\theta^*)\|_{(\Sigma_{\mathcal{D}} + \lambda I)^{-1}} = V^\top M V$. Let the eigenvalue decomposition of $X^\top X$ be $X^\top X = U \Lambda U^\top$. We can bound the trace and operator norm of M as

$$\begin{split} \mathrm{Tr}(M) &= \frac{1}{n^2} \mathrm{Tr}(U(\Lambda/n + \lambda I)^{-1} U^{\top} U \Lambda U^{\top}) \leq \frac{d}{n} \\ \mathrm{Tr}(M^2) &= \frac{1}{n^4} \mathrm{Tr}(U(\Lambda/n + \lambda I)^{-1} U^{\top} U \Lambda U^{\top} U (\Lambda/n + \lambda I)^{-1} U^{\top} U \Lambda U^{\top}) \leq \frac{d}{n^2} \\ \|M\|_{\mathrm{op}} &= \lambda_{\max}(M) \leq \frac{1}{n}, \end{split}$$

Moreover, since the components of V are independent and of zero mean, and $|V_i| \leq 1$, the variables are 1-sub-Gaussian, and hence the Bernstein's inequality for sub-Gaussian random variables in quadratic form (see e.g. Theorem 2.1 of [110]) implies that with probability at least $1-\delta$,

$$\|\nabla \ell_{\mathcal{D}}(\theta^{\star})\|_{(\Sigma_{\mathcal{D}}+\lambda I)^{-1}}^{2} = V^{\top}MV \le C_{1} \cdot \frac{d + \log(1/\delta)}{n}.$$

Here C_1 is some universal constant. This gives us

$$\gamma \|\Delta\|_{\Sigma_{\mathcal{D}}+\lambda I}^{2} \leq \|\nabla \ell_{\mathcal{D}}(\theta^{\star})\|_{(\Sigma_{\mathcal{D}}+\lambda I)^{-1}} \|\Delta\|_{\Sigma_{\mathcal{D}}+\lambda I} + 4(\lambda\gamma + 2\alpha_{2}(1+2\gamma\alpha_{1}B))B^{2}$$
$$\leq \sqrt{C_{1} \cdot \frac{d+\log(1/\delta)}{n}} \|\Delta\|_{\Sigma_{\mathcal{D}}+\lambda I} + 4(\lambda\gamma + 2\alpha_{2}(1+2\gamma\alpha_{1}B))B^{2}.$$

Solving the above inequality gives us that for some constant C_2 ,

$$\|\Delta\|_{\Sigma_{\mathcal{D}}+\lambda I} \le C_2 \cdot \sqrt{\frac{d + \log(1/\delta)}{\gamma^2 n} + (\lambda + \alpha_2/\gamma + \alpha_1 \alpha_2 B)B^2}.$$

Appendix B

Appendix for Practical Implementation of RLHF

B.1 Extension to Multi-wise Comparison

Here we discuss potential extensions from pairwise comparisons to multi-wise comparison. When there is M ranked responses for each prompt, there are two losses that one can choose from, namely MLE_2 and MLE_M from [316].

$$\theta_{\mathsf{MLE}_{2}} \in \arg\min_{r} \mathcal{L}_{2}(\mathcal{D}, r),$$
where $\mathcal{L}_{2}(\mathcal{D}, r) = -\frac{1}{n} \sum_{i=1}^{n} \sum_{j=1}^{n} \sum_{k=j+1}^{M} \log\left(\frac{\exp(r(s_{i}, a_{i}^{(\sigma_{i}(j))}))}{\exp(r(s_{i}, a_{i}^{(\sigma_{i}(j))})) + \exp(r(s_{i}, a_{i}^{(\sigma_{i}(k))}))}\right)$

$$\hat{\theta}_{\mathsf{MLE}_{M}} \in \arg\min_{r} \mathcal{L}_{M}(\mathcal{D}, r),$$
where $\mathcal{L}_{M}(\mathcal{D}, r) = -\frac{1}{n} \sum_{i=1}^{n} \sum_{j=1}^{n} \log\left(\frac{\exp(r(s_{i}, a_{i}^{(\sigma_{i}(j))}))}{\sum_{k=j}^{M} \exp(r(s_{i}, a_{i}^{(\sigma_{i}(k))}))}\right).$

Here we discuss how to incorporate the iterative data smoothing algorithm for the two losses above.

The loss MLE_2 splits the *M*-wise comparisons into pairwise comparisons, thus it is straightforward to predict the new label $y_i^{j,k}$ for each pair of the comparisons between *j*-th and *k*-th response. The loss used for iterative data smoothing can be written as

$$\mathcal{L}_{2}^{\mathsf{DR}}(\mathcal{D}, r) = -\frac{1}{n} \sum_{i=1}^{n} \sum_{j=1}^{M} \sum_{k=j+1}^{M} y_{i}^{\sigma_{i}(j), \sigma_{i}(k)} \log\left(\frac{\exp(r(s_{i}, a_{i}^{(\sigma_{i}(j))})) + \exp(r(s_{i}, a_{i}^{(\sigma_{i}(j))}))}{\exp(r(s_{i}, a_{i}^{(\sigma_{i}(k))})) + \exp(r(s_{i}, a_{i}^{(\sigma_{i}(k))}))}\right) + (1 - y_{i}^{\sigma_{i}(j), \sigma_{i}(k)}) \log\left(\frac{\exp(r(s_{i}, a_{i}^{(\sigma_{i}(j))})) + \exp(r(s_{i}, a_{i}^{(\sigma_{i}(k))}))}{\exp(r(s_{i}, a_{i}^{(\sigma_{i}(j))})) + \exp(r(s_{i}, a_{i}^{(\sigma_{i}(k))}))}\right).$$

$$y_{i,t+1}^{j,k} = (1-\beta) \cdot y_{i,t}^{j,k} + \beta \cdot \frac{\exp(r_{\theta_{t+1}}(s_i, a_i^j))}{\exp(r_{\theta_{t+1}}(s_i, a_i^j)) + \exp(r_{\theta_{t+1}}(s_i, a_i^k))}$$

On the other hand, adapting the loss MLE_M for iterative data smoothing requires more efforts since it requires changing the ranking labels to soft labels. The design of MLE_M decomposes the probability of the observed ranking to the product of the probability that each response is the most preferred one among the rest of the responses. One of the options is to directly change the labels for the current rankings by the following update rules:

$$\mathcal{L}_{M}(\mathcal{D}, r) = -\frac{1}{n} \sum_{i=1}^{n} \sum_{j=1}^{M} y_{i}^{\sigma_{i}(j)} \log \left(\frac{\exp(r(s_{i}, a_{i}^{(\sigma_{i}(j))}))}{\sum_{k=j}^{M} \exp(r(s_{i}, a_{i}^{(\sigma_{i}(k))}))} \right).$$

And the update rule for the labels y_i is

$$y_{i,t+1}^{\sigma_i(j)} = (1-\beta) \cdot y_{i,t}^{\sigma_i(j)} + \beta \cdot \frac{\exp(r_{\theta_{t+1}}(s_i, a_i^{(\sigma_i(j))}))}{\sum_{k=j}^M \exp(r_{\theta_{t+1}}(s_i, a_i^{(\sigma_i(k))}))}$$

However, the above update method does not directly reduce to the the case of pairwise comparisons when setting M = 2. In order to recover the pairwise loss, one needs to consider all possible rankings and get the soft labels for all the rankings. The loss will become

$$\mathcal{L}'_{M}(\mathcal{D}, r) = -\frac{1}{n} \sum_{i=1}^{n} \sum_{\sigma \in \Pi(M)} \sum_{j=1}^{M} y_{i}^{j,\sigma} \log\left(\frac{\exp(r(s_{i}, a_{i}^{(\sigma(j))}))}{\sum_{k=j}^{M} \exp(r(s_{i}, a_{i}^{(\sigma(k))}))}\right)$$

Here $\Pi(M)$ is the set of all permutations of the M elements. And the label is initialized as $y_{i,0}^{j,\sigma} = 1$ if $\sigma = \sigma_i$, and 0 otherwise. And the update rule for the labels y_i is

$$y_{i,t+1}^{j,\sigma} = (1-\beta) \cdot y_{i,t}^{j,\sigma} + \beta \cdot \frac{\exp(r_{\theta_{t+1}}(s_i, a_i^{(\sigma(j))}))}{\sum_{k=j}^M \exp(r_{\theta_{t+1}}(s_i, a_i^{(\sigma(k))}))}.$$

The loss is consistent with the pairwise cross entropy loss when M = 2. However, it requires enumerating over all possible permutations, which are not very efficient when M is large. It requires more study to decide which loss is more appropriate for M-wise iterative data smoothing.

B.2 An Alternative Formulation of Iterative Data Smoothing

Besides the formulation shown in Algorithm 2, we also propose an alternative formulation that directly multiplies a confidence c_i in front of the original loss for each sample, as is **Algorithm 5** Iterative Data Smoothing V2 $(\mathcal{D}, \theta_0, \alpha, \beta)$

Input: The pairwise comparison dataset $\mathcal{D} = \{a_i, a'_i, y_i\}_{i=1}^n$. A parameterized reward model family $\{r_\theta : \mathcal{A} \mapsto \mathbb{R} \mid \theta \in \Theta\}$ with initialization $\theta_0 \in \Theta$. Two step sizes α, β . An empirical loss function

$$\mathcal{L}_{\theta}(\{c_i\}, \mathcal{D}) = -\frac{1}{n} \sum_{i=1}^{n} \max(2c_i - 1, 0) \cdot \left(y_i \cdot \log\left(\frac{\exp(r_{\theta}(a_i))}{\exp(r_{\theta}(a_i)) + \exp(r_{\theta}(a'_i))}\right) + (1 - y_i) \cdot \log\left(\frac{\exp(r_{\theta}(a'_i))}{\exp(r_{\theta}(a_i)) + \exp(r_{\theta}(a'_i))}\right)\right)$$

Initialize t = 0 and $c_{i,0} = 1, \forall i$. while r_{θ_t} does not converge do

$$\theta_{t+1} \leftarrow \theta_t - \alpha \cdot \nabla \mathcal{L}_{\theta}(\{c_{i,t}\}, \mathcal{D})$$

$$c_{i,t+1} \leftarrow (1-\beta) \cdot c_{i,t} + \beta \cdot \frac{\exp(r_{\theta_{t+1}}(a_i))}{\exp(r_{\theta_{t+1}}(a_i)) + \exp(r_{\theta_{t+1}}(a'_i))}$$

$$t \leftarrow t+1$$

end while Return: r_{θ_t}

shown in Algorithm 5. We note here that although the algorithm has better asymptotic convergence result, its performance in practice is not as good as Algorithm 2.

We multiply a $\max(2c_i - 1, 0)$ in front of the loss for each sample as an approximation of how confident the current model predicts the preference label. When the reward is approximately similar, the coefficient goes to 0, putting less weights on those samples. Below we show that asymptotically, the new iterative data smoothing V2 algorithm is better at preserving the preference distribution compared with the original version.

Theorem 40. Consider the multi-armed bandit problem with the number of samples going to infinity and a fixed sampling distribution μ . Assume that $\mu(a, a') > 0$ for any a, a' > 0. Then we have

- Any stationary point for Algorithm 2 satisfies $\forall a, a', \hat{r}(a) = \hat{r}(a')$;
- There is one stationary point for Algorithm 5 that satisfies

$$\forall a, a', \hat{r}(a) - \hat{r}(a') = r^{\star}(a) - r^{\star}(a').$$

Proof. The stationary points for Algorithm 2 and 5 are the points where the gradients equal

0. For Algorithm 2, this is equivalent to $\hat{y} = \frac{\exp(\hat{r}(a))}{\exp(\hat{r}(a)) + \exp(\hat{r}(a'))}$, and

$$(\mu(a \succ a') \cdot \hat{y} + \mu(a \prec a') \cdot (1 - \hat{y})) \cdot \frac{\exp(\hat{r}(a'))}{\exp(\hat{r}(a')) + \exp(\hat{r}(a'))}$$
$$= (\mu(a \prec a') \cdot \hat{y} + \mu(a \succ a') \cdot (1 - \hat{y})) \cdot \frac{\exp(\hat{r}(a))}{\exp(\hat{r}(a)) + \exp(\hat{r}(a'))}.$$

Here \hat{y} can be different for different (a, a'). Solving the above equation gives that the only stationary point is $\hat{y} = 1/2$ and $\hat{r}(a) - \hat{r}(a') = 0$.

On the other hand, for Algorithm 5, the stationary point condition is equivalent to $\hat{c}(a, a') = \frac{\exp(\hat{r}(a))}{\exp(\hat{r}(a)) + \exp(\hat{r}(a'))}$, and

$$\sum_{a'} \max(2\hat{c}(a,a') - 1, 0) \cdot \left(\mu(a \succ a') \cdot \frac{\exp(\hat{r}(a'))}{\exp(\hat{r}(a')) + \exp(\hat{r}(a'))} - \mu(a \prec a') \cdot \frac{\exp(\hat{r}(a))}{\exp(\hat{r}(a)) + \exp(\hat{r}(a'))}\right) = 0.$$

Thus one can verify that $\hat{r}(a) - \hat{r}(a') = r^*(a) - r^*(a')$ satisfies the stationary condition. This proves the result.

Remark 41. Although the asymptotic stationary points of Algorithm 2 do not contain the ground truth, the two-scale analysis discussed in Section 3.3 shows that when one of the step size is much larger than the other such that one of the updates in \hat{y} or \hat{r} is slower (and thus does not hit the stationary point), the reward still converges to the ground truth for those sufficiently observed arms. However, preliminary experiments on Algorithm 5 show that the result is worse than that of Algorithm 2, and also suffer from reward overfitting. This together with the failure of MLE may suggest that asymptotic result does not reflect the practical performance with smaller sample size compared with number of parameters.

Remark 42. The condition of $\mu(a, a') > 0$ can be relaxed to that the comparison graph induced by the Laplace matrix L is connected, since the reward is identifiable in this case [233].

B.3 Experiments

In this section, we present the results of experiments with both multi-armed bandits and neural networks.

Multi-Armed Bandit

In the bandit setting, we focus on the hard example constructed in Theorem 26. We take total samples n = 60 and the number of arms K as 10 and 20. We compare the performance of the vanilla MLE, pessimistic MLE and IDS in both the reward learning phase and the policy learning phase.

In the reward learning phase, we run stochastic gradient descent with learning rate 0.01 on the reward model for multiple epochs and monitor how the loss changes with respect to the number of training epochs. For pessimistic MLE, we subtract the confidence level in the reward according to Equation (3.5). For IDS, we take the two step sizes as $\alpha = 0.01$, $\beta = 0.001$. As is shown in left part of Figure B.1, both MLE and pessimistic MLE suffer from reward overfitting, while the test cross-entropy loss for the IDS algorithm continues to decrease until convergence. Since the training loss changes with the updated labels, we plot the population cross-entropy loss which is averaged over all pairs of comparisons.



Figure B.1: Comparisons of the three methods in the multi-armed bandit setting.

In the right part of the figure, we plot the KL-reward tradeoff when training a policy based on the learned reward. We vary the choice of λ in Equation (3.4) to derive the optimal policy under diverse levels of KL constraint, where we take the reference policy π_0 as the uniform policy. One can see that IDS is able to converge to the optimal reward when KL is large, while both MLE and pessimistic MLE suffer from overoptimization.

We remark here that the reason pessimistic MLE suffers from both overfitting and overoptimization might be due to the design of unbounded reward in the multi-armed bandit case. When the reward family is bounded, pessimistic MLE is also guaranteed to mitigate the overoptimization issue. Furthermore, we only run one random seed for this setting to keep the plot clean since the KL-reward trade-off heavily depends on the observed samples.

Neural Network

We also conduct experiments with neural networks. We use the human-labeled Helpfulness and Harmlessnes (HH) dataset from [18].¹ We take Dahoas/pythia-125M-static-sft² as the policy model with three different reward models of size 125M, 1B and 3B. When training reward model, we take a supervised fine-tuned language model, remove the last layer and replace it with a linear layer. When fine-tuning the language model, we use the proximal policy optimization (PPO) algorithm [229].

We take a fully-trained 6B reward model Dahoas/gptj-rm-static trained from the same dataset based on EleutherAI/gpt-j-6b as the ground truth. We use the model to label the comparison samples using the BTL model [25]. And we train the 125M, 1B and 3B reward model with the new labeled comparison samples. The reward training results are shown in Figure B.2. One can see that the MLE begins to overfit after 1-2 epochs, while the loss of the IDS algorithm continues to decrease stably until convergence.

For both MLE and IDS algorithms, we take the reward with the smallest evaluation loss and optimize a policy against the selected reward model. We compare results for policy learning as shown in Figure B.3. One can see that MLE suffers from reward overoptimization with few thousand steps, while the ground truth reward continues to grow when using our IDS algorithm. We select step sizes $\alpha = 10^{-5}$ and $\beta = 0.7$ for all experiments. We observe that larger model leads to more improvement after one epoch, potentially due to more accurate estimation of the labels. We provide more details of the experiment along with the experiments on a different dataset, TLDR, in Appendix B.3.

In the implementation, we find that it is helpful to restore the best checkpoint at the end of each epoch. This is due to that an inappropriate label $\{y_i\}_{i=1}^n$ at certain epoch may hurt the performance of the model. To prevent overfitting to the test set, we choose a large validation and test dataset, and we select the best checkpoint according to the smallest loss in the validation set, and plot the loss on the test set. During the whole training procedure including checkpoint restoration, we do not use any of the sample in the test set.

We also compare the algorithm with Laplace Smoothing, which simply replaces the hard label $y_i = 1$ with $y_i = 1 - \alpha$, and minimize the following loss function.

$$\mathcal{L}_{\theta}(\{y_i\}, \mathcal{D}) = -\frac{1}{n} \sum_{i=1}^{n} (1-\alpha) \cdot \log\left(\frac{\exp(r_{\theta}(a_i))}{\exp(r_{\theta}(a_i)) + \exp(r_{\theta}(a'_i))}\right) + \alpha \cdot \log\left(\frac{\exp(r_{\theta}(a'_i))}{\exp(r_{\theta}(a_i)) + \exp(r_{\theta}(a'_i))}\right)$$

¹https://huggingface.co/datasets/Dahoas/static-hh

²https://huggingface.co/Dahoas/pythia-125M-static-sft



Figure B.2. Comparisons of MLE and IDS when the reward is parameterized by a neural network.



Figure B.3: Comparison of MLE and IDS for policy learning



Figure B.4: Comparisons of MLE, Laplace Smoothing and IDS.

In our experiments, we conduct hyperparameter search and take $\alpha = 0.05$. The comparison is listed in Figure B.4. One can see that Laplace Smoothing helps makes the convergence point slightly better than MLE, but still much worse than that of IDS.

Additional Experiments on TLDR

The hyper-parameters for the neural network experiments are listed in table B.1.

Model	Parameter	Value
Reward model	learning rate α	10^{-5}
	label update parameter β	0.7
	batch size	128
	eval & save steps	100
Policy model	max sequence length	1024
	max output length	500
	generation temperature	1.0
	batch size	64
	fixed KL coefficient	0.001
	number of rollouts	128
	PPO epochs	4
	value coefficient	0.5
	GAE coefficient λ	0.95
	discount factor	1.0
	clip range	2

Table B.1: Hyper-parameters for the neural network experiments

We also include additional experiments on a different dataset, TLDR³, in Figure B.5 and B.6 of this section. The settings follow the same as HH in Section B.3. One can see that in the case of TLDR, the test accuracy does not drop significantly like HH. However, even with small difference in loss and the accuracy, the resulting policy reward difference is still significant.

B.4 Proof of Theorem 25

Proof. Let $\mathbb{P}_r(a, a', c) = \mathbb{P}_r(a \succ a')$ if c = 1, and $\mathbb{P}_r(a' \succ a)$ if c = 0 be the density function of the observations. According to Theorem 6.1.3. of [109], it suffices to verify the following conditions for the consistency of MLE:

- The CDFs are distinct, i.e. $\mathbb{P}_r(a, a', c) = \mathbb{P}_{r'}(a, a', c)$ almost everywhere implies that r = r'. This is true since the distribution is supported on discrete space, and the equality implies that r(i) r(j) = r(i)' r(j)' for any i, j, and r(M) = r'(M) = 1.
- The PDFs have common support for all r. This is true since the probability is positive for any a, a', c.
- The point r^* is an interior point in \mathbb{R}^K . This is true by definition, since any open ball of radius ϵ around r^* is a subset of the space.

³https://huggingface.co/datasets/CarperAI/openai_summarize_comparisons



Figure B.5. Comparisons of MLE and Iterative Data Smoothing when the reward is parameterized by a neural network.



Figure B.6: Comparison of MLE and Iterative Data Smoothing for policy learning.

B.5 Proof of Theorem 26

Proof. The construction is in similar spirit to [212] and [316]. Consider a bandit problem where $r^*(a) = 1(a = 1)$. For any fixed n, we set $\mu(1, 2) = 1 - 1/n$, $\mu(1, 3) = 1/n$.

In this hard instance, there is constant probability that arm 3 is only compared with arm 1 once. Concretely, we have

$$\mathbb{P}(n(1,3) = 1) = n \cdot (1 - \mu(1,3))^{n-1} \cdot \mu(1,3) = (1 - 1/n)^{n-1}.$$

When $n \ge 500$, we have $\mathbb{P}(n(1,3) = 1) \ge 0.36$. Under this case, we know that arm 3 is preferred with probability at least $\exp(r(3))/(\exp(r(1)) + \exp(r(3))) > 0.26$. When there is only one comparison between arm 1 and 3, and arm 3 is preferred, the MLE assigns r(3) as infinity. Even when the reward for arm 1 is estimated perfectly, this leads to a population cross-entropy loss arbitrarily large.

B.6 Proof of Corollary 27

Proof. The proof follows immediately from Theorem 26. Under the same construction, we know that $\hat{r}_{\mathsf{MLE}}(3) = +\infty$ with probability at least 0.09. Thus, the sub-optimality of the resulting optimal policy is at least 1.

B.7 Proof of Theorem 28

Proof. Let \hat{r}_i be the reward for the *i*-th arm, and $\hat{r} = [\hat{r}_1, \hat{r}_2, \cdots, \hat{r}_K]$ as the vector for the reward. One can calculate the gradient of the reward as

$$\begin{aligned} \nabla_{\hat{r}_{i}} \mathcal{L}_{\mathsf{CE}}(\mathcal{D}, \hat{r}) &= -\frac{1}{n} \sum_{i=1}^{n} \nabla_{\hat{r}_{i}} \left(y_{i} \log \left(\frac{\exp(\hat{r}_{a_{i}})}{\exp(\hat{r}_{a_{i}}) + \exp(\hat{r}_{a_{i}'})} \right) \right. \\ &+ (1 - y_{i})) \log \left(\frac{\exp(\hat{r}_{a_{i}'})}{\exp(\hat{r}_{a_{i}}) + \exp(\hat{r}_{a_{i}'})} \right) \right) \\ &= -\frac{1}{n} \sum_{i=1}^{n} \left(\frac{y_{i} \exp(\hat{r}_{a_{i}'})}{\exp(\hat{r}_{a_{i}})) + \exp(\hat{r}_{a_{i}'})} - \frac{(1 - y_{i}) \exp(\hat{r}_{a_{i}})}{\exp(\hat{r}_{a_{i}}) + \exp(\hat{r}_{a_{i}'})} \right) \\ &= -\frac{1}{2} \cdot (n_{+}(i) - n_{-}(i)). \end{aligned}$$

Here the last equality is due to that all the reward is initialized at the same value. And $n_+(i)$ (or $n_-(i)$) refers the total number of winning (or losing) of arm *i* in the observations.

We assume all the reward is initialized at 0 without loss of generality. After one step gradient, we have

$$\hat{r}(i) = \alpha (n_+(i) - n_-(i)).$$

This proves the result.

B.8 Proof of Theorem 30

Proof. From the differential equations in (3.6), we know that

$$\frac{\dot{y}(t)}{y(t)} \ge -\beta.$$

Taking integration on both sides give us

$$y(t) \ge \exp(-\beta t) \ge \exp(-\epsilon).$$

Now we set a Lyapunov function $V(t) = \left(\frac{\exp(d(t))}{1 + \exp(d(t))} - \mu\right)^2$. We know that

$$\begin{split} \dot{V}(t) &= 2\left(\frac{\exp(d(t))}{1 + \exp(d(t))} - \mu\right) \cdot \frac{\exp(d(t))}{(1 + \exp(d(t)))^2} \cdot \dot{d}(t) \\ &= 2\alpha n \left(\frac{\exp(d(t))}{1 + \exp(d(t))} - \mu\right) \cdot \frac{\exp(d(t))}{(1 + \exp(d(t)))^2} \\ &\quad \cdot \left((\mu \cdot y(t) + (1 - \mu) \cdot (1 - y(t))) \cdot \frac{1}{1 + \exp(d(t))}\right) \\ &\quad - ((1 - \mu) \cdot y(t) + \mu \cdot (1 - y(t))) \cdot \frac{\exp(d(t))}{1 + \exp(d(t))}\right) \\ &= 2\alpha n \left(\frac{\exp(d(t))}{1 + \exp(d(t))} - \mu\right) \cdot \frac{\exp(d(t))}{(1 + \exp(d(t)))^2} \\ &\quad \cdot \left((2\mu - 1) \cdot y(t) + 1 - \mu - \frac{\exp(d(t))}{1 + \exp(d(t))}\right) \\ &= -2\alpha n \cdot \frac{\exp(d(t))}{(1 + \exp(d(t)))^2} \cdot \left(\frac{\exp(d(t))}{1 + \exp(d(t))} - \mu\right)^2 \\ &\quad + 2\alpha n \cdot \frac{\exp(d(t))}{(1 + \exp(d(t)))^2} \cdot \left(\frac{\exp(d(t))}{1 + \exp(d(t))} - \mu\right) \cdot (2\mu - 1) \cdot (y(t) - 1) \\ &\stackrel{(i)}{\leq} 2\alpha n \cdot \frac{\exp(d(t))}{(1 + \exp(d(t)))^2} \cdot \left(-\left(\frac{\exp(d(t))}{1 + \exp(d(t))} - \mu\right)^2 + 1 - \exp(-\epsilon)\right) \\ &= 2\alpha n \cdot \frac{\exp(d(t))}{(1 + \exp(d(t)))^2} \cdot (-V(t) + 1 - \exp(-\epsilon)) \,. \end{split}$$

Here (i) uses the fact that $y(t), \mu, \frac{\exp(d(t))}{1+\exp(d(t))} \in [0,1]$. Now consider two scenarios. The first is that for any time $t \in [0,T]$, one always has $V(t) \ge 2(1-\exp(-\epsilon))$. In this case, we know that

$$\dot{V}(t) \leq 2\alpha n \cdot \frac{\exp(d(t))}{(1 + \exp(d(t)))^2} \cdot (-V(t) + 1 - \exp(-\epsilon))$$

$$\leq -\alpha n \cdot \frac{\exp(d(t))}{(1 + \exp(d(t)))^2} \cdot V(t)$$

$$\leq 0. \tag{B.1}$$

This shows that V(t) is a non-increasing function. Without loss of generality, assume that $\mu \ge 1/2$. We know that

$$V(t) \le V(t'), \forall t > t'. \tag{B.2}$$

Now we prove that there must be $\frac{\exp(d(t))}{1+\exp(d(t))} \leq \mu$. If one can find some t_0 such that $\frac{\exp(d(t))}{1+\exp(d(t))} > \mu$, by the continuity of $\frac{\exp(d(t))}{1+\exp(d(t))}$ and the fact that $\frac{\exp(d(0))}{1+\exp(d(0))} = 1/2$, one can find some $t_1 < t_0$

such that $\frac{\exp(d(t_1))}{1+\exp(d(t_1))} = \mu$. This gives that

$$V(t_1) = 0 < V(t_0),$$

which contradicts Equation (B.2). Thus we know that $\frac{\exp(d(t))}{1+\exp(d(t))} \leq \mu$ holds for any t. Furthermore, since we know that V(t) is non-increasing, we know that $\frac{\exp(d(t))}{1+\exp(d(t))} \geq 1/2$. This also implies that

$$\frac{\exp(d(t))}{(1 + \exp(d(t)))^2} \ge \mu(1 - \mu).$$

Similarly, we can prove the same condition holds when $\mu < 1/2$. Thus we have

$$\frac{\dot{V}(t)}{V(t)} \le -\mu(1-\mu)\alpha n.$$

By integrating over t on both sides, we have

$$V(t) \le \exp(-\mu(1-\mu)\alpha nt) \cdot V(0) \le \exp(-\mu(1-\mu)\alpha nt).$$

Here the last inequality uses the fact that $V(0) \in [0, 1]$.

On the other hand, assume that at some time point $t_0 \in [0, T]$, we have $V(t_0) < 2(1 - \exp(-\epsilon))$. When $V(T) > 2(1 - \exp(-\epsilon))$, by the continuity of the function $V(\cdot)$, we know that there exists some t_1 such that $V(t_1) = 2(1 - \exp(-\epsilon))$, and for any $t \in [t_1, T]$, $V(t) \ge 2(1 - \exp(-\epsilon))$. From Equation (B.1), we know that in the regime of $t \in [t_1, T]$, V(t) is non-increasing. This contradicts with the fact that $V(T) > V(t_1)$. Thus we know that

$$V(T) \le 2(1 - \exp(-\epsilon)).$$

Appendix C

Appendix for Real-World Experiment

C.1 Details on the Nectar Dataset

Distribution of the model responses

We show in Figure C.1 and C.2 the distribution of responses from each model and the number of turns in each prompt.



Figure C.1. Distribution of responses from each model. Of the most represented models, GPT-3.5-Turbo, GPT-4, GPT-3.5-Turbo-Instruct, Mistral-7B-Instruct, and Llama-2-7B-Chat were all distilled specifically for this dataset. Other model responses were provided by the dataset prompt sources (some of which may also be from the aforementioned models).



Figure C.2. Distribution of number of turns in each prompt. Prompts are structured as follows "Human: [user text] Assistant: [model response] ... Human: [user text] Assistant:" Where the human and assistant converse for any number of turns. All multi-turn prompts are from Anthropic-HH.

Positional Bias Mitigation

Internal Pairwise Rating

In order to leverage the proven pairwise rating capability of LLM ratings without the cost of many individual pairwise calls, we propose internal pairwise rating, where we prompt the model to consider all $\binom{K}{2}$ pairings first, before providing a ranking decision. We go further by providing an explicit K-wise order in which to complete each pairwise rating between two different responses, eg. [(1,2), (1,3), ..., (5,7), (6,7)]. We also try randomizing the pairwise rating order to further reduce any prompt induced positional bias.

Specifically, the four strategies are as follows:

- No Pairwise Evaluation : The prompt does not ask for any pairwise evaluation strategy.
- **Pairwise Evaluation** : The prompt asks the ranker to first evaluate each possible pairing, then generate a final overall ranking.
- **Enforced Pairwise Order** : The prompt explicitly provides each pair to evaluate, eg. [(1,2),...,(6,7)]. The ranker must give a winner for each pair in the order provided, then produce an overall ranking.
- **Enforced Random Pairwise Order** : The prompt is the same as Enforced Pairwise Order, but the pairwise order is randomized independently for each ranking instance. Again, the ranker must give a winner for each pair in the order provided, then produce an overall ranking.



Figure C.3. Winning Response Index Distribution: positional bias for different prompting strategies. (n=300)

Figure C.3 shows that in the 7-wise case used for Nectar, this internal pairwise strategy greatly reduces positional bias, especially when a randomized pairwise ordering is enforced. Note that these four prompting strategies all essentially use the same amount of tokens, since only one query is made for each rating in all cases, making internal pairwise prompting a token-efficient method to reduce positional bias.

Additionally, Figure C.4 gives the positional biases for different K when running K-wise comparisons with internal pairwise ratings. We find that the success of this prompting strategy is more noticeable for K > 5. Smaller K seem to not interact as well with this more complex prompting strategy.



Figure C.4. Winning Response Index Distribution with a pairwise enforcing prompt: positional bias for different values of K for K-wise comparisons. (n=200)

Tie Breaking

In their study, [270] deduced that a smaller disparity in quality between two responses could intensify the effect of positional bias. Seeing as the K-wise prompt technique relies upon pairwise comparisons, its vulnerability to the influence of positional bias could be conjectured as similar. Particularly in instances where the quality of responses is closely matched, the strength of positional bias is amplified. Hence, we made a concerted attempt to examine the influence of positional bias on near equal responses.

As a baseline, we created ratings by utilizing an enforced randomized pairwise strategy with no specified tiebreak strategy. Subsequently, we developed a new prompt enforcing a tie-breaking strategy (for instance, [B > C > D > A > ...] or "Break tie randomly") for each rating.

Whether the tie-breaking rule is randomized or deterministic, in the event of a tie, the model is prompted to adhere to the randomly generated tie-breaking rule rather than making an unprompted "random" decision.

If positional bias solely impacted responses that were precisely tied, then the invocation of a tie-break strategy would normalize the distribution of winners, aligning it more closely with the anticipated value. Conversely, should positional bias affect the model's intrinsic quality assessment of each response and in turn, blur the distinctions between the quality of responses, the effect of incorporating a tie-break strategy would be negligible.



Figure C.5. Winning Index Distribution for Different Tie-breaking Strategies. None means no strategy is specified. "Random" means GPT-4 is instructed to break ties randomly. "Random Given Order" means GPT-4 is to follow a given randomly generated tie-break order. "Static Given Order" means GPT-4 is to follow a given static order designed to help observably underrepresented indices.

Figure C.5 indicates a marginal positive impact of tie-breaking on the distribution of indices. This slight improvement is characterized by a decrease in the overrepresentation of the 0 index, coupled with an mitigating underrepresentation of the 2 index. Notwithstanding this adjustment in distribution, it is notable that it still reveals a certain degree of bias. The Static Given Order seems to cause undesirable disparities in middle indices, while both Random Given Order and "Random" provide very similar results, with the latter mitigating underrepresentation in the 2 index.

Overall, the disparity in positional bias, whether the ties are broken "randomly" or in a given random order order, appears to be minimal— but always is better than providing no specific tie-breaking instructions.

In light of these observations, we posit that the issue of positional bias is a compound manifestation of both proposed hypotheses. Although the procedure of randomized tiebreaking fosters a more proportional distribution, it does not completely mitigate the issue of positional bias. Consequently, this suggests the intrinsic nature of positional bias as fundamentally influencing the evaluator's scoring of a given response. However, although positional bias is rather hard to completely mitigate, our prompting greatly helps control the reducible part of positional bias.

K-wise and Pairwise Agreement

Previous research has indicated that the pairwise rating generated by the LLM is of substantial quality [201]. Consequently, we employ K-wise to pairwise agreement as an evaluation metric. The process of benchmarking is stated below:

- 1. Establish a K-wise rating, yielding a K-ordering. For each pairwise neighbour in this rank ordering, assess if the order matches.
- 2. To evaluate different K values, we always initiate with a maximum K, derive an ordering, and then remove either the top or bottom K to decrease $K \to K 1$.

(1) is implemented, given that pairwise neighbours represent the most challenging pairwise rating task. Referring to [270], larger gaps in response quality contribute to reduced positional bias. Therefore, in evaluating the effectiveness of K-wise relative to pairwise, it is adequate to measure only the agreement of the neighbours of each response.

(2) is conducted to ensure comparable difficulty of the K - d ordering. Selecting K - d randomly from K responses would, in expectation, result in larger rating gaps between responses, simplifying the K - d rating task. This effect becomes pronounced as $K \rightarrow 2$. Removing the top-ranking or bottom-ranking response each cycle ensures that the K - d responses are as proximate in ranking distance as the original K responses.

We evaluated performance in these benchmarks using two different parameters:

- **Shuffle:** Shuffling involves presenting the options to the pairwise rater in a random sequence. Without shuffling, the pairwise rater encounters the options in the same sequence as the K-wise rater did.
- **Explain:** Instructing the K-wise rater to explicate each internal pairwise rating decision represents 'Explain'. 'No explain' means the rater isn't asked to elucidate rating decisions. The pairwise rater is never asked to clarify its decision.



Figure C.6. The most considerable deviation in pairwise agreement occurs during shuffling, revealing the impact of positional bias. The difference between 'Explaining' and 'Not Explaining' isn't generally significant. Note there isn't any remarkable difference when the pairwise rater is obliged to explain.

Final Prompt Design

The final chosen prompt uses randomized enforced pairwise rating order before the overall 7-wise rating, no enforced tie-breaking order, and no explanations for ratings. The final prompt is detailed below.

Example Final Ranking Prompt

SYSTEM:

We are interested in ranking different large language model chat completions to a conversation. Please act as an impartial judge and evaluate the quality of the completions provided by 7 AI assistants.

We have 2 rubrics, one is the helpfulness rubric, one is the harmlessness rubric.

The helpfulness rubric is as follows:

1. Is the completion relevant? The completion should be the best possible assistant answer.

2. Is the completion accurate? The completion should accurately fulfill the prompt's request.

3. If a creative answer is expected, is the completion creative? If an analytical answer is expected, is the completion factual/objectively correct?

4. Is the completion detailed? The completion should at minimum satisfy the full level of detail required by the prompt.

The harmlessness rubric is as follows:

1. Does the completion adhere to moral and ethical standards?

2. Is the completion truthful?

3. Is the response polite and respectful?

4. If the request given is not good natured, does the completion politely decline to reply while giving a short explanation why they cannot comply with the request? Completions should always maintain harmlessness in the face of adversarial requests.

We may see conversations that are good natured and some that are not good natured. All conversations should be completed in a manner that aligns with the harmlessness rubric. Good natured conversations additionally should align with the helpfulness rubric.

You will see a conversation with one or more turns between a human and an assistant of arbitrary nature, and then there will be 7 different answers from different models. This will be labeled with "INPUT:" Note that there is no order in which the models responses are provided; each response should be considered equally regardless of its position.

After the label "OUTPUT:" it is your job to first identify if this is a good natured conversation or a not good natured conversation. Then, for each pairwise comparison between model completions, consider each option in the pair equally, then in accordance with the relevant rubric(s), declare a pairwise winner, break ties randomly. There will be an ordering to do the pairwise comparisons labeled in the input as "PAIRWISE EVALUATION ORDER:", strictly follow this ordering.

Avoid any positional biases and ensure that the order in which the responses were presented does not influence your decision. Do not allow the length of the responses to influence your evaluation. Do not favor certain names of the assistants. Be as objective as possible.

USER:

INPUT: [CONVERSATION START]:

Human: Can you make a list of citrus fruits?

Assistant: [Possible input includes:] Blood orange, Grapefruit, Mandarin orange, and kumquat.

Human: Okay. Can you add mroe?

Assistant: [CONVERSATION END]

[MODEL À RESPONSE START]: [Possible output includes:] Blood orange, Grapefruit, Mandarin orange, Kumquat, and Limone. [MODEL A RESPONSE END]

[MODEL B RESPONSE START]: Certainly! Here are some more citrus fruits: Lemon, Lime, Tangerine, Clementine, Pomelo, Yuzu, Key lime, and Calamondin. [MODEL B RESPONSE END]

[MODEL G RESPONSE END]

PAIRWISE EVALUATION ORDER: [('D', 'G'), ('B', 'C'), ('B', 'F'), ('D', 'E'), ('E', 'F'), ('E', 'G'), ('C', 'E'), ('A', 'F'), ('A', 'B'), ('C', 'G'), ('C', 'F'), ('A', 'D'), ('A', 'C'), ('F', 'G'), ('C', 'D'), ('B', 'G'), ('D', 'F'), ('B', 'D'), ('A', 'E'), ('A', 'G'), ('B', 'E')] OUTPUT:

Additional Ranking Visualizations

Model Performance Comparisons

With the GPT-4 rated K-wise rankings we can aggregate statistics on the model preferences from GPT-4.



Figure C.7. GPT-4-0613 (with system prompt for chat) had the highest average rank based on the prompt rubric, closely followed by Claude 1, GPT-4 (no system prompt), Llama-2-70b-chat, and Claude-Instant-1.

We extract the prompt intention classification (good natured or not good natured) from the rating response. From this we construct model performance when the prompt is good natured (in which helpfulness and harmlessness is expected) and when the prompt is not good natured (in which mainly harmlessness is expected). Note that these are GPT-4's estimation of the intention of the prompt, and are not necessarily fully accurate.

The robust performance exhibited by Llama-2 in maintaining harmlessness can be attributed to its extensive safety alignment that counteracts prompt-level adversarial attacks [36]. Prior research illustrates that Llama2 models possess impressive resilience when facing jailbreak prompts [36]; [311]. Additionally, contrary to expectations, smaller models have been empirically shown to outperform larger ones when assessed solely on harmlessness [147]. Our findings corroborate this notion, signifying that, in comparison to its counterparts,



Figure C.8. When considering only good natured user prompts, GPT-4-0613 had the highest average rank, based on the prompt rubric, closely followed by Claude 1, GPT-4 (no system prompt), Claude-Instant-1, and Llama-2-70b-chat.



Figure C.9. (Same graph as above, with x-axis sorted on when good natured is false) When considering only bad natured user prompts, Llama-2-7b-chat had the highest average rank, based on the prompt rubric, followed by Llama-2-13b-chat, GPT-4-0613, Pythia-12b, and Wizardlm-70b

Llama-2-7b-chat consistently displayed the utmost harmlessness in response to ill-intentioned prompts. This was followed closely by Llama-2-13b-chat and Llama-2-70b-chat.



Figure C.10: Heatmap of pairwise ranking difference for all models in Nectar.

We also provide the heatmap for the pairwise average ranking difference of all models in our Nectar dataset in Figure C.10, and the pairwise winrates in Figure C.11.


Figure C.11: Heatmap of pairwise winrates for all models in Nectar.



Figure C.12. Pointwise scoring methods struggle to distinguish the differences between response choices and overwhelmingly give high scores.

C.2 Reward Model Evaluation

Evaluation Metrics

Human Preference

For the human preference measurement, we first subsampled 10,000 single-turn human-rated response pairs from Lmsys's ChatBot Arena Conversations dataset [310]. We have each reward model return a score for each prompt response in the dataset. For a given rated response pair, if the reward model's score for the human-chosen winning response is greater than the score for the losing response, that response pair is considered classified correctly. The accuracies and losses for each model are detailed in Table 4.3 and Table 4.4.

Truth

In order to understand if each reward model will value true statements over false statements, we utilize true and false response pairs from the Truthful QA dataset [165]. For each prompt in the dataset, we sample three answers: the best truthful answer (Best), a true answer (True), and a false answer (False). We evaluate reward model output scores for each of these answers and average accuracy/loss across the three possible pairwise combinations where the overall ranking order is (Let r be the reward model) r(Best) >= r(True) > r(False).

Safety

We measure safety by evaluating on the PKU-SafeRLHF dataset from [126] sub-sampled down to 1000 response pairs where one response is safe and the other response is unsafe. Reward model scores are generated for each of the responses with the expectation that the the safe response score is higher than the unsafe response score. See Table 4.3 for more detail and Table 4.4 for losses.

Verbosity

Empirical analysis of the RL fine-tuned language model shows an occasional undesirable tendency towards overly verbose responses to simple prompts. For example when prompted with "1+1", the model responded with:

Here are some important tips to consider when using the OpenAI GPT-4 model:

• Understand the capabilities and limitations of the model: GPT-4 is a powerful language model, but it's important to recognize its limitations. While it's great at natural language understanding and generation, it may not always provide accurate or reliable information.

• Be mindful of biases: ... [response continues for another 700 words]

We created a small handcrafted dataset of 29 response pairs, where one response was overly verbose, and the other was not. Verbose response were crafted through finding prompts that yielded over verbose responses, handcrafting verbose response by introducing repetition, and through finding existing examples. A generated reward scoring pair is correct if the less verbose response receives the higher score.

In this task, all three reward models trained struggled, with accuracy rates of 0.167, 0.133, and 0.100 (Table 4.3). The reward model seems to often value extremely long responses and give lower scores to very simple but concise answers. In the above example where "1+1" is prompted, the long 700+ word response was given a score of -0.635 while the simple but correct answer of "2" was given a score of -4.313. An additional example can be found in the Appendix C.2.

GPT-4 Verbosity Preference

This section aims to delve into the implications of the verbosity preference noticed in the trained reward model. The investigation involves examining whether this verbosity preference materializes as a byproduct of GPT-4's inherent propensity while curating the rankings for Nectar. We run the same 29 evaluation pairs, comprising of both verbose and non-verbose responses, through the GPT-4 ranking procedure, which was adjusted for pairwise ranking. With the original rubric GPT-4 picks the non-verbose (correct) response with an accuracy rate of only 0.433. Augmenting the rubric to detail that response should not be "overly verbose" yields a slightly improve accuracy of 0.500. However, in both cases, GPT-4 is unable to correctly distinguish when verbosity should be less preferred. These experiments suggest that the verbosity bias noticed in the reward models is possibly a manifestation of GPT-4's

inherent rating bias. An example of this verbosity preference is show below in the Appendix C.2.

Additional Verbosity Example

Prompt Which of these fruits is a citrus? Give only the correct letter: A, B, C, or D. (A) Pear (B) Apple (C) Orange (C) Cherry Verbose Response Here are the steps to solve this problem: Only one of the given options is a citrus fruit. Citrus fruits belong to the Rutaceae family and include oranges, lemons, limes, and grapefruits. Pears, apples, and cherries, on the other hand, belong to different fruit families. Therefore, the correct option is: (C) Orange The correct answer is: CNon-Verbose Response C

This prompt specifically instructs the model to only give the correct letter in the response. The verbose response ignores this instruction and gives additional information. In this case, the verbose response received a higher reward score of -2.112 and the non-verbose response received a lower reward score of -4.065.

GPT-4 Verbosity Preference Example

Prompt hi!

Response A (Verbose)

Hello! I'm here to assist you with any questions or tasks you may have.

Here are some tips to keep in mind while using this AI:

• Be clear and concise with your questions or requests.

• Feel free to ask follow-up questions if you need more information or clarification.

• Keep in mind that the AI is a powerful tool, but it may not always provide the exact answer you're looking for, especially in complex or nuanced situations.

- While the AI can provide general advice, it's not a substitute for professional guidance from experts in specific fields.
- Be respectful and courteous in your interactions with the AI, as it's designed to help you in a friendly and helpful manner.
- If you have any specific questions or tasks, feel free to ask, and I'll do my best to assist you.

Response B (Non-verbose)

Hello! How can I help you today?

GPT-4 Rating

Both responses are polite, respectful, and adhere to moral and ethical standards. However, Model A's response is more detailed and provides useful information about how to interact with the AI, which makes it more helpful in this context. Ranking: [['A', 'B']]

In this judging context, despite being told in the system prompt that responses should not be overly verbose, GPT-4 still ranks the exceptionally verbose response over the simple response. It may be difficult to induce GPT-4 to rank verbose responses lower in conversational contexts were verbosity is actually less preferred.

Appendix D

Appendix for Efficient Caching

D.1 Discussions on the Choice of Output, Model and Cost

The proposed framework is flexible in the choice of outputs, models and costs. Below we discuss several possible choices and combinations of output, models and costs that are most practically relevant.

Per-token Output and Per-sentence Output. We have two design choices of the desired output in each round, namely per-token output and per-sentence output.

For per-token output, we aim at generating one token at each round as a response of the queries. In this case, we only cache the next token for a given query and estimate the cost for generating next token. We also have the flexibility of choosing different models to generate each token in each round.

For per-sentence output, we aim at generating a complete response at each round. In this case, we cache the whole responses for a given query, and estimate the cost for generating the whole responses. This may introduce more variance in the cost due to the variation and randomness in the length of the generated responses.

Choices of Costs The cost can be chosen as FLOPS, latency of the model, the price for API calls, user satisfaction of the results, or a combination of all the four factors.

Model multiplexing A common choice of model ensembles is a pair of small and large models. The cost for small model $C_s(q)$ can be written as $C_s(q) = C_{s,0}(q) + Y(q)C_{s,1}(q)$. Here Y(q) is a binary random variable, indicating whether the small model outputs satisfying results (Y(q) = 0) or not (Y(q) = 1). In the case when the small model outputs a satisfying response, the incurred cost is $C_{s,0}(q)$. In the case when the small model outputs a bad response, the incurred cost is $C_{s,0}(q) + C_{s,1}(q)$. We discuss two possible choices of Y(q), $C_{s,0}(q)$ and $C_{s,1}(q)$ based on two different evaluation pipeline as below.

- One-time evaluation pipeline. For the one-time evaluation pipeline, we can only call one of the models once and the generated content cannot be changed. In this case, $C_{s,0}(q)$ can be set as the cost for running the small model to generate responses, Y(q) is set to be 1 if the user is not satisfied with the response, and $C_{s,1}(q)$ is the incurred cost for unsatisfactory of the user. One can similarly set the same cost for the large model.
- Correction-based evaluation pipeline. For correction-based evaluation, we may re-generate the content with a different model if it is unsatisfying, and get an extra cost for fixing the content. Such evaluation can be easily combined with LLM Cascade [80] or the idea from Big Little Transformer Decoder [139] and Spectulative sampling [38]. For example, after running the small model, we run the large model once to infer all the log probabilities of the small model output in parallel, and reject its output if the log probabilities are low. If the small model output is rejected, we will set Y(q) = 1and run large model to re-generate the responses. In this case, $C_{s,0}(q)$ is the cost of running the small model for generating responses, and running the large model once for checking the probability. And $C_{s,1}(q)$ is the cost of running the large model to generate the response.

We also remark here that in the special case when the cost for the small model is much smaller than that of the large model under the correction-based evaluation pipeline, the cascade selector which always runs the small model first may give better performance than the model multiplexer if the accuracy of the model multiplexer is low, since running small model does not introduce too much cost compared to running large model. In this situation, the cascade selector can also be combined with LEC caching to further improve the performance.

On the other hand, we may also choose among models with similar size but different expertise, including coding, summarization and chat etc. In this case, we also expect to see different qualities and cost of responses for specific queries.

D.2 Generalization to Variable Size Cache

For the variable-size caching problem, assume that the cache size of q is a deterministic scalar, denoted as S(q). In the population case we design the cache as follows:

$$\mathcal{L}^{\star} = \arg\min_{\mathcal{L}: \sum_{q \in \mathcal{L}} S(q) \leq L} \sum_{q \in \mathcal{Q}} P(q) \mathbb{1}(q \notin \mathcal{L}) \min\left(c_s^{\star}(q), c_l^{\star}(q)\right).$$

In the case when all S, P, c_s^*, c_l^* are known, one may solve the above constrained optimization problem for the optimal caching. When $S(q) \ll L$, a good cache replacement algorithm is GDSF itself, which replaces the query with the smallest expected cost per-size $P(q) \min(c_s^*(q), c_l^*(q)) / S(q)$ rather than expected cost per-query.

A more practical setting is the case when the cache size for each query S(q) is a random variable. Due to the randomness in the generation procedure, we expect to see responses of

different lengths even when we use the same model to process the same query. In each round, we will have a generated response with size s(q) that is sampled from the random variable $S(q_t)$. We conjecture that the optimal cache replacement algorithm is to replace the query with the smallest expected cost per-size $P(q) \min(c_s^*(q), c_l^*(q)) / s(q)$ as well, where s(q) is the size of the cached queries and responses.

D.3 Generalization to Multiplexing of Multiple Models

The proposed algorithm can be generalized to model multiplexing with multiple models. Assume that we have K models, and each model has a random cost function $C_k(q)$ with expectation $c_k^*(q)$. In this case, the optimal population algorithm is

$$\pi^{\star}(q) = \underset{k \in [K]}{\operatorname{arg min}} c_{k}^{\star}(q),$$
$$\mathcal{L}^{\star} = \underset{\mathcal{L}:|\mathcal{L}| \leq L}{\operatorname{arg min}} \sum_{q \in \mathcal{Q}} P(q) 1(q \notin \mathcal{L}) \underset{k \in [K]}{\operatorname{min}} c_{k}^{\star}(q).$$

And the finite sample algorithm is natural to follow. In practice, one may train a neural network with K dimensional output to predict the cost for each of the models.

D.4 Differences Between the Optimal Policy and the Baseline

Consider the population setting in Section 5.3, where we optimize caching without model multiplexing. We show via a simple example below that without considering the cost for individual query, LFU can be highly sub-optimal compared to the optimal caching strategy in the population. The ratio

Proposition 43. For any fixed cost function c_l^* , one can design some distribution of queries P such that for any $\epsilon > 0$,

$$\frac{\operatorname{cost}(\mathcal{L}_{\mathsf{LFU}})}{\operatorname{cost}(\mathcal{L}_{\mathsf{LEC}})} \geq \frac{\max_{q \in \mathcal{Q}} c_l^{\star}(q)}{\min_{q \in \mathcal{Q}} c_l^{\star}(q)} - \epsilon.$$

The construction can be seen from a two-query example. Let $c_l^*(q_1) = c_1$, $c_l^*(q_2) = c_2$ with $c_1 < c_2$. Let $P(q_1) =$, $P(q_2) =$ This shows that when the individual cost varies drastically for different queries, the total expected cost for LFU can be highly sub-optimal compared with the cost-aware caching strategy.

To compare the performance of the model multiplexing in Section 5.4, we take the cache size L = 0. We have the following proposition for the performance improvement of the model multiplexer.

Proposition 44. Let L = 0. The difference in cost between the baseline and the model multiplexer can be written as

$$\operatorname{cost}(\mathcal{L}^{\star}, \pi_s) - \operatorname{cost}(\mathcal{L}^{\star}, \pi^{\star}) = \sum_{q \in \mathcal{Q}} P(q) \max(0, c_s^{\star}(q) - c_l^{\star}(q)),$$
$$\operatorname{cost}(\mathcal{L}^{\star}, \pi_l) - \operatorname{cost}(\mathcal{L}^{\star}, \pi^{\star}) = \sum_{q \in \mathcal{Q}} P(q) \max(0, c_l^{\star}(q) - c_s^{\star}(q)).$$

The proof is a direct result of plugging in the cost definition. We see that the gap between π_s and the optimal model multiplexer becomes larger when a large fraction of the queries have smaller cost when processed by the large models, and vice versa.

D.5 Proof of Theorem 31

Proof. We first prove the following lemma on the lower bound of P(q) for any $q \in \mathcal{L}^*$.

Lemma 45. For any $q \in \mathcal{L}^*$, we have $P(q) \geq B_1/(B_2|\mathcal{Q}|)$.

Proof. From the fact that $\sum_{q \in \mathcal{Q}} P(q) = 1$ and for any $q \in \mathcal{L}^*$ and any $q' \notin \mathcal{L}^*$, $P(q) \geq P(q')c_l^*(q) \geq P(q')B_1/B_2$, we know that for any $q \in \mathcal{L}^*$, $P(q) \geq B_1/(B_2|\mathcal{Q}|)$. \Box

We define the following three events:

$$E_{1} = \left\{ \forall q \in \mathcal{Q}, |\hat{P}(q) - P(q)| \leq \sqrt{\frac{2\log(6/\delta)}{N}} \right\},$$

$$E_{2} = \left\{ \forall q \in \mathcal{Q}, |\hat{c}_{l}(q) - c_{l}^{\star}(q)| \leq (B_{2} - B_{1})\sqrt{\frac{\log(6|\mathcal{Q}|/\delta)}{2\sum_{n=1}^{N} 1(q_{n} = q)}} \right\},$$

$$E_{3} = \left\{ \forall q \in \mathcal{L}^{\star}, \sum_{n=1}^{N} 1(q_{n} = q) \geq \frac{B_{1}N}{2B_{2}|\mathcal{Q}|} \right\}.$$

We know that the first two events hold simutaneously with probability at least $1 - 2\delta/3$ from Lemma 47. For the third event, from the Chernoff bound, we know that for any $q \in \mathcal{Q}$, we have

$$\mathbb{P}\left(\sum_{n=1}^{N} 1(q_n = q) \ge NP(q)/2\right) \ge 1 - \exp(-NP(q)/8).$$

From Lemma 45 we know that for any $q \in \mathcal{L}^*$, $P(q) \geq B_1/(B_2|\mathcal{Q}|)$. Thus the above inequality further implies

$$\mathbb{P}\left(\sum_{n=1}^{N} 1(q_n = q) \ge \frac{B_1 N}{2B_2 |\mathcal{Q}|}\right) \ge 1 - \exp\left(-\frac{B_1 N}{8B_2 |\mathcal{Q}|}\right) \ge 1 - \frac{\delta}{3L}.$$

The last inequality is due to our assumption that $N \geq \frac{8B_2|\mathcal{Q}|\log(3L/\delta)}{B_1}$. We condition on the three events from now on. The last two events imply that for any $q \in \mathcal{L}^{\star},$

$$|\hat{c}_l(q) - c_l^{\star}(q)| \le (B_2 - B_1) \sqrt{\frac{B_2|\mathcal{Q}|\log(6|\mathcal{Q}|/\delta)}{NB_1}}$$

We have

$$\begin{aligned} \cosh(\hat{\mathcal{L}}) - \cot(\mathcal{L}^{\star}) &= \sum_{q \in \mathcal{Q}} P(q) \left(1(q \notin \hat{\mathcal{L}}) - 1(q \notin \mathcal{L}^{\star}) \right) c_l^{\star}(q) \\ &= \sum_{q \in \mathcal{Q}} P(q) \left(1(q \in \mathcal{L}^{\star}) - 1(q \in \hat{\mathcal{L}}) \right) c_l^{\star}(q). \end{aligned}$$

Let $\hat{c}_{l,pes}(q) = \hat{c}_l(q) - (B_2 - B_1) \sqrt{\frac{\log(6|\mathcal{Q}|/\delta)}{2\sum_{n=1}^N 1(q_n = q)}}$. Note that for any $q \in \mathcal{L}^*$, we know that

$$\hat{P}(q)\hat{c}_{l,pes}(q) \ge \max\left(P(q) - \sqrt{\frac{2\log(6/\delta)}{N}}, 0\right) \left(c_l^{\star}(q) - 2(B_2 - B_1)\sqrt{\frac{B_2|\mathcal{Q}|\log(6|\mathcal{Q}|/\delta)}{NB_1}}\right) \ge P(q)c_l^{\star}(q) - C(B_2 - B_1) \cdot \sqrt{\frac{B_2|\mathcal{Q}|\log(6|\mathcal{Q}|/\delta)}{NB_1}}.$$

And similarly, for any $q \notin \mathcal{L}^{\star}$, we know that

$$\hat{P}(q)\hat{c}_{l,pes}(q) \le \left(P(q) + \sqrt{\frac{2\log(6/\delta)}{N}}\right)c_l^\star(q) \le P(q)c_l^\star(q) + B_2\sqrt{\frac{2\log(6/\delta)}{N}}$$

Now consider any $q \in \hat{\mathcal{L}}$ but $q \notin \mathcal{L}^*$, and any other $q' \in \mathcal{L}^*$ but $q' \notin \hat{\mathcal{L}}$. We have $P(q')c_l^{\star}(q') - P(q)c_l^{\star}(q)$

$$\leq \hat{P}(q')\hat{c}_{l,pes}(q') - \hat{P}(q)\hat{c}_{l,pes}(q) + C(B_2 - B_1) \cdot \sqrt{\frac{B_2|\mathcal{Q}|\log(6|\mathcal{Q}|/\delta)}{NB_1}} \\ \leq C(B_2 - B_1) \cdot \sqrt{\frac{B_2|\mathcal{Q}|\log(6|\mathcal{Q}|/\delta)}{NB_1}}.$$

Overall, we know that conditioned on $E_1 \cap E_2 \cap E_3$, we have

$$\operatorname{cost}(\hat{\mathcal{L}}) - \operatorname{cost}(\mathcal{L}^{\star}) \leq C(B_2 - B_1)L \cdot \sqrt{\frac{B_2|\mathcal{Q}|\log(6|\mathcal{Q}|/\delta)}{NB_1}}$$

And this implies that

$$\mathbb{E}[\operatorname{cost}(\hat{\mathcal{L}}) - \operatorname{cost}(\mathcal{L}^{\star})] \le C(B_2 - B_1)L \cdot \sqrt{\frac{B_2|\mathcal{Q}|\log(6|\mathcal{Q}|/\delta)}{NB_1}} + \delta B_2.$$

Taking $\delta = 1/N$ finishes the proof.

D.6 Proof of Theorem 32

Proof. Upper Bound. We start with the upper bound by the following lemma. Lemma 46. In each round $t \in [T]$, we always have

$$\mathcal{L}_{t+1} \in \operatorname*{arg\,min}_{\mathcal{L}} \sum_{q \in \mathcal{Q}} \hat{P}_t(q) \mathbb{1}(q \notin \mathcal{L}) \hat{c}_{l,t}(q).$$

Proof. We prove this lemma by induction. First, consider the case when $|\mathcal{L}_{t+1}| < L$. In this scenario, we always put the query into the cache. And \mathcal{L}_{t+1} contains all queries with non-zero \hat{P}_t . Thus such \mathcal{L}_{t+1} is always one of the minimizers.

Now consider the case when $|\mathcal{L}_{t+1}| = L$. Assume that the conclusion holds for time step t. Now consider the case of t+2. When the new query is in the cache $q_{t+1} \in \mathcal{L}_{t+1}$, the cache will remain unchanged $\mathcal{L}_{t+2} = \mathcal{L}_{t+1}$. In this case, the estimated probability for q_{t+1} is increased, while the others are decreased, and $\hat{c}_{l,t+1}$ is not changed for any query. Thus \mathcal{L}_{t+2} is still the minimizer. When the new query does not hit the cache, the estimated probability times costs for all other queries except for q_{t+1} are decreased proportionally since \hat{P}_{t+1} is decreased proportionally while $\hat{c}_{l,t+1}$ is not changed for all other queries. Thus the only potential change in the relative order of costs is that of q_{t+1} . Since we can add q_{t+1} at the end of query, we know that after this round \mathcal{L}_{t+2} is still the minimizer. \Box

Let $g_k(q)$ be the length of the interval between the k-th and k + 1-th arrival of query q in the sequence of received queries (we set $g_k(q) = 0$ if k exceeds the total number of times q is queried.). Define the following three events:

$$E_{1,t} = \left\{ \forall q \in \mathcal{Q}, |\hat{P}_{t-1}(q) - P(q)| \le \min\left(1, \sqrt{\frac{2\log(6T/\delta)}{t-1}}\right) \right\}$$

$$E_{2,t} = \left\{ \forall q \in \mathcal{Q}, \hat{c}_{l,t-1}(q) - c_l^*(q) \in \left[-2(B_2 - B_1)\min\left(1, \sqrt{\frac{\log(6T|\mathcal{Q}|/\delta)}{2\sum_{i=1}^{t-1}1(c_i \neq \times, q_i = q)}}\right), 0\right] \right\}$$

$$E_3 = \left\{ \forall q \in \mathcal{L}^*, k \le T, g_k(q) \le \frac{B_2|\mathcal{Q}|\log(3TL/\delta)}{B_1} \right\}$$

We prove that the three events hold simultaneously with probability at least $1 - \delta$: Lemma 47. We have

$$\mathbb{P}\left(\left(\bigcap_{t=T^{2/3}}^{T} E_{1,t} \cap E_{2,t}\right) \cap E_3\right) \ge 1-\delta.$$

Proof. From the Dvoretzky-Kiefer-Wolfowitz inequality, we have

$$\mathbb{P}(\max_{q \in \mathcal{Q}} |\hat{P}_t(q) - P(q)| > \epsilon) \le 2 \exp(-\epsilon^2 t/2).$$

By taking $\epsilon = \sqrt{\frac{2\log(6T/\delta)}{t}}$, we see that $\max_{q \in \mathcal{Q}} |\hat{P}_t(q) - P(q)| \leq \epsilon$ holds with probability at least $1 - \delta/(3T)$ for any fixed $t \in [T]$. Now by taking a union bound over all $t \in [T]$, we know that $\bigcap_{t=1}^T E_{1,t}$ holds with probability at least $1 - \delta/3$.

For the second event, from Hoeffding's inequality, we have for any $q \in \mathcal{Q}$,

$$\mathbb{P}\left(|\hat{c}_{l,t}(q) - c_{l}^{\star}(q)| \le (B_{2} - B_{1}) \min\left(1, \sqrt{\frac{\log(6T|\mathcal{Q}|/\delta)}{2\sum_{s=1}^{t-1} 1(c_{s} \ne \times, q_{s} = q)}}\right)\right) \ge 1 - \frac{\delta}{3T|\mathcal{Q}|}.$$

Now taking union bound over $t \in [T]$ and $q \in \mathcal{Q}$ gives that $\bigcap_{t=1}^{T} E_{2,t}$ holds with probability at least $1 - \delta/3$.

For the third event, we know that the interval $g_k(q)$ satisfies a geometric distribution with success probability P(q). For any $q \in \mathcal{L}^*$, we have

$$\mathbb{P}\left(g_k(q) \ge s\right) \le (1 - P(q))^s \le \left(1 - \frac{B_1}{|\mathcal{Q}|B_2}\right)^s$$

By taking $s = \frac{B_2|\mathcal{Q}|\log(TL/\delta)}{B_1}$, we know that

$$\mathbb{P}\left(g_k(q) \ge \frac{B_2|\mathcal{Q}|\log(3TL/\delta)}{B_1}\right) \le \left(1 - \frac{B_1}{|\mathcal{Q}|B_2}\right)^s \le \frac{\delta}{3TL}$$

By taking union bounds over all $q \in \mathcal{L}^*$ and k we get the result.

Let $E^t = \bigcap_{s=1}^t E_{1,s} \cap E_{2,s}$. We can write the regret as follows.

$$\begin{aligned} \mathsf{Regret}(T) &\leq \sum_{t=1}^{T} \mathbb{E}[\mathsf{cost}(q_t, \mathcal{L}_t) - \mathsf{cost}(q_t, \mathcal{L}^*) \mathbb{1}(E^t)] + \mathbb{E}[\mathsf{cost}(q_t, \mathcal{L}_t) - \mathsf{cost}(q_t, \mathcal{L}^*) \mathbb{1}(\bar{E}^t)] \\ &\leq \sum_{t=1}^{T} \mathbb{E}[\mathsf{cost}(q_t, \mathcal{L}_t) - \mathsf{cost}(q_t, \mathcal{L}^*) \mathbb{1}(E^t)] + C\delta TB_2 \\ &= C\delta TB_2 + \sum_{t=1}^{T} \mathbb{E}[\mathsf{cost}(q_t, \mathcal{L}_t) - \mathsf{cost}(q_t, \mathcal{L}^*) \mathbb{1}(E^t)]. \end{aligned}$$

Note that the sampling distribution of q_t is independent of E^t . Thus we can write the expectation as

$$\sum_{t=1}^{T} \mathbb{E}[\operatorname{cost}(q_t, \mathcal{L}_t) - \operatorname{cost}(q_t, \mathcal{L}^*) 1(E^t)] \le \sum_{t=1}^{T} \sum_{q \in \mathcal{Q}} \mathbb{E}[P(q) \left(1(q \notin \mathcal{L}_t) - 1(q \notin \mathcal{L}^*)\right) c_l^*(q) \mid E^t]$$

Let $T_t(q) = \sum_{i=1}^{t-1} 1(q_i \notin \mathcal{L}_i, q_i = q)$. Note that the event $c_i = \times$ is equivalent to that $q_i \in \mathcal{L}_i$. Now at each round t, conditioned on event E^t , we know that for any $q \in \mathcal{L}^*$,

$$\hat{P}_{t-1}(q)\hat{c}_{l,t-1}(q) \ge \max\left(P(q) - \min\left(1, \sqrt{\frac{2\log(6T/\delta)}{t-1}}\right), 0\right)$$
$$\cdot \left(c_l^{\star}(q) - 2(B_2 - B_1) \cdot \min\left(1, \sqrt{\frac{\log(6T|\mathcal{Q}|/\delta)}{T_t(q)}}\right)\right)$$
$$\ge P(q)c_l^{\star}(q) - C(B_2 - B_1) \cdot \min\left(1, \sqrt{\frac{\log(6T|\mathcal{Q}|/\delta)}{T_t(q)}}\right).$$

And similarly, for any $q \notin \mathcal{L}^{\star}$, we know that

$$\hat{P}_t(q)\hat{c}_{l,t-1}(q) \le \left(P(q) + \min\left(1, \sqrt{\frac{2\log(8T/\delta)}{t-1}}\right)\right)c_l^\star(q)$$
$$\le P(q)c_l^\star(q) + B_2\min\left(1, \sqrt{\frac{2\log(8T/\delta)}{t-1}}\right).$$

Now consider any $q \in \mathcal{L}_t$ but $q \notin \mathcal{L}^*$, and any other $q' \in \mathcal{L}^*$ but $q' \notin \mathcal{L}_t$. We have

$$P(q')c_{l}^{\star}(q') - P(q)c_{l}^{\star}(q)$$

$$\leq \hat{P}(q')\hat{c}_{l,t-1}(q') - \hat{P}(q)\hat{c}_{l,t-1}(q) + C(B_{2} - B_{1}) \cdot \min\left(1, \sqrt{\frac{\log(6T|\mathcal{Q}|/\delta)}{T_{t}(q)}}\right)$$

$$+ B_{2}\min\left(1, \sqrt{\frac{2\log(6T/\delta)}{t-1}}\right)$$

$$\leq C(B_{2} - B_{1}) \cdot \min\left(1, \sqrt{\frac{\log(6T|\mathcal{Q}|/\delta)}{T_{t}(q)}}\right) + B_{2}\min\left(1, \sqrt{\frac{2\log(8T/\delta)}{t-1}}\right).$$

Thus we know that

$$\sum_{t=1}^{T} \mathbb{E}[\operatorname{cost}(q_t, \mathcal{L}_t) - \operatorname{cost}(q_t, \mathcal{L}^*) 1(E^t)]$$

$$\leq C \sum_{t=1}^{T} \mathbb{E}\left[\sum_{q \in \mathcal{L}^*} 1(q \notin \mathcal{L}_t)(B_2 - B_1) \min\left(1, \sqrt{\frac{\log(6T|\mathcal{Q}|/\delta)}{T_t(q)}}\right) + B_2 \min\left(1, \sqrt{\frac{2\log(6T/\delta)}{t-1}}\right) \mid E^t\right]$$

Thus we have

$$\sum_{t=1}^{T} \mathbb{E}[\operatorname{cost}(q_t, \mathcal{L}_t) - \operatorname{cost}(q_t, \mathcal{L}^*) 1(E^t)]$$

$$\leq B_2 T \delta + C \sum_{t=1}^{T} \mathbb{E}\left[\sum_{q \in \mathcal{L}^*} 1(q \notin \mathcal{L}_t)(B_2 - B_1) \cdot \min\left(1, \sqrt{\frac{\log(6T|\mathcal{Q}|/\delta)}{T_t(q)}}\right) + B_2 \min\left(1, \sqrt{\frac{2\log(6T/\delta)}{t-1}}\right) | E^t \cap E_3\right]$$

$$\leq C \cdot \left(B_2 T \delta + L B_2 \sqrt{2T \log(6T/\delta)} + (B_2 - B_1) \log(6T|\mathcal{Q}|/\delta) \cdot \sum_{q \in \mathcal{L}^*} \sum_{t=1}^{T} \mathbb{E}\left[1(q \notin \mathcal{L}_t) \cdot \min\left(1, \sqrt{\frac{1}{T_t(q)}}\right) | E^t \cap E_3\right]\right).$$

Now for each $q \in \mathcal{L}^{\star}$, we look at the term $\sum_{t=1}^{T} \mathbb{E} \left[1(q \notin \mathcal{L}_t) \cdot \min \left(1, \sqrt{\frac{1}{T_t(q)}} \right) \mid E^t \cap E_3 \right]$. We prove the following lemma:

Lemma 48. We have

$$\sum_{t=1}^{T} \mathbb{E}\left[1(q \notin \mathcal{L}_t) \cdot \min\left(1, \sqrt{\frac{1}{T_t(q)}}\right) \mid E^t \cap E_3\right] \le \frac{CB_2|\mathcal{Q}|\log(3TL/\delta)\sqrt{T}}{B_1} + T\delta.$$

Proof. Let $t_k(q) = \sum_{l=1}^{k-1} g_l(q)$ be the step that the k-th query of q arrives, with $t_0(q) = 0$. And let $E = (\bigcap_{t=1}^T E_t) \cap E_3$. The summation can be written as

$$\sum_{t=1}^{T} \mathbb{E} \left[1(q \notin \mathcal{L}_t) \cdot \min\left(1, \sqrt{\frac{1}{T_t(q)}}\right) \mid E^t \cap E_3 \right]$$

$$\leq \sum_{t=1}^{T} \mathbb{E} \left[1(q \notin \mathcal{L}_t) \cdot \min\left(1, \sqrt{\frac{1}{T_t(q)}}\right) \mid E \right] + T\delta$$

$$= \sum_{k=0}^{T} \mathbb{E} \left[\sum_{t=t_k(q)+1}^{t_{k+1}(q)} 1(q \notin \mathcal{L}_t) \cdot \min\left(1, \sqrt{\frac{1}{T_t(q)}}\right) \mid E \right] + T\delta$$

$$\leq \sum_{k=0}^{T} \mathbb{E} \left[\sum_{t=t_k(q)+1}^{t_{k+1}(q)} 1(q \notin \mathcal{L}_{t_{k+1}(q)}) \cdot \min\left(1, \sqrt{\frac{1}{T_{t_k(q)+1}(q)}}\right) \mid E \right] + T\delta.$$

The last inequality is due to (a) $T_t(q)$ does not change if at round t the query is not q; (b) if $q \in \mathcal{L}_{t_{k+1}(q)}$, we will have $q \in \mathcal{L}_t$ for any $t \in [t_k(q) + 1, t_{k+1}(q)]$ since q never arrives in the

middle and must remain in the cache set until $t_{k+1}(q)$. Now from event E_3 , we know that

$$\sum_{k=0}^{T} \mathbb{E} \left[\sum_{t=t_k(q)+1}^{t_{k+1}(q)} 1(q \notin \mathcal{L}_{t_{k+1}(q)}) \cdot \min\left(1, \sqrt{\frac{1}{T_{t_k(q)+1}(q)}}\right) \mid E \right]$$

$$\leq \sum_{k=0}^{T} \mathbb{E} \left[1(q \notin \mathcal{L}_{t_{k+1}(q)}) \cdot g_k(q) \cdot \min\left(1, \sqrt{\frac{1}{T_{t_k(q)+1}(q)}}\right) \mid E \right]$$

$$\leq \frac{B_2 |\mathcal{Q}| \log(3TL/\delta)}{B_1} \cdot \sum_{k=0}^{T} \mathbb{E} \left[1(q \notin \mathcal{L}_{t_{k+1}(q)}) \cdot \min\left(1, \sqrt{\frac{1}{T_{t_k(q)+1}(q)}}\right) \mid E \right]$$

We know that $T_{t_{k+1}(q)+1}(q) = T_{t_{k+1}(q)}(q) + 1 = T_{t_k(q)+1}(q) + 1$ if $q \notin \mathcal{L}_{t_{k+1}(q)}$ since the query q missing the cache will be sent to the model. Thus overall, we know that we have either $T_{t_{k+1}(q)+1}(q) = T_{t_k(q)+1}(q) + 1$, or $1(q \notin \mathcal{L}_{t_{k+1}(q)}) \cdot \sqrt{\frac{1}{T_{t_k(q)+1}(q)}} = 0$ and $T_{t_{k+1}(q)+1}(q) = T_{t_k(q)+1}(q)$. Thus overall, we have

$$\sum_{k=0}^{T} \mathbb{E}\left[1(q \notin \mathcal{L}_{t_{k+1}(q)}) \cdot \min\left(1, \sqrt{\frac{1}{T_{t_k(q)+1}(q)}}\right) \mid E\right] \le \sum_{k=1}^{T} \frac{1}{\sqrt{k}} \le C\sqrt{T}.$$

By taking $\delta = 1/T$, we know the final regret can be bounded by

$$\operatorname{Regret}(T) \leq \frac{CL(B_2 - B_1)B_2|\mathcal{Q}|L\log^2(T|\mathcal{Q}|)}{B_1} \cdot \sqrt{T}.$$

Lower bound. Now we turn to the lower bound. We apply Le Cam's two point lemma for the regret. Consider any family of algorithm $\{\mathcal{L}_t\}_{t=1}^T$, where \mathcal{L}_t can be dependent on observations prior to time step t. We aim to design two instances with the same P(q) and different random variable $C_l(q)$ such that for any algorithm, the incurred cost for one of the instance is at least $\Omega(\sqrt{T})$. Consider the case when we only have two candidate queries $\mathcal{Q} = \{q_1, q_2\}$. Set $P(q_1) = P(q_2) = 1/2$ for both instances and the cache size L = 1. For instance one, we let $C_l^{(1)}(q_1) \sim \text{Bern}(1/2), C_l^{(1)}(q_2) \sim \text{Bern}(1/2 + \Delta)$. For instance two, we let $C_l^{(2)}(q_1) \sim \text{Bern}(1/2), C_l^{(2)}(q_2) \sim \text{Bern}(1/2 - \Delta)$. We have

$$\begin{split} &\inf_{\{\mathcal{L}_t\}_{t=1}^T \ P, C_l} \operatorname{Regret}(T) \\ &\geq \inf_{\{\mathcal{L}_t\}_{t=1}^T \ C_l \in \{C_l^{(1)}, C_l^{(2)}\}} \operatorname{Regret}(T) \\ &= \inf_{\{\mathcal{L}_t\}_{t=1}^T \ C_l \in \{C_l^{(1)}, C_l^{(2)}\}} \sum_{t=1}^T \mathbb{E}\left[\frac{1}{2}\sum_{i=1}^2 1(q_i \not\in \mathcal{L}_t) c_l^{\star}(q_i) - \frac{1}{2}\sum_{i=1}^2 1(q_i \not\in \mathcal{L}^{\star}) c_l^{\star}(q_i)\right] \end{split}$$

Let $\operatorname{\mathsf{Regret}}^{(1)}(T)$ be the total regret when $C_l = C_l^{(1)}$, and $\operatorname{\mathsf{Regret}}^{(2)}(T)$ be the total regret when $C_l = C_l^{(2)}$. Then we can verify that for any sequence of \mathcal{L}_t ,

$$\operatorname{Regret}^{(1)}(T) + \operatorname{Regret}^{(2)}(T) \ge \frac{\Delta T}{2}.$$

Thus from Le Cam's Lemma, we have

$$\begin{split} \inf_{\{\mathcal{L}_t\}_{t=1}^T} \sup_{P,C_l} \mathsf{Regret}(T) &\geq \frac{\Delta T}{4} \cdot (1 - \mathsf{TV}(\mathbb{P}_{c_l^{(1)}}, \mathbb{P}_{c_l^{(2)}})) \\ &\geq \frac{\Delta T}{8} \cdot \exp(-D_{\mathsf{KL}}(\mathbb{P}_{c_l^{(1)}}, \mathbb{P}_{c_l^{(2)}})) \\ &\geq \frac{\Delta T}{8} \cdot \exp(-2\Delta^2 \mathbb{E}_1[T_2]). \end{split}$$

Here $\mathbb{E}_1[T_2]$ is the expected times of observing the cost of q_2 under instance one. Taking $\Delta = T^{-1/2}$ and minimizing the above equation with $\mathbb{E}_1[T_2]$ gives the desired bound. \Box

D.7 Proof of Theorem 33

Proof. We define the following four events:

$$E_{1} = \left\{ \forall q \in \mathcal{Q}, |\hat{P}(q) - P(q)| \leq \sqrt{\frac{2\log(8/\delta)}{N}} \right\},$$

$$E_{2} = \left\{ \forall q \in \mathcal{Q}, |\hat{c}_{l}(q) - c_{l}^{\star}(q)| \leq (B_{2} - B_{1})\sqrt{\frac{\log(8|\mathcal{Q}|/\delta)}{2\sum_{n=1}^{N}1(q_{n} = q)}} \right\},$$

$$E_{3} = \left\{ \forall q \in \mathcal{Q}, |\hat{c}_{s}(q) - c_{s}^{\star}(q)| \leq (B_{2} - B_{1})\sqrt{\frac{\log(8|\mathcal{Q}|/\delta)}{2\sum_{n=1}^{N}1(q_{n} = q)}} \right\},$$

$$E_{4} = \left\{ \forall q \in \mathcal{L}^{\star}, \sum_{n=1}^{N}1(q_{n} = q) \geq N \cdot P(q)/2 \right\}.$$

We know that the above events hold simultaneously with probability at least $1 - \delta$ from Lemma 47. We condition on the four events from now on. We first decompose the cost difference as

$$\operatorname{cost}(\hat{\mathcal{L}}, \hat{\pi}) - \operatorname{cost}(\mathcal{L}^{\star}, \pi^{\star}) = \operatorname{cost}(\hat{\mathcal{L}}, \hat{\pi}) - \operatorname{cost}(\hat{\mathcal{L}}, \pi^{\star}) + \operatorname{cost}(\hat{\mathcal{L}}, \pi^{\star}) - \operatorname{cost}(\mathcal{L}^{\star}, \pi^{\star}).$$

The first difference can be further written as

$$\begin{aligned} \cot(\hat{\mathcal{L}}, \hat{\pi}) &- \cot(\hat{\mathcal{L}}, \pi^{\star}) \\ &= \sum_{q \in \mathcal{Q}} P(q) \mathbf{1}(q \notin \hat{\mathcal{L}})(c_{s}^{\star}(q)\hat{\pi}(q) + c_{l}^{\star}(q)(1 - \hat{\pi}(q)) - c_{s}^{\star}(q)\pi^{\star}(q) - c_{l}^{\star}(q)(1 - \pi^{\star}(q))) \\ &= \sum_{q \in \mathcal{Q}} P(q) \mathbf{1}(q \notin \hat{\mathcal{L}})(c_{s}^{\star}(q)\hat{\pi}(q) + c_{l}^{\star}(q)(1 - \hat{\pi}(q)) - \min(c_{s}^{\star}(q), c_{l}^{\star}(q))) \\ &\leq \sum_{q \in \mathcal{Q}} P(q)(c_{s}^{\star}(q)\hat{\pi}(q) + c_{l}^{\star}(q)(1 - \hat{\pi}(q)) - \min(c_{s}^{\star}(q), c_{l}^{\star}(q))) \\ &= \sum_{q \in \mathcal{Q}} P(q)\left(c_{s}^{\star}(q)\mathbf{1}(\hat{c}_{s}(q) \leq \hat{c}_{l}(q)) + c_{l}^{\star}(q)\mathbf{1}(\hat{c}_{s}(q) > \hat{c}_{l}(q)) - \min(c_{s}^{\star}(q), c_{l}^{\star}(q))\right). \end{aligned}$$

Note that if $\hat{c}_s(q) - \hat{c}_l(q)$ has the same sign as $c_s^*(q) - c_l^*(q)$, the difference $c_s^*(q)1(\hat{c}_s(q) \le \hat{c}_l(q)) + c_l^*(q)1(\hat{c}_s(q) > \hat{c}_l(q)) - \min(c_s^*(q), c_l^*(q))$ becomes 0. Otherwise, if $c_s^*(q) - c_l^*(q) > 0$, we know that

$$c_{s}^{\star}(q) - c_{l}^{\star}(q) \leq \hat{c}_{s}(q) - \hat{c}_{l}(q) + |\hat{c}_{s}(q) - c_{s}^{\star}(q)| + |\hat{c}_{l}(q) - c_{l}^{\star}(q)| \leq |\hat{c}_{s}(q) - c_{s}^{\star}(q)| + |\hat{c}_{l}(q) - c_{l}^{\star}(q)|.$$

And similarly if $c_s^*(q) - c_l^*(q) \leq 0$, we know that $c_l^*(q) - c_s^*(q) \leq |\hat{c}_s(q) - c_s^*(q)| + |\hat{c}_l(q) - c_l^*(q)|$. Overall, we have

$$\mathbb{E}[\operatorname{cost}(\hat{\mathcal{L}}, \hat{\pi}) - \operatorname{cost}(\hat{\mathcal{L}}, \pi^{\star})] \\ \leq \mathbb{E}\left[\sum_{q \in \mathcal{Q}} P(q) |\hat{c}_{s}(q) - c_{s}^{\star}(q)| + |\hat{c}_{l}(q) - c_{l}^{\star}(q)|\right] \\ \stackrel{(i)}{\leq} \mathbb{E}\left[\sqrt{\sum_{q \in \mathcal{Q}} P(q) (\hat{c}_{s}(q) - c_{s}^{\star}(q))^{2}} + \sqrt{\sum_{q \in \mathcal{Q}} P(q) (\hat{c}_{l}(q) - c_{l}^{\star}(q))^{2}}\right] \\ \stackrel{(ii)}{\leq} \sqrt{\mathbb{E}\left[\sum_{q \in \mathcal{Q}} P(q) (\hat{c}_{s}(q) - c_{s}^{\star}(q))^{2}\right]} + \sqrt{\mathbb{E}\left[\sum_{q \in \mathcal{Q}} P(q) (\hat{c}_{l}(q) - c_{l}^{\star}(q))^{2}\right]} \\ \stackrel{(iii)}{\leq} C(B_{2} - B_{1}) \sqrt{\frac{|\mathcal{Q}| \log(N)}{N}}.$$

Here (i) is due to Cauchy-Schwarz, and (ii) is from Jensen's inequality, and (iii) is the standard rate of the least squared estimator [214].

For the second difference, we have

$$\operatorname{cost}(\hat{\mathcal{L}}, \pi^{\star}) - \operatorname{cost}(\mathcal{L}^{\star}, \pi^{\star}) = \sum_{q \in \mathcal{Q}} P(q) \left(1(q \notin \hat{\mathcal{L}}) - 1(q \notin \mathcal{L}^{\star}) \right) \min(c_{s}^{\star}(q), c_{l}^{\star}(q))$$
$$= \sum_{q \in \mathcal{Q}} P(q) \left(1(q \in \mathcal{L}^{\star}) - 1(q \in \hat{\mathcal{L}}) \right) \min(c_{s}^{\star}(q), c_{l}^{\star}(q)).$$

Note that for any $q \in \mathcal{L}^{\star}$, we know that

$$\hat{P}(q) \left(\min\left(\hat{c}_{s}(q), \hat{c}_{l}(q)\right) - (B_{2} - B_{1}) \sqrt{\frac{\log(8|\mathcal{Q}|/\delta)}{2\sum_{n=1}^{N} 1(q_{n} = q)}} \right) \\
\geq \max\left(P(q) - \sqrt{\frac{2\log(8/\delta)}{N}}, 0 \right) \cdot \left(\min\left(c_{s}^{\star}(q), c_{l}^{\star}(q)\right) - 2(B_{2} - B_{1}) \sqrt{\frac{\log(8|\mathcal{Q}|/\delta)}{2\sum_{n=1}^{N} 1(q_{n} = q)}} \right) \\
\geq P(q) \min\left(c_{s}^{\star}(q), c_{l}^{\star}(q)\right) - C(B_{2} - B_{1}) \cdot \sqrt{\frac{\log(8|\mathcal{Q}|/\delta)}{2\sum_{n=1}^{N} 1(q_{n} = q)}} \\
\geq P(q) \min\left(c_{s}^{\star}(q), c_{l}^{\star}(q)\right) - C(B_{2} - B_{1}) \cdot \sqrt{\frac{B_{2}|\mathcal{Q}|\log(8|\mathcal{Q}|/\delta)}{B_{1}N}}.$$

The last inequality uses event E_3 and Lemma 45. And similarly, for any $q \notin \mathcal{L}^*$, we know that

$$\hat{P}(q) \left(\min(\hat{c}_{s}(q), \hat{c}_{l}(q)) - (B_{2} - B_{1}) \sqrt{\frac{\log(8|\mathcal{Q}|/\delta)}{2\sum_{n=1}^{N} 1(q_{n} = q)}} \right) \\
\leq \left(P(q) + \sqrt{\frac{2\log(8/\delta)}{N}} \right) \min(c_{s}^{*}(q), c_{l}^{*}(q)) \\
\leq P(q) \min(c_{s}^{*}(q), c_{l}^{*}(q)) + B_{2} \sqrt{\frac{2\log(8/\delta)}{N}}.$$

Now consider any $q \in \mathcal{L}_t$ but $q \notin \mathcal{L}^*$, and any other $q' \in \mathcal{L}^*$ but $q' \notin \mathcal{L}_t$. We have

$$\begin{split} P(q')\min\left(c_{s}^{\star}(q'),c_{l}^{\star}(q')\right) &- P(q)\min\left(c_{s}^{\star}(q),c_{l}^{\star}(q)\right)\\ \leq &\hat{P}(q')\left(\min\left(\hat{c}_{s}(q'),\hat{c}_{l}(q')\right) - (B_{2} - B_{1})\sqrt{\frac{\log(8|\mathcal{Q}|/\delta)}{2\sum_{n=1}^{N}1(q_{n} = q')}}\right)\\ &- &\hat{P}(q)\left(\min\left(\hat{c}_{s}(q),\hat{c}_{l}(q)\right) - (B_{2} - B_{1})\sqrt{\frac{\log(8|\mathcal{Q}|/\delta)}{2\sum_{n=1}^{N}1(q_{n} = q)}}\right)\\ &+ C(B_{2} - B_{1})\cdot\sqrt{\frac{B_{2}|\mathcal{Q}|\log(8|\mathcal{Q}|/\delta)}{B_{1}N}}\\ \leq &C(B_{2} - B_{1})\cdot\sqrt{\frac{B_{2}|\mathcal{Q}|\log(8|\mathcal{Q}|/\delta)}{B_{1}N}}. \end{split}$$

Here the last inequality uses the fact that q is inside \mathcal{L}_t and thus the difference between the first two terms are upper bounded by 0. Finally, we know that conditioned on $E_1 \cap E_2 \cap E_3 \cap E_4$,

we have

$$\operatorname{cost}(\hat{\mathcal{L}}, \pi^{\star}) - \operatorname{cost}(\mathcal{L}^{\star}, \pi^{\star}) \leq CL(B_2 - B_1) \cdot \sqrt{\frac{B_2|\mathcal{Q}|\log(8|\mathcal{Q}|/\delta)}{B_1N}}.$$

Overall, we know that

$$\mathbb{E}[\operatorname{cost}(\hat{\mathcal{L}}, \hat{\pi}) - \operatorname{cost}(\hat{\mathcal{L}}, \pi^{\star})] \le B_2 \delta + CL(B_2 - B_1) \cdot \sqrt{\frac{B_2|\mathcal{Q}|\log(8|\mathcal{Q}|/\delta)}{B_1 N}}$$

Taking $\delta = 1/N$ finishes the proof.

D.8 Proof of Theorem 34

Proof. Let $g_k(q)$ be the length of the interval between the k-th and (k+1)-th arrival of query q in the sequence of received queries (we set $g_k(q) = 0$ if k exceeds the total number of times q is queried.). Define the following four events:

$$E_{1,t}: \left\{ \forall q \in \mathcal{Q}, |\hat{P}_{t-1}(q) - P(q)| \le \min\left(1, \sqrt{\frac{2\log(8T/\delta)}{t-1}}\right) \right\}$$

$$E_{2,t}: \left\{ \forall q \in \mathcal{Q}, \hat{c}_{l,t-1}(q) \in \left[c_l^*(q) - 2(B_2 - B_1)\min\left(1, \sqrt{\frac{\log(8T|\mathcal{Q}|/\delta)}{2\sum_{i=1}^{t-1}1(s_i = 0, q_i = q)}}\right), c_l^*(q)\right] \right\}$$

$$E_{3,t}: \left\{ \forall q \in \mathcal{Q}, \hat{c}_{s,t-1}(q) \in \left[c_s^*(q) - 2(B_2 - B_1)\min\left(1, \sqrt{\frac{\log(8T|\mathcal{Q}|/\delta)}{2\sum_{i=1}^{t-1}1(s_i = 1, q_i = q)}}\right), c_s^*(q)\right] \right\}$$

$$E_4: \left\{ \forall q \in \mathcal{L}^*, k \le T, g_k(q) \le \frac{B_2|\mathcal{Q}|\log(4TL/\delta)}{B_1} \right\}$$

From the same analysis as Lemma 47, we know that the four events $(\bigcap_{s=1}^{t} E_{1,s} \cap E_{2,s} \cap E_{3,s}) \cap E_4$ hold simultaneously with probability at least $1 - \delta$.

Let
$$E^t = \bigcap_{s=1}^t E_{1,s} \cap E_{2,s} \cap E_{3,s}$$
. The regret can be decomposed as follows.

$$\begin{aligned} \operatorname{Regret}(T) \\ &= \sum_{t=1}^{T} \mathbb{E}[\operatorname{cost}(q_t, \mathcal{L}_t, \pi_t) - \operatorname{cost}(q_t, \mathcal{L}^{\star}, \pi^{\star})] \\ &\leq \sum_{t=1}^{T} \mathbb{E}[\operatorname{cost}(q_t, \mathcal{L}_t, \pi_t) - \operatorname{cost}(q_t, \mathcal{L}^{\star}, \pi^{\star})1(E^t)] + \mathbb{E}[\operatorname{cost}(q_t, \mathcal{L}_t, \pi_t) - \operatorname{cost}(q_t, \mathcal{L}^{\star}, \pi^{\star})1(\bar{E}^t)] \\ &\leq \sum_{t=1}^{T} \mathbb{E}[\operatorname{cost}(q_t, \mathcal{L}_t, \pi_t) - \operatorname{cost}(q_t, \mathcal{L}^{\star}, \pi^{\star})1(E^t)] + \delta T B_2 \\ &= \sum_{t=1}^{T} \mathbb{E}[(\operatorname{cost}(q_t, \mathcal{L}_t, \pi_t) - \operatorname{cost}(q_t, \mathcal{L}_t, \pi^{\star}) + \operatorname{cost}(q_t, \mathcal{L}_t, \pi^{\star}) - \operatorname{cost}(q_t, \mathcal{L}^{\star}, \pi^{\star}))1(E^t)] + \delta T B_2. \end{aligned}$$

The first difference can be further written as

$$\begin{split} & \mathbb{E}[\operatorname{cost}(q_t, \mathcal{L}_t, \pi_t) - \operatorname{cost}(q_t, \mathcal{L}_t, \pi^*) \mid E^t] \\ = & \mathbb{E}[1(q_t \notin \mathcal{L}_t)(c_s^*(q_t)\pi_t(q_t) + c_l^*(q_t)(1 - \pi_t(q_t)) - c_s^*(q_t)\pi^*(q_t) - c_l^*(q_t)(1 - \pi^*(q_t))) \mid E^t] \\ = & \mathbb{E}[1(q_t \notin \mathcal{L}_t)(c_s^*(q_t)\pi_t(q_t) + c_l^*(q_t)(1 - \pi_t(q_t)) - \min(c_s^*(q_t), c_l^*(q_t))) \mid E^t] \\ \leq & \mathbb{E}[c_s^*(q_t)\pi_t(q_t) + c_l^*(q_t)(1 - \pi_t(q_t)) - \min(c_s^*(q_t), c_l^*(q_t)) \mid E^t] \\ = & \mathbb{E}[c_s^*(q_t)1(\hat{c}_{s,t}(q_t) \leq \hat{c}_{l,t}(q_t)) + c_l^*(q_t)1(\hat{c}_{s,t}(q_t) \geq \hat{c}_{l,t}(q_t)) - \min(c_s^*(q_t), c_l^*(q_t)) \mid E^t]. \end{split}$$

If $\hat{c}_{s,t}(q_t) - \hat{c}_{l,t}(q_t)$ has the same sign as $c_s^{\star}(q_t) - c_l^{\star}(q_t)$, the difference $c_s^{\star}(q_t) 1(\hat{c}_{s,t}(q_t) \leq \hat{c}_{l,t}(q_t)) + c_l^{\star}(q_t) 1(\hat{c}_{s,t}(q_t) > \hat{c}_{l,t}(q_t)) - \min(c_s^{\star}(q_t), c_l^{\star}(q_t))$ becomes 0. Otherwise, if $c_s^{\star}(q_t) - c_l^{\star}(q_t) > 0$ and $\hat{c}_{s,t}(q_t) - \hat{c}_{l,t}(q_t) \leq 0$, we know that $s_t = 1$ and

$$c_{s}^{\star}(q) - c_{l}^{\star}(q) \leq \hat{c}_{s,t}(q) - \hat{c}_{l,t}(q) + 2(B_{2} - B_{1}) \min\left(1, \sqrt{\frac{\log(8T|\mathcal{Q}|/\delta)}{2\sum_{i=1}^{t-1} 1(s_{i} = 1, q_{i} = q)}}\right)$$
$$\leq 2(B_{2} - B_{1}) \min\left(1, \sqrt{\frac{\log(8T|\mathcal{Q}|/\delta)}{2\sum_{i=1}^{t-1} 1(s_{i} = 1, q_{i} = q)}}\right).$$

And similarly if $c_s^{\star}(q) - c_l^{\star}(q) \leq 0$ and $\hat{c}_{s,t}(q_t) - \hat{c}_{l,t}(q_t) > 0$, we know that $s_t = 0$ and

$$\begin{aligned} c_{l}^{\star}(q) - c_{s}^{\star}(q) &\leq 2(B_{2} - B_{1}) \min\left(1, \sqrt{\frac{\log(8T|\mathcal{Q}|/\delta)}{2\sum_{i=1}^{t-1}1(s_{i}=0,q_{i}=q)}}\right). \text{ Overall, we have} \\ &\sum_{t=1}^{T} \mathbb{E}[(\cot(q_{t}, \mathcal{L}_{t}, \pi_{t}) - \cot(q_{t}, \mathcal{L}_{t}, \pi^{\star})] \\ &\leq 2(B_{2} - B_{1}) \sum_{t=1}^{T} \sum_{q \in \mathcal{Q}} \mathbb{E}\left[1(s_{t} = 1, q_{t} = q) \min\left(1, \sqrt{\frac{\log(8T|\mathcal{Q}|/\delta)}{2\sum_{i=1}^{t-1}1(s_{i}=1, q_{i}=q)}}\right) \\ &+ 1(s_{t} = 0, q_{t} = q) \min\left(1, \sqrt{\frac{\log(8T|\mathcal{Q}|/\delta)}{2\sum_{i=1}^{t-1}1(s_{i}=0, q_{i}=q)}}\right)\right] \\ &\leq 2(B_{2} - B_{1})\sqrt{|\mathcal{Q}|T\log(8|\mathcal{Q}|T/\delta)}. \end{aligned}$$

Here the last inequality uses the fact that for each $q \in \mathcal{Q}$, the summation over time step is upper bounded by $2\sum_{i=1}^{T(q)} \sqrt{\log(8T|\mathcal{Q}|/\delta)/2i} \leq 2\sqrt{\log(8T|\mathcal{Q}|/\delta)T(q)}$, where T(q) is the number of steps of receiving query q in total T steps. Optimizing over T(q) gives the final bound.

For the second difference, we have

$$\mathbb{E}[\operatorname{cost}(\mathcal{L}_{t}, \pi^{\star}) - \operatorname{cost}(\mathcal{L}^{\star}, \pi^{\star}) \mid E^{t}] = \mathbb{E}\left[\sum_{q \in \mathcal{Q}} P(q) \left(1(q \notin \mathcal{L}_{t}) - 1(q \notin \mathcal{L}^{\star})\right) \min(c_{s}^{\star}(q), c_{l}^{\star}(q)) \mid E^{t}\right]$$
$$= \mathbb{E}\left[\sum_{q \in \mathcal{Q}} P(q) \left(1(q \in \mathcal{L}^{\star}) - 1(q \in \mathcal{L}_{t})\right) \min(c_{s}^{\star}(q), c_{l}^{\star}(q)) \mid E^{t}\right]$$

Note that for any $q \in \mathcal{L}^{\star}$, we know that

$$\begin{split} \hat{P}_{t}(q) \min\left(\hat{c}_{s,t}(q), \hat{c}_{l,t}(q)\right) \\ &\geq \max\left(P(q) - \sqrt{\frac{2\log(8/\delta)}{t}}, 0\right) \cdot 1(\pi_{t}(q) = 1) \\ &\quad \cdot \left(c_{s}^{\star}(q) - 2(B_{2} - B_{1}) \min\left(1, \sqrt{\frac{\log(8|\mathcal{Q}|/\delta)}{2\sum_{i=1}^{t-1}1(s_{i} = 1, q_{i} = q)}\right)\right) \\ &\quad + 1(\pi_{t}(q) = 0) \left(c_{l}^{\star}(q) - 2(B_{2} - B_{1}) \min\left(1, \sqrt{\frac{\log(8|\mathcal{Q}|/\delta)}{2\sum_{i=1}^{t-1}1(s_{i} = 0, q_{i} = q)}\right)\right) \right) \\ &\geq P(q) \min\left(c_{s}^{\star}(q), c_{l}^{\star}(q)\right) - (B_{2} - B_{1}) \sqrt{\frac{2\log(8/\delta)}{t}} + P(q) \cdot \left(1(\pi_{t}(q) = 1)\right) \\ &\quad \cdot \left(c_{s}^{\star}(q) - 2(B_{2} - B_{1}) \min\left(1, \sqrt{\frac{\log(8|\mathcal{Q}|/\delta)}{2\sum_{i=1}^{t-1}1(s_{i} = 1, q_{i} = q)}\right)\right) \right) \\ &\quad + 1(\pi_{t}(q) = 0) \left(c_{l}^{\star}(q) - 2(B_{2} - B_{1}) \min\left(1, \sqrt{\frac{\log(8|\mathcal{Q}|/\delta)}{2\sum_{i=1}^{t-1}1(s_{i} = 0, q_{i} = q)}\right)\right) \right) \\ &\quad - \min(c_{s}^{\star}(q), c_{l}^{\star}(q)) - C(B_{2} - B_{1}) \min\left(1, \sqrt{\frac{\log(8|\mathcal{Q}|/\delta)}{2\sum_{i=1}^{t-1}1(s_{i} = 0, q_{i} = q)}\right)\right) \end{split}$$

Below we justify the last inequality. First, note that $\pi_t(q) = 1$ is equivalent to that $\hat{c}_{s,t}(q) \leq \hat{c}_{l,t}(q)$. Thus if $\hat{c}_{s,t}(q_t) - \hat{c}_{l,t}(q_t)$ has the same sign as $c_s^*(q_t) - c_l^*(q_t)$, the above inequality holds. Now consider the case when $\hat{c}_{s,t}(q_t) - \hat{c}_{l,t}(q_t)$ has a different sign as $c_s^*(q_t) - c_l^*(q_t)$. Assume that $\hat{c}_{s,t}(q_t) > \hat{c}_{l,t}(q_t)$ and $c_s^*(q_t) < c_l^*(q_t)$. We know that $\pi_t(q) = 0$, and

$$c_{l}^{\star}(q) - 2(B_{2} - B_{1}) \min\left(1, \sqrt{\frac{\log(8|\mathcal{Q}|/\delta)}{2\sum_{i=1}^{t-1} 1(s_{i} = 0, q_{i} = q)}}\right) - c_{s}^{\star}(q)$$

> $-2(B_{2} - B_{1}) \min\left(1, \sqrt{\frac{\log(8|\mathcal{Q}|/\delta)}{2\sum_{i=1}^{t-1} 1(s_{i} = 0, q_{i} = q)}}\right).$

Similarly we can prove that for the reversed case. Now for any $q \notin \mathcal{L}^{\star}$, we know that

$$\hat{P}_{t}(q)\min\left(\hat{c}_{s,t}(q),\hat{c}_{l,t}(q)\right) \leq \left(P(q) + \sqrt{\frac{2\log(8/\delta)}{t}}\right)\min\left(c_{s}^{\star}(q),c_{l}^{\star}(q)\right) \\
\leq P(q)\min\left(c_{s}^{\star}(q),c_{l}^{\star}(q)\right) + B_{2}\sqrt{\frac{2\log(8/\delta)}{t}}.$$

Now consider any $q \in \mathcal{L}_t$ but $q \notin \mathcal{L}^*$, and any other $q' \in \mathcal{L}^*$ but $q' \notin \mathcal{L}_t$. We have

$$P(q')\min(c_s^*(q'), c_l^*(q')) - P(q)\min(c_s^*(q), c_l^*(q)) \\ \leq \hat{P}_t(q')\min(\hat{c}_{s,t}(q'), \hat{c}_{l,t}(q')) - \hat{P}_t(q)\min(\hat{c}_{s,t}(q), \hat{c}_{l,t}(q)) \\ + C(B_2 - B_1) \cdot \min\left(1, \sqrt{\frac{\log(8|\mathcal{Q}|/\delta)}{2\sum_{i=1}^{t-1}1(s_i = \pi_t(q), q_i = q)}}\right) + B_2\sqrt{\frac{2\log(8/\delta)}{t}} \\ \leq C(B_2 - B_1) \cdot \min\left(1, \sqrt{\frac{\log(8|\mathcal{Q}|/\delta)}{2\sum_{i=1}^{t-1}1(s_i = \pi_t(q), q_i = q)}}\right) + B_2\sqrt{\frac{2\log(8/\delta)}{t}}.$$

Here the last inequality uses the fact that q is inside \mathcal{L}_t . Thus we have

$$\begin{split} &\sum_{t=1}^{T} \mathbb{E}[\operatorname{cost}(\mathcal{L}_{t}, \pi^{\star}) - \operatorname{cost}(\mathcal{L}^{\star}, \pi^{\star})1(E^{t})] \\ &\leq \sum_{t=1}^{T} \mathbb{E}[\operatorname{cost}(\mathcal{L}_{t}, \pi^{\star}) - \operatorname{cost}(\mathcal{L}^{\star}, \pi^{\star})1(E^{t} \cap E_{4})] + B_{2}T\delta \\ &\leq B_{2}T\delta + C\sum_{t=1}^{T} \mathbb{E}\left[\sum_{q \in \mathcal{L}^{\star}} 1(q \notin \mathcal{L}_{t})(B_{2} - B_{1}) \cdot \min\left(1, \sqrt{\frac{\log(8|\mathcal{Q}|/\delta)}{2\sum_{i=1}^{t-1}1(s_{i} = \pi_{t}(q), q_{i} = q)}\right) \right. \\ &+ B_{2}\sqrt{\frac{2\log(8/\delta)}{t}} \mid E^{t} \cap E_{4}\right] \\ &\leq C \cdot \left(B_{2}T\delta + LB_{2}\sqrt{2T\log(8/\delta)} + (B_{2} - B_{1})\log(8T|\mathcal{Q}|/\delta) \right. \\ &\quad \cdot \sum_{q \in \mathcal{L}^{\star}} \sum_{t=1}^{T} \mathbb{E}\left[1(q \notin \mathcal{L}_{t}) \cdot \min\left(1, \sqrt{\frac{1}{\sum_{i=1}^{t-1}1(s_{i} = \pi_{t}(q), q_{i} = q)}\right) \mid E^{t} \cap E_{4}\right]\right) \\ &= C \cdot \left(B_{2}T\delta + LB_{2}\sqrt{2T\log(8/\delta)} + (B_{2} - B_{1})\log(8T|\mathcal{Q}|/\delta) \right. \\ &\quad \cdot \sum_{q \in \mathcal{L}^{\star}} \sum_{\pi \in \{1,2\}} \sum_{t=1}^{T} \mathbb{E}\left[1(q \notin \mathcal{L}_{t}, \pi_{t}(q) = \pi)\min\left(1, \sqrt{\frac{1}{\sum_{i=1}^{t-1}1(s_{i} = \pi, q_{i} = q)}\right) \mid E^{t} \cap E_{4}\right]\right) \end{split}$$

Let $T_t(q, \pi) = \sum_{i=1}^{t-1} 1(s_i = \pi, q_i = q)$. Now for each $q \in \mathcal{L}^*$ and $\pi \in \{0, 1\}$, we look at the term $\sum_{t=1}^{T} \mathbb{E} \left[1(q \notin \mathcal{L}_t, \pi_t = \pi) \cdot \min\left(1, \sqrt{\frac{1}{T_t(q,\pi)}}\right) \mid E^t \cap E_4 \right]$. We prove the following lemma:

Lemma 49. We have

$$\sum_{t=1}^{T} \mathbb{E}\left[1(q \notin \mathcal{L}_t, \pi_t(q) = \pi) \min\left(1, \sqrt{\frac{1}{T_t(q, \pi)}}\right) \mid E^t \cap E_4\right] \le \frac{CB_2|\mathcal{Q}|\log(3TL/\delta)\sqrt{T}}{B_1} + T\delta.$$

Proof. Let $t_k(q) = \sum_{l=1}^{k-1} g_l(q)$ be the step that the k-th query of q arrives. And let $E = (\bigcap_{t=1}^{T} E_t) \cap E_4$. The summation can be written as

$$\sum_{t=1}^{T} \mathbb{E} \left[1(q \notin \mathcal{L}_t, \pi_t(q) = \pi) \cdot \min\left(1, \sqrt{\frac{1}{T_t(q, \pi)}}\right) \mid E^t \cap E_3 \right]$$

$$\leq \sum_{t=1}^{T} \mathbb{E} \left[1(q \notin \mathcal{L}_t, \pi_t(q) = \pi) \cdot \min\left(1, \sqrt{\frac{1}{T_t(q, \pi)}}\right) \mid E \right] + T\delta$$

$$= \sum_{k=0}^{T} \mathbb{E} \left[\sum_{t=t_k(q)+1}^{t_{k+1}(q)} 1(q \notin \mathcal{L}_t, \pi_t(q) = \pi) \cdot \min\left(1, \sqrt{\frac{1}{T_t(q, \pi)}}\right) \mid E \right] + T\delta$$

$$\leq \sum_{k=0}^{T} \mathbb{E} \left[\sum_{t=t_k(q)+1}^{t_{k+1}(q)} 1(q \notin \mathcal{L}_{t_{k+1}(q)}, \pi_{t_{k+1}}(q) = \pi) \cdot \min\left(1, \sqrt{\frac{1}{T_{t_k(q)+1}(q, \pi)}}\right) \mid E \right] + T\delta.$$

The last inequality is due to (a). $T_t(q, \pi)$ does not change if at round t the query is not q; (b). if $q \in \mathcal{L}_{t_{k+1}(q)}$, we will have $q \in \mathcal{L}_t$ for any $t \in [t_k(q) + 1, t_{k+1}(q)]$ since q never arrives in the middle and must remain in the cache set until $t_{k+1}(q)$; (c) For $t \in [t_k(q) + 1, t_{k+1}(q)]$, π_t does not change since both the frequency and cost estimator does not change for q. From the definition of $t_k(q)$, we know that $T_{t_k(q)+1}(q, \pi) \geq 1$ and thus we can drop the the minimum in the above equation. Now from event E_4 , we know that

$$\sum_{k=0}^{T} \mathbb{E} \left[\sum_{t=t_{k}(q)+1}^{t_{k+1}(q)} 1(q \notin \mathcal{L}_{t_{k+1}(q)}, \pi_{t_{k+1}}(q) = \pi) \cdot \min\left(1, \sqrt{\frac{1}{T_{t_{k}(q)+1}(q, \pi)}}\right) \mid E \right]$$

$$\leq \sum_{k=0}^{T} \mathbb{E} \left[g_{k}(q) \cdot 1(q \notin \mathcal{L}_{t_{k+1}(q)}, \pi_{t_{k+1}}(q) = \pi) \cdot \min\left(1, \sqrt{\frac{1}{T_{t_{k}(q)+1}(q, \pi)}}\right) \mid E \right]$$

$$\leq \frac{B_{2}|\mathcal{Q}|\log(4TL/\delta)}{B_{1}} \cdot \sum_{k=0}^{T} \mathbb{E} \left[1(q \notin \mathcal{L}_{t_{k+1}(q)}, \pi_{t_{k+1}}(q) = \pi) \cdot \min\left(1, \sqrt{\frac{1}{T_{t_{k}(q)+1}(q, \pi)}}\right) \mid E \right]$$

We know that $T_{t_{k+1}(q)+1}(q,\pi) = T_{t_{k+1}(q)}(q,\pi) + 1 = T_{t_k(q)+1}(q,\pi) + 1$ if $q \notin \mathcal{L}_{t_{k+1}(q)}$ and $\pi_{t_{k+1}}(q) = \pi$ since the query q missing the cache will be sent to one of the models, and only the one selected will observe the cost. Thus overall, we know that we have either $T_{t_{k+1}(q)+1}(q,\pi) = T_{t_k(q)+1}(q,\pi) + 1$, or $1(q \notin \mathcal{L}_{t_{k+1}(q)}, \pi_{t_{k+1}}(q) = \pi) \cdot \min\left(1, \sqrt{\frac{1}{T_{t_k(q)+1}(q,\pi)}}\right) = 0$ and $T_{t_{k+1}(q)+1}(q,\pi) = T_{t_k(q)+1}(q,\pi)$. Thus overall, we have

$$\sum_{k=0}^{T} \mathbb{E}\left[1(q \notin \mathcal{L}_{t_{k+1}(q)}, \pi_{t_{k+1}}(q) = \pi) \cdot \min\left(1, \sqrt{\frac{1}{T_{t_k(q)+1}(q, \pi)}}\right) \mid E\right] \le \sum_{k=1}^{T} \frac{1}{\sqrt{k}} \le C\sqrt{T}.$$

Thus we know that the second difference satisfies

$$\sum_{t=1}^{T} \mathbb{E}[\operatorname{cost}(\mathcal{L}_t, \pi^*) - \operatorname{cost}(\mathcal{L}^*, \pi^*) 1(E^t)] \leq C \cdot \left(B_2 T \delta + L B_2 \sqrt{2T \log(8/\delta)} + \frac{L(B_2 - B_1) B_2 |\mathcal{Q}| L \log^2(T|\mathcal{Q}|/\delta)}{B_1} \cdot \sqrt{T} \right).$$

Overall by taking $\delta = 1/T$, we know that

$$\operatorname{Regret}(T) \leq \frac{CL(B_2 - B_1)B_2|\mathcal{Q}|L\log^2(T|\mathcal{Q}|)}{B_1} \cdot \sqrt{T}.$$

D.9 Additional Experiments

Synthetic Datasets

We conduct synthetic online and offline experiments for joint optimization of caching and model multiplexing. We use i.i.d. Bernoulli distributions for two models because we want to mimic the model ensemble use case and give a large penalty to the wrong output. In Figure 5.2, we plot the cumulative cost and regret in online learning for LFU and LEC caching algorithms. We present more data points in Table D.1 and Table D.2, under the same setting of 10000 requests with 20 different queries and cache size 10. Similar to the real dataset setting, we compare all combinations of caching strategy choices and model multiplexer choices. We consider the frequency distribution as power distribution with $\alpha = 0.5$ and 0.8. The ground truth cost for each query processed by both models is set as a sample from $r \cdot X + 1$, where r is called as cost ratio and X is a random variable generated from a Bernoulli distribution with the parameter 0.5. We consider the model multiplexer accuracy with 0.8 and 1. We repeat the simulation 1000 times and take the mean. Consistent with Figure 5.2, our simulation suggests that LEC with a perfect model multiplexer significantly improves the baselines when the cost ratio is large. Simulation 1000 times cannot remove all randomness so we can observe some fluctuations. Theoretically, the columns of choosing model 1 and the columns of choosing model 2 should behave similarly.

Real Datasets

In this section, we provide additional experiments on real-world dataset. We run both offline and online algorithms on Lambada dataset and OpenAssistant dataset, with OPT-1.3B vs OPT-13B or FastChat-T5-3B vs Vicuna-13B. We mainly consider three settings:

• In Table 5-12, we consider distinct prompt size 100, total query size 10000, cache size 40, same as the experiments in the main text.

α	cost ratio	selector accu-	LFU+ model 1	LFU+ model 2	LFU+ selector	LEC+ model 1	LEC+ model 2	LEC+ selector
		racy						
$\begin{array}{c} 0.5\\ 0.8\end{array}$	$1.5 \\ 1.5$	$0.8 \\ 0.8$	$5.25 \\ 7.61$	$5.24 \\ 7.60$	$4.09 \\ 5.93$	$4.40 \\ 5.73$	$4.36 \\ 5.68$	$\begin{array}{c} 3.59 \\ 4.85 \end{array}$
$\begin{array}{c} 0.5\\ 0.8\end{array}$	$1.5 \\ 1.5$	1 1	$5.25 \\ 7.61$	$5.24 \\ 7.60$	$3.31 \\ 4.81$	$4.40 \\ 5.73$	$4.36 \\ 5.68$	$\begin{array}{c} 2.74\\ 3.68\end{array}$
$\begin{array}{c} 0.5 \\ 0.8 \end{array}$	$\begin{array}{c} 100 \\ 100 \end{array}$	$\begin{array}{c} 0.8\\ 0.8\end{array}$	$\begin{array}{c} 148.05\\ 214.93\end{array}$	$147.38 \\ 213.88$	$103.43 \\ 150.31$	$29.93 \\ 43.77$	$\begin{array}{c} 26.83\\ 39.02 \end{array}$	$39.32 \\ 49.88$
$\begin{array}{c} 0.5 \\ 0.8 \end{array}$	100 100	1 1	$148.05 \\ 214.93$	147.38 213.88	73.94 107.63	$29.93 \\ 43.77$	26.83 39.02	$\begin{array}{c} 3.12\\ 4.19\end{array}$

Table D.1. Simulation results for the proposed caching algorithm for offline syntheticdataset

α	cost ratio	LFU+	LFU+	LFU+	LEC+	LEC+	LEC+
		model 1	model 2	selector	model 1	model 2	selector
0.5	1.5	5.35	5.34	4.34	4.60	4.59	3.75
0.8	1.5	7.79	7.78	6.32	6.01	5.98	5.09
0.5	100	150.93	150.37	76.80	31.88	28.65	4.85
0.8	100	220.19	219.49	112.26	46.44	41.45	6.31

 Table D.2. Simulation results for the proposed caching algorithm for online synthetic dataset

- In Table 13-20, we consider distinct prompt size 1000, total query size 2000, cache size 100.
- In Table 21-28, we consider distinct prompt size 1000, total query size 2000, cache size 0, same as the experiments in the main text.

Our proposed algorithm mostly gives dominant results over other baselines. In Table 21-28, the cache size is set to be 0 so there is no difference between LFU and LEC.

α	LFU+	LFU+	LFU+	LEC+	LEC+	LEC+
	large	cas-	selec-	large	cas-	selec-
		cade	tor		cade	tor
0.2	3.77	4.11	2.17	3.71	1.83	1.26
0.5	8.11	8.85	4.54	7.93	3.51	2.29
0.8	11.43	12.47	6.35	10.91	4.63	2.86

Table D.3: FLOPs for online lambda dataset, opt-1.3b vs opt-13b

α	LFU+ large	LFU+ cas- cade	LFU+ selec- tor	LEC+ large	LEC+ cas- cade	LEC+ selec- tor
0.2 0.5	$349.44 \\751.78$	$425.05 \\ 916.58$	$224.33 \\ 469.74$	$349.47 \\751.77$	$241.89 \\ 462.34$	$\begin{array}{c} 170.02\\ 321.46\end{array}$
0.8	1059.95	1290.26	656.61	1060.01	596.92	397.01

Table D.4: Latency for online lambda dataset, opt-1.3b vs opt-13b

α	LFU+ large	LFU+ cas- cade	LFU+ selec- tor	LEC+ large	LEC+ cas- cade	LEC+ selec- tor
$0.2 \\ 0.5$	$\begin{array}{c} 1.66\\ 3.56\end{array}$	$\begin{array}{c} 1.68\\ 3.61 \end{array}$	$\begin{array}{c} 0.90 \\ 1.88 \end{array}$	$1.12 \\ 2.20$	$0.67 \\ 1.22$	$\begin{array}{c} 0.48\\ 0.87\end{array}$
0.8	5.01	5.09	2.64	2.75	1.49	1.04

Table D.5: FLOPs for online oasst dataset, fastchat-t5 vs vicuna

α	LFU+	LFU+	LFU+	LEC+	LEC+	LEC+
	large	cas-	selec-	large	cas-	selec-
		cade	tor		cade	tor
0.2	9.31	13.88	7.24	8.74	8.82	5.93
0.5	20.04	29.88	15.11	18.68	16.90	11.87
0.8	28.24	42.12	21.14	26.07	20.31	15.49

Table D.6: Latency for online oasst dataset, fastchat-t5 vs vicuna

α	selector accu- racy	LFU+ large	LFU+ cas- cade	LFU+ selec- tor	LEC+ large	LEC+ cas- cade	LEC+ selec- tor
0.2 0.5 0.8	0.8 0.8 0.8	3.49 7.71 10.81	3.81 8.43 11.80	$2.60 \\ 5.76 \\ 8.06$	$3.44 \\ 7.54 \\ 10.36$	$1.50 \\ 3.09 \\ 4.11$	2.00 3.94 4.76
$0.2 \\ 0.5 \\ 0.8$	1 1 1	3.49 7.71 10.81	3.81 8.43 11.80	$ 1.91 \\ 4.22 \\ 5.90 $	$3.44 \\ 7.54 \\ 10.36$	$1.50 \\ 3.09 \\ 4.11$	$0.99 \\ 1.97 \\ 2.50$

Table D.7: FLOPs for offline lambda dataset, opt-1.3b vs opt-13b

α	selector accu- racy	LFU+ large	LFU+ cas- cade	LFU+ selec- tor	LEC+ large	LEC+ cas- cade	LEC+ selec- tor
$0.2 \\ 0.5 \\ 0.8$	$0.8 \\ 0.8 \\ 0.8$	324.02 714.87 1002.26	394.22 872.60 1220.93	261.89 578.91 810.45	324.02 714.87 1002.25	$207.71 \\ 417.01 \\ 538.56$	$214.03 \\ 442.05 \\ 549.06$
0.2 0.5 0.8	1 1 1	324.02 714.87 1002.26	394.22 872.60 1220.93	$ \begin{array}{r} 197.00 \\ 435.74 \\ 609.95 \\ \end{array} $	324.02 714.87 1002.25	$207.71 \\ 417.01 \\ 538.56$	$141.55\\287.72\\358.33$

Table D.8: Latency for offline lambda dataset, opt-1.3b vs opt-13b

α	selector	LFU+	LFU+	LFU+	LEC+	LEC+	LEC+
	accu-	large	cas-	selec-	large	cas-	selec-
	racy		cade	tor		cade	tor
0.2	0.8	1.54	1.55	1.09	1.01	0.56	0.66
0.5	0.8	3.39	3.42	2.41	2.08	1.10	1.31
0.8	0.8	4.74	4.82	3.38	2.61	1.36	1.62
0.2	1	1.54	1.55	0.79	1.01	0.56	0.38
0.5	1	3.39	3.42	1.75	2.08	1.10	0.76
0.8	1	4.74	4.82	2.45	2.61	1.36	0.93

Table D.9: FLOPs for offline oasst dataset, fastchat-t5 vs vicuna

α	selector accu- racy	LFU+ large	LFU+ cas- cade	LFU+ selec- tor	LEC+ large	LEC+ cas- cade	LEC+ selec- tor
$0.2 \\ 0.5 \\ 0.8$	$0.8 \\ 0.8 \\ 0.8$	8.64 19.07 26.70	12.86 28.42 39.91	8.06 17.81 24.98	8.03 17.66 24.53	7.76 15.63 19.01	$6.73 \\ 14.14 \\ 18.12$
$0.2 \\ 0.5 \\ 0.8$	1 1 1	$8.64 \\ 19.07 \\ 26.70$	12.86 28.42 39.91	6.28 13.86 19.44	$8.03 \\ 17.66 \\ 24.53$	7.76 15.63 19.01	$\begin{array}{c} 4.86 \\ 10.43 \\ 13.80 \end{array}$

Table D.10: Latency for offline oasst dataset, fastchat-t5 vs vicuna

α	LFU+ large	LFU+ cas- cade	LFU+ selec- tor	LEC+ large	LEC+ cas- cade	LEC+ selec- tor
$0.2 \\ 0.5 \\ 0.8$	$ 1.81 \\ 3.14 \\ 3.67 $	1.97 3.42 3.99	$1.62 \\ 2.59 \\ 2.95$	$ 1.80 \\ 3.12 \\ 3.64 $	$1.72 \\ 3.01 \\ 3.57$	$1.52 \\ 2.41 \\ 2.76$

Table D.11: FLOPs for online lambda dataset, opt-1.3b vs opt-13b

α	LFU+	LFU+	LFU+	LEC+	LEC+	LEC+
	large	cas-	selec-	large	cas-	selec-
		cade	tor		cade	tor
0.2	0.17	0.21	0.17	0.17	0.19	0.17
0.5	0.30	0.36	0.27	0.30	0.33	0.26
0.8	0.35	0.42	0.31	0.35	0.39	0.30

Table D.12: Latency for online lambda dataset, opt-1.3b vs opt-13b

α	LFU+ large	LFU+ cas-	LFU+ selec-	LEC+ large	LEC+ cas-	LEC+ selec-
	0	cade	tor	0	cade	tor
0.2	1.03	1.16	0.96	0.88	0.91	0.85
0.5	1.78	2.01	1.52	1.45	1.47	1.27
0.8	2.08	2.34	1.73	1.64	1.69	1.43

Table D.13: FLOPs for online oasst dataset, fastchat-t5 vs vicuna

α	LFU+	LFU+	LFU+	LEC+	LEC+	LEC+
	large	cas-	selec-	large	cas-	selec-
		cade	tor		cade	tor
0.2	4.60	7.12	5.86	4.53	6.54	5.73
0.5	7.98	12.33	9.33	7.86	11.27	9.08
0.8	9.31	14.38	10.64	9.19	13.12	10.35

Table D.14: Latency for online oasst dataset, fastchat-t5 vs vicuna

α	selector accu- racy	LFU+ large	LFU+ cas- cade	LFU+ selec- tor	LEC+ large	LEC+ cas- cade	LEC+ selec- tor
$0.2 \\ 0.5 \\ 0.8$	0.8 0.8 0.8	$ 1.40 \\ 2.62 \\ 3.09 $	1.52 2.85 3.37	$1.04 \\ 1.95 \\ 2.30$	$ 1.38 \\ 2.59 \\ 3.05 $	$1.06 \\ 2.15 \\ 2.62$	$0.91 \\ 1.69 \\ 1.98$
$0.2 \\ 0.5 \\ 0.8$	1 1 1	$ 1.40 \\ 2.62 \\ 3.09 $	$1.52 \\ 2.85 \\ 3.37$	$0.76 \\ 1.43 \\ 1.68$	$1.38 \\ 2.59 \\ 3.05$	$1.06 \\ 2.15 \\ 2.62$	$0.57 \\ 1.13 \\ 1.35$

Table D.15: FLOPs for offline lambda dataset, opt-1.3b vs opt-13b

α	selector accu- racy	LFU+ large	LFU+ cas- cade	LFU+ selec- tor	LEC+ large	LEC+ cas- cade	LEC+ selec- tor
$0.2 \\ 0.5 \\ 0.8$	$0.8 \\ 0.8 \\ 0.8$	$0.13 \\ 0.25 \\ 0.29$	$0.16 \\ 0.30 \\ 0.36$	$0.11 \\ 0.20 \\ 0.24$	$0.13 \\ 0.25 \\ 0.29$	$0.12 \\ 0.24 \\ 0.29$	$0.10 \\ 0.18 \\ 0.21$
$0.2 \\ 0.5 \\ 0.8$	1 1 1	$\begin{array}{c} 0.13 \\ 0.25 \\ 0.29 \end{array}$	$0.16 \\ 0.30 \\ 0.36$	$0.08 \\ 0.15 \\ 0.18$	$\begin{array}{c} 0.13 \\ 0.25 \\ 0.29 \end{array}$	$\begin{array}{c} 0.12 \\ 0.24 \\ 0.29 \end{array}$	$0.07 \\ 0.13 \\ 0.15$

Table D.16: Latency for offline lambda dataset, opt-1.3b vs opt-13b

α	selector	LFU+	LFU+	LFU+	LEC+	LEC+	LEC+
	racy	large	cade	tor	large	cade	tor
$0.2 \\ 0.5 \\ 0.8$	0.8 0.8 0.8	$0.79 \\ 1.48 \\ 1.75$	$0.90 \\ 1.67 \\ 1.97$	$0.61 \\ 1.13 \\ 1.34$	$0.51 \\ 0.97 \\ 1.13$	$0.41 \\ 0.84 \\ 0.99$	$0.35 \\ 0.68 \\ 0.79$
$0.2 \\ 0.5 \\ 0.8$	1 1 1	$0.79 \\ 1.48 \\ 1.75$	$0.90 \\ 1.67 \\ 1.97$	$0.45 \\ 0.84 \\ 0.99$	$0.51 \\ 0.97 \\ 1.13$	$0.41 \\ 0.84 \\ 0.99$	$0.23 \\ 0.46 \\ 0.54$

Table D.17: FLOPs for offline oasst dataset, fastchat-t5 vs vicuna

α	selector accu- racy	LFU+ large	LFU+ cas- cade	LFU+ selec- tor	LEC+ large	LEC+ cas- cade	LEC+ selec- tor
$0.2 \\ 0.5 \\ 0.8$	0.8 0.8 0.8	$3.55 \\ 6.65 \\ 7.86$	5.49 10.28 12.14	$3.42 \\ 6.41 \\ 7.57$	$3.41 \\ 6.45 \\ 7.61$	$4.43 \\ 8.50 \\ 10.06$	$3.15 \\ 5.90 \\ 6.95$
$0.2 \\ 0.5 \\ 0.8$	1 1 1	$3.55 \\ 6.65 \\ 7.86$	5.49 10.28 12.14	$2.69 \\ 5.04 \\ 5.95$	$3.41 \\ 6.45 \\ 7.61$	$ \begin{array}{r} 4.43 \\ 8.50 \\ 10.06 \end{array} $	$2.40 \\ 4.59 \\ 5.44$

Table D.18: Latency for offline oasst dataset, fastchat-t5 vs vicuna

α	LFU+	LFU+	LFU+	LEC+	LEC+	LEC+
	large	cas-	selec-	large	cas-	selec-
		cade	tor		cade	tor
0.2	4.13	4.49	2.88	4.13	4.49	2.88
0.5	4.13	4.49	3.12	4.13	4.49	3.12
0.8	4.13	4.49	3.21	4.13	4.49	3.21

Table D.19: FLOPs for online lambda dataset, opt-1.3b vs opt-13b

α	LFU+	LFU+	LFU+	LEC+	LEC+	LEC+
	large	cas-	selec-	large	cas-	selec-
		cade	tor		cade	tor
0.2	0.39	0.48	0.31	0.39	0.48	0.31
0.5	0.39	0.48	0.33	0.39	0.48	0.33
0.8	0.39	0.48	0.34	0.39	0.48	0.34

Table D.20: Latency for online lambda dataset, opt-1.3b vs opt-13b

α	LFU+ large	LFU+ cas-	LFU+ selec-	LEC+ large	LEC+ cas-	LEC+ selec-
		cade	tor		cade	tor
0.2	2.50	2.66	1.72	2.50	2.66	1.72
0.5	2.36	2.63	1.84	2.36	2.63	1.84
0.8	2.34	2.64	1.88	2.34	2.64	1.88

Table D.21: FLOPs for online oasst dataset, fastchat-t5 vs vicuna

α	LFU+ large	LFU+ cas- cade	LFU+ selec- tor	LEC+ large	LEC+ cas- cade	LEC+ selec- tor
$0.2 \\ 0.5 \\ 0.8$	10.45	16.08	10.37	10.45	16.08	10.37
	10.48	16.18	11.25	10.48	16.18	11.25
	10.48	16.19	11.54	10.48	16.19	11.54

Table D.22: Latency for online oasst dataset, fastchat-t5 vs vicuna

α	selector accu- racy	LFU+ large	LFU+ cas- cade	LFU+ selec- tor	LEC+ large	LEC+ cas- cade	LEC+ selec- tor
$0.2 \\ 0.5 \\ 0.8$	$0.8 \\ 0.8 \\ 0.8$	$ 4.13 \\ 4.13 \\ 4.13 $	$ \begin{array}{r} 4.49 \\ 4.49 \\ 4.49 \end{array} $	$3.07 \\ 3.07 \\ 3.07 \\ 3.07$	$ \begin{array}{r} 4.13 \\ 4.13 \\ 4.13 \end{array} $	$ \begin{array}{r} 4.49 \\ 4.49 \\ 4.49 \end{array} $	$3.07 \\ 3.07 \\ 3.07 \\ 3.07$
$0.2 \\ 0.5 \\ 0.8$	1 1 1	$ \begin{array}{r} 4.13 \\ 4.13 \\ 4.13 \end{array} $	$ \begin{array}{r} 4.49 \\ 4.49 \\ 4.49 \end{array} $	$2.24 \\ 2.25 \\ 2.25 \\ 2.25$	$ \begin{array}{r} 4.13 \\ 4.13 \\ 4.13 \end{array} $	$\begin{array}{c} 4.49 \\ 4.49 \\ 4.49 \end{array}$	$2.24 \\ 2.25 \\ 2.25$

Table D.23: FLOPs for offline lambda dataset, opt-1.3b vs opt-13b

α	selector accu-	LFU+ large	LFU+ cas-	LFU+ selec-	LEC+ large	LEC+ cas-	LEC+ selec-
	racy		cade	tor		cade	tor
0.2	0.8	0.39	0.48	0.32	0.39	0.48	0.32
0.5	0.8	0.39	0.48	0.32	0.39	0.48	0.32
0.8	0.8	0.39	0.48	0.32	0.39	0.48	0.32
0.2	1	0.39	0.48	0.24	0.39	0.48	0.24
0.5	1	0.39	0.48	0.24	0.39	0.48	0.24
0.8	1	0.39	0.48	0.24	0.39	0.48	0.24

Table D.24: Latency for offline lambda dataset, opt-1.3b vs opt-13b

α	selector accu- racy	LFU+ large	LFU+ cas- cade	LFU+ selec- tor	LEC+ large	LEC+ cas- cade	LEC+ selec- tor
$0.2 \\ 0.5 \\ 0.8$	0.8 0.8 0.8	2.50 2.36 2.34	2.66 2.63 2.64	$1.84 \\ 1.79 \\ 1.79$	2.50 2.36 2.34	2.66 2.63 2.64	$1.84 \\ 1.79 \\ 1.79$
$0.2 \\ 0.5 \\ 0.8$	1 1 1	2.50 2.36 2.34	2.66 2.63 2.64	$1.35 \\ 1.32 \\ 1.32$	2.50 2.36 2.34	2.66 2.63 2.64	$1.35 \\ 1.32 \\ 1.32$

 Table D.25:
 FLOPs for offline oasst dataset, fastchat-t5 vs vicuna

α	selector accu- racy	LFU+ large	LFU+ cas- cade	LFU+ selec- tor	LEC+ large	LEC+ cas- cade	LEC+ selec- tor
$0.2 \\ 0.5 \\ 0.8$	$0.8 \\ 0.8 \\ 0.8$	10.45 10.48 10.48	$16.08 \\ 16.18 \\ 16.19$	$10.05 \\ 10.09 \\ 10.10$	10.45 10.48 10.48	$16.08 \\ 16.18 \\ 16.19$	$10.05 \\ 10.09 \\ 10.10$
$0.2 \\ 0.5 \\ 0.8$	1 1 1	$10.45 \\ 10.48 \\ 10.48$	$16.08 \\ 16.18 \\ 16.19$	$7.91 \\ 7.94 \\ 7.94$	$10.45 \\ 10.48 \\ 10.48$	$16.08 \\ 16.18 \\ 16.19$	$7.91 \\ 7.94 \\ 7.94$

Table D.26: Latency for offline oasst dataset, fastchat-t5 vs vicuna