An Early Exploration of Causality Analysis for Irregularly Sampled Time Series via the Frequency Domain

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Temgine: enabling scalable and generic stochastic process analysis for irregularly observed time series with a duality of representations

by Francois Belletti

Research Project

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Abstract

Linear causal analysis is central to a wide range of important application spanning finance, the physical sciences, and engineering. Much of the existing literature in linear causal analysis operates in the time domain. Unfortunately, the direct application of time domain linear causal analysis to many real-world time series presents three critical challenges: irregular temporal sampling, long range dependencies, and scale. Moreover, real-world data is often collected at irregular time intervals across vast arrays of decentralized sensors and with long range dependencies [1] which make naive time domain correlation estimators spurious [2]. In this paper we present a frequency domain based estimation framework which naturally handles irregularly sampled data and long range dependencies while enabled memory and communication efficient distributed processing of time series data. By operating in the frequency domain we eliminate the need to interpolate and help mitigate the effects of long range dependencies. We implement and evaluate our new workflow in the distributed setting using Apache Spark and demonstrate on both Monte Carlo simulations and high-frequency financial trading that we can accurately recover causal structure at scale.

1 Introduction

The analysis of time series is central to applications ranging from statistical finance [3,4] to climate studies [5] or cyberphysical systems such as the transportation network [6]. In many of these applications one is interested in estimating the mutual linear predictive properties of events from time series data corresponding to a collection of data streams each of which is a series of pairs (timestamp, observation).

Figure 1: Frequency domain causal analysis workflow for two irregularly sampled correlated and lagged Brownian motion increments and the derived cross-correlogram of increments which highlights their causal structure. Dotted lines represent 5th, 50th and 95th percentiles respectively. Frequency domain estimation is treated in Section 2 and LRD erasure in Section 3. Once the cross-correlogram has been estimated, practitioners read out the direction of linear causality in the asymmetry of the curve. The lag value for which cross-correlation reaches a maximum can be interpreted as the characteristic delay of causation of the two processes.

In most applications, observations occur at random, unevenly spaced and unaligned time stamps. In such a setting we therefore consider two underlying processes \((X_t)_{t \in \mathbb{R}}\) and \((Y_t)_{t \in \mathbb{R}}\) that are only observed at discrete and finite timestamps in the form of two collections of data points: \((x_{t_x})_{t_x \in I_x}, (y_{t_y})_{t_y \in I_y}\). We adapt our definition of causality to this different theoretical framework. Let \(\phi\) be a causal convolution kernel with delay \(\tau\) (i.e. \(\phi(t) = 0\) whenever \(t < \tau\)), \((X)\) and \((Y)\) two continuous time stochastic processes, for instance, two Wiener processes. A popular instance of a causation kernel is for instance the exponential kernel: \(\phi(t), \theta = (\alpha, \beta) = \alpha \exp(-\beta(t-\tau))\) if \(t > \tau\), 0 otherwise. We assume for instance that \((X_t)\) is an Ornstein-Uhlenbeck process or a Brownian motion (Wiener process) [7] and

\[
dY_t = dW_t^Y + \int_{s \in \mathbb{R}} \phi(s, t-\tau) dX_{(t-s)}
\]

where \(W^Y\) and \(W^X\) are two independent Brownian mo-
tions whose increments are classically referred to as the innovation process. In the following, the parameter $\tau$ will be referred to as lag. We will consider that $(Y)$ is lagging with characteristic delay $\tau$ behind $(X)$ which is causing it.

We adopt the cross-correlogram based causality estimation approach developed in [8], in order to be consistent with Granger’s definition of causality as linear predictive ability of $(dX_{s<\tau})$ and $(dY_{s<\tau})$ for the random variable $dX_t$ [9].

Let $(X)$ and $(Y)$ be two Wiener processes. We consider that $(X)$ has a causal effect on $(Y)$ if $(dX_{s<\tau})$ is a more accurate linear predictor of $dY_t$ in square norm error than $(dY_{s<\tau})$ is an accurate linear predictor of $dX_t$. In other words $(X)$ causes $(Y)$ if and only if

$$E\left((dX_t - E(dX_t|dY_s, s < t))^2\right) > E\left((dY_t - E(dY_t|dX_s, s < t))^2\right).$$

(2)

In order to quantify the magnitude of this statistical causation, Huth and Abergel introduced in [8] the Lead-Lag Ratio (LLR) between $(X)$ and $(Y)$ as

$$LLR_{X \Rightarrow Y} = \frac{\sum_{h > 0} \rho_{XY}^2(h)}{\sum_{h < 0} \rho_{XY}^2(h)} \quad (3)$$

where $\rho_{XY}(\cdot)$ is the cross-correlation between the second order stationary processes $(X)$ and $(Y)$. The analysis conducted in [8] proved $(X)$ causes $(Y)$ is equivalent to

$$LLR_{X \Rightarrow Y} < 1$$

thereby yielding an indicator of causation intensity between processes which depends $\phi_{\tau,\phi}$ through (1).

1.1 Challenges with real world data:

Unfortunately, in practical applications, time series data sets often present three main challenges that hinder the estimation of even linear causal dependencies:

- **Irregular Sampling**: Observations are collected at irregular intervals both within and across processes complicating the application of standard causal inference techniques that rely on evenly spaced timestamps that align across processes.

- **Long Range Dependencies (LRD)**: Long range dependencies can result in increased and non vanishing variance in correlation estimates.

- **Scale**: Real-world time series are often very large and high dimensional and are therefore often stored in distributed fashion and require communication over limited channels to process.

In the following we show as in [8] that naive interpolation of irregularly sampled data may yields spurious causality inference measurements. We also prove that eliminating LRD is crucial in order to obtain consistent correlation estimates. Unfortunately, standard time domain LRD erasure requires sorting the data chronologically and is therefore costly in the distributed setting. These costs are further exacerbated by time domain fractional differentiation which scales quadratically with the numbers of samples.

To address these three critical challenge we propose a Fourier transform based approach to causal inference. Projecting on a Fourier basis can be done with a simple sum operator for irregularly sampled data as described in [10]. A novel and salient byproduct of our estimation technique is that there is no need to sort the data chronologically or gather the data of different sensors on the same computing node. We use Fourier transforms as a signal compressing representation where cross-correlations and causal dependencies can be estimated with sound statistical methods all while minimizing memory and communication overhead. In contrast to sub-sampling in which aliasing obscures short-range interactions, our methods does not introduce aliasing enabling the study of short-range interactions. An exciting aspect of compressing by Fourier transforms is that it only affects the variance of the cross-correlogram without destroying the opportunity to study inter-dependencies at a small time scale.

In section 2 we show that leveraging the frequency domain representation we present communication avoiding consistent spectral estimators [11] for cross-dependencies. We first compress the time series by projecting without interpolation or reordering directly onto a reduced Fourier basis, thereby locally compressing the data. Spectral estimation then occurs in the frequency domain prior to being translated back into the time domain with an inverse Fourier transform. The resulting output can be used to compute unbiased Lead-Lag Ratios and thereby identify statistical causation.

In section 3, we provide a method to approximately erase LRD in the frequency domain, which has tremendous computational advantages as opposed to time domain based methods. Our analysis of LRD erasure as fractional pole elimination in frequency domain guarantees the causal estimates we obtain are not spurious unlike those calculated naively on LRD processes [1, 2]. Finally, we apply these methods to synthetic data and several terabytes of real financial market trade tables.

In section 4, we present a novel analysis of the trade-off between estimator variance and communication bandwidth which precisely assesses the cost of compressing time series prior to analyzing them. A three-fold analysis establishes the statistical soundness of the contributions that address the three issues mentioned above. Studying data on compressed representations comes at an expected cost. In our setting this supplementary variance can be decreased
in an iterative manner and with bounded memory cost on a single machine. These properties cannot be replicated to the best of our knowledge by time domain based sub-sampling.

2 Interpolation and spurious lead-lag

In this section, we first review existing techniques for interpolated time-domain estimation of second-order statistics in the context of sparse and random sampling along the time axis. Interpolating data is a usual solution in order to be able to use classic time series analysis [10, 12–14]. Unfortunately it is not always suitable, as it can create spurious causality estimates and implies a supplementary memory burden.

2.1 Second order statistics and interpolated data

In order to infer a linear model from cross-correlogram estimates by solving the Yule-Walker equations [15] or to compute a LLR (Eq. (3)) one needs to estimate the cross-correlation structure of two time series. Let \((X)\) and \((Y)\) be two centered stochastic processes whose cross-covariance structure is stationary:

\[
\gamma_{XY} (h) = E(X_{t-h} Y_t).
\]  

(4)

If data is sampled regularly \((x_n \Delta t, y_n \Delta t))_{n=0...N-1}\) a consistent estimator for \(\gamma_{XY} (h)\) is:

\[
\hat{\gamma}_{XY} (h) = \frac{1}{N-h-1} \sum_{n=h}^{N-1} x_{n-h} \Delta t y_n \Delta t
\]  

(5)

(we use \(\hat{A}\) to denote an estimator for \(A\)). Classically, cross-correlation estimates can subsequently be computed as

\[
\rho_{XY} (h) = \frac{\hat{\gamma}_{XY} (h)}{\sqrt{\hat{\gamma}_{XX} (0) \hat{\gamma}_{YY} (0)}
\]  

(6)

using any consistent cross-covariance estimator.

Interpolating irregular records:

The standard consistent estimator Eq. (5) cannot be computed when \((x)\) and \((y)\) do not share common timestamps. A classical way to circumvent the irregular sampling issue is to interpolate the records \((x_{t_x} \nu_{t_x} \epsilon_{t_x})\) and \((y_{t_y} \nu_{t_y} \epsilon_{t_y})\) onto the set of timestamps \((n \Delta t))_{n=0...N-1}\) therefore yielding two approximations \((\widetilde{x}_n \Delta t))_{n=0...N-1}\) and \((\widetilde{y}_n \Delta t))_{n=0...N-1}\) that can be studied as a synchronous multivariate time series. An adapted cross-covariance estimate is then:

\[
\hat{\gamma}_{XY} (h) = \frac{1}{N-h-1} \sum_{n=h}^{N-1} \hat{\gamma}_{\widetilde{x}_{(n-h) \Delta t} \Delta t} \hat{\gamma}_{\widetilde{y}_{n \Delta t}}
\]

While there are many interpolation techniques, a commonly used method is last observation carried forward (LOCF). Note that interpolation may require substantial additional memory to render each time series at the resolution of interactions which can be millisecond scale in many crucial applications such as studying stock market interactions.

We now consider the causality inference framework introduced in [8] and show how the LOCF interpolation technique creates spurious causality estimates.

Bias in LLR with irregularly sampled data: The LLR can be computed by several methods. Cross-correlation measurements on a symmetric centered interval are sufficient statistics for this estimator. Therefore one can use synchronous cross-correlation estimates on interpolated data in order to compute the LLR. Carrying the last observation forward (LOCF) has been proven to create a bias in lag estimation in [8]. The LOCF interpolation method introduces a causality estimation bias in which a process sampled at a higher frequency will be seen as causing another process which is sampled less frequently although these correspond to Brownian motions with simultaneously correlated increments.

2.2 Interpolation-free causality assessment

The Hayashi-Yoshida (HY) estimator was introduced in [16] to address this spurious causality estimation issue. The HY estimator of cross-correlation does not require data interpolation and has been proven to be consistent with processes sampled on a quantized grid of values [17] in the context of High Frequency statistics in finance.

Correlation of Brownian motions: HY is adapted to measuring cross-correlations between irregularly sampled Brownian motions. Considering the successor operator \(s\) for the series of timestamps of a given process, let \([t, s(t)]\) be the set of intervals delimited by consecutive observations of \(x\) and \(y\) respectively. The Hayashi-Yoshida covariance estimator over the covariation of \((X)\) and \((Y)\) [7] is defined as

\[
HY_{[0,t]}(x, y) = \sum_{t \in I_x, s(t) \in I_y \cap ov(t, t')} (x_{s(t)} - x_t) \cdot (y_{s(t')} - y_{t'})
\]

(7)

where \(ov(t, t')\) is true if and only if \([t, s(t)]\) and \([t', s(t')]\) overlap. The estimator can be trivially normalized so as to yield a correlation estimate.

HY and fractional Brownian motions: No interpolation is required with HY but unfortunately this estimator is only designed to handle full differentiation of standard Brownian motions. Figure 2 shows how HY fails to estimate cross-correlation of increments on a fractional Brownian motion whereas the technique we present succeeds. In the following, we show how our frequency domain based analysis naturally handles irregular observations and is able to fractionally differentiate the underlying continuous time process. This is in particular necessary when one studies fractional Brownian motions with correlated increments. In the interest of concision, we refer the reader to [18] for the
Indeed, the power spectrum basis that allows us to compute cross-covariance estimates is based on the definition of the Fourier transform of a stochastic process. Considering a continuous time stochastic process \( \{X_t\}_{t \in [0, T]} \) and a frequency \( f \in [0, 2\pi] \), the Fourier projection of \( X \) for the frequency \( f \) is defined as

\[
P_f(X) = \int_0^T X_t e^{-i ft} dt \quad (8)
\]

where \( i \) is the imaginary number. Much attention has been focused on the benefits of the FFT algorithm which has been designed for the very particular base of ordered and regularly sampled observations. Our key insight is to go back to the very definition of the Fourier transform as an integral and express it empirically in summation form \([10, 11]\). Moreover, if the process \( X \) is observed at times \((t_1, \ldots, t_N)\), one can estimate the Fourier projection by

\[
\hat{P}_f(x) = \sum_{n=1}^N x_{tn} e^{-ift_n}. \quad (9)
\]

Therefore, we propose the following simple framework for frequency domain based linear causal inference:

1. Project \( x \) and \( y \) on to a reduced Fourier basis.
2. Estimate the cross-spectrum of \( X \) and \( Y \) in the frequency domain.
3. Apply the inverse Fourier transform to the cross-spectrum to recover the cross-correlogram and infer the linear causal structure.

The intuition behind this estimation method is a change of basis that allows us to compute cross-covariance estimates without needing to address the irregularity of timestamps. Indeed the power spectrum \( f(\cdot) \) is the element-wise Fourier transform of

\[
\gamma(\cdot) = \begin{bmatrix}
\gamma_{XX}(\cdot) & \gamma_{XY}(\cdot) \\
\gamma_{XY}(\cdot) & \gamma_{YY}(\cdot)
\end{bmatrix}.
\]

Therefore, in order to estimate this function one may infer what corresponds to its frequency domain representation and then compute the inverse Fourier transform of the result.

**Projecting onto Reduced Fourier Basis:** We first project \((X)\) and \((Y)\) onto the elements of the Fourier basis of frequencies \((l \Delta f)_{l=0, \ldots, P}\), namely the pair \((P_l \Delta f(X))_{l=0, \ldots, P}\) and \((P_l \Delta f(Y))_{l=0, \ldots, P}\). By projecting onto a single relatively small set of orthonormal functions, we are able to compress and effectively re-align the observations \((x)\) and \((y)\). In practice using only a few thousand basis functions we are able to accurately recover the cross-correlogram. Finally, this computation is sufficiently fast to execute interactively on a single laptop and can be easily expressed using the map-reduce framework.

**Estimating the Cross-spectra:** Computing projections onto a reduce Fourier basis enables exploratory data analysis through the study of the cross-spectrum of \((X)\) and \((Y)\)

\[
(I_{XY}(l \Delta f))_{l=0, \ldots, P} = (P_l \Delta f(X) \times P_l \Delta f(Y))_{l=0, \ldots, P}. \quad (10)
\]

An inconsistent estimator for the cross-spectrum is:

\[
(I_{XYh}(l \Delta f))_{l=0, \ldots, P} = \left(\hat{P}_l \Delta f(x) \times \hat{P}_l \Delta f(y)\right)_{l=0, \ldots, P}. \quad (11)
\]

Local averaging of Eq. (11) with respect to frequencies is widely used \([10, 11, 15]\) in cross-spectral analysis to identify the characteristic frequencies at which stochastic processes interact although they are observed at irregular times. Unfortunately, to compute characteristic delays or LLR (crucial steps in linear causal inference) we still need to estimate the cross-correlogram.

**Estimating the Cross-correlogram:** To estimate the cross-correlogram we can take the inverse Fourier transform of the cross-spectrum \((I_{XY}(l \Delta f))_{l=0, \ldots, P}\) which translates frequency analysis back into the time domain:

\[
\gamma_{XY}^P(h) = \frac{1}{P} \sum_{l=0}^P I_{XY}(l \Delta f) e^{i l \Delta fh}. \quad (12)
\]

Using the following consistent estimator:

\[
\gamma_{XYh}^P(h) = \frac{1}{P} \sum_{l=0}^P \hat{I}_{xy}(l \Delta f) e^{i l \Delta fh} \quad (13)
\]

of the cross-covariance we can directly compute a consistent estimator of the cross-correlation using equation Eq. (6). The cross-correlation between \((X)\) and \((Y)\) can now be estimated in the time domain with a discrete grid \(G_h\) of lag values ranging from \(-L \Delta h\) to \(L \Delta h\) with a resolution \(\Delta h\). As expected, aliasing will occur if the user specifies a resolution in the cross-correlation estimate that is
much higher than the average sampling frequency of the time series [10].

In contrast to more cumbersome time domain synchroniza-
tion relying on interpolation based methods (LOCF) or in-
terval matching based estimations (HY), our method ele-
gantly addresses time synchronization in the frequency do-
main. While earlier work [10,11] has considered the ap-
plication of frequency domain analytics to irregularly sam-
pled data, our method is the first to translate back to the
time domain to recover a consistent estimator of the linear
causal structure. Alternatively, Lomb-Scargle periodogram
[19,20] also enables the frequency domain analysis of ir-
regularly observed data but suffers from the supplementary
cost of a least square regression. To the best of our knowl-
edge we are the first to use frequency domain projections
in order to compute the cross-correlogram in order to infer
linear causal structure.

### 2.4 The statistical cost of compression

Central to the communication and memory performance of
our technique is the ability to use a small number of Fourier
projections relative to the number of observations and still
accurately recover the cross-correlogram.

**Cross-correlogram Estimator Consistency:** We can char-
acterize the statistical properties of the cross-spectral esti-
mator [10,11,15]. In particular, it is well known that for
two distinct non-zero frequencies \( f_1 \) and \( f_2 \) the esti-
mators \( I_{XY}(f_1) \) and \( I_{XY}(f_2) \) are asymptotically independent.
Consequently, to obtain an estimator with variance \( O(V) \)
the user will need to project on \( \frac{1}{V} \) frequencies. We con-
firm this result numerically in Figure 6. The element-wise
product of Fourier transforms is converted into the time
domain by the inverse Fourier transform to yield a cross-
correlogram. With very large datasets in which \( N \gg \frac{1}{V} \)
we obtain the suitable compression property of our algo-

**Issues with Non-smooth Cross-correlograms:** As ex-
pected, deterministic lags or seasonal components can re-
sult in Fourier compression artifacts in the inverse Fourier
transform. However, statistical estimation and removal of
these deterministic components is standard in time series
analysis [11,15]. In the context of estimating non-trivial
stochastic causal relationships (e.g., social networks, pairs
of stock prices the financial markets, cyberphysical sys-
tems) random perturbations affect the causation delay. In
these settings, the theoretical cross-correlation function is
smooth. As a consequence, a few Fourier projections suf-
fice to accurately represent the cross-correlogram in fre-
quency domain.

### 2.5 Example of time domain exploratory data
analysis through the frequency domain

The time domain exploratory analysis we enable makes
lead-lag relationships self-explanatory as shown in Fig-
ure 1. We show in the following that it is not hindered by
biases related to the fact that one process is sampled more
seldom than the other.

**Numerical assessment of frequency domain based cor-
relation measurements:** We demonstrate, through simu-
lation, that the spurious causation issue that plagues the
LOCF interpolation [8] does not appear in our proposed
method. We consider two synthetic correlated Brownian
motions that do not feature any lead-lag and compare the
estimation of LLR provided by two time domain interpola-
tion methods and our approach. After having sampled these
at random timestamps, in Table 1 and Figure 3 we compare
the cross-correlation and LLR estimates obtained by LOCF
interpolation and our proposed frequency domain analysis
technique confirming that our method does not introduce
spurious causal estimation bias.

<table>
<thead>
<tr>
<th>( \frac{N_1}{N_2} )</th>
<th>LOCF interpolation LLR</th>
<th>Fourier transform LLR</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Avg + std</td>
<td>Avg + std</td>
</tr>
<tr>
<td>1</td>
<td>0.998 + 0.135</td>
<td>1.021 + 0.166</td>
</tr>
<tr>
<td>4.5</td>
<td>6.863 + 1.678</td>
<td>1.053 + 0.320</td>
</tr>
<tr>
<td>10</td>
<td>7.277 + 1.854</td>
<td>1.107 + 0.391</td>
</tr>
</tbody>
</table>

Table 1: Comparison of LLR ratios with LOCF and Fourier
transforms (1000 projections) for simultaneously correlated
Brownian motions with different sampling frequencies. The LLR
ratios should all be 1, one can observe the bias in the LOCF
method.

### 3 The Long Range Dependence (LRD) Issue

A stochastic process is said to be long range dependent
if it features cross-correlation magnitudes whose sum is
infinite [1]. Many issues arise in that case with correla-
tion estimates becoming spurious. This phenomenon was
first discover when Granger studied the concept of cointe-
gration between Brownian motions (integrated time series)
[2]. On sorted Brownian motion data, this effect can be ad-
dressed by differentiating the time series, namely comput-
ing \( (\Delta X_t)_{t \in \mathbb{Z}} = (X_t - X_{t-1})_{t \in \mathbb{Z}} \). For fractional Brown-
ian motion and LRD time series, the fractional differentia-
tion operator needs to be computed. It is defined as

\[
(\Delta^\alpha X_t)_{t \in \mathbb{Z}} = \left( \sum_{h=0}^{\infty} \prod_{j=0}^{b-1} (\alpha - j)(-X_{t-h})^h \right)_{t \in \mathbb{Z}}.
\]

Therefore, to study the cross-correlogram structure of two
integrated or fractionally integrated time series, one would
have to compute \( (\Delta^\alpha X_t)_{t \in \mathbb{Z}} \) or \( (\Delta^\alpha X_t)_{t \in \mathbb{Z}} \). The latter re-
quires chronologically sorted data and synchronous times-
Erasing memory through fractional pole elimination:

The power spectrum of a fractional Brownian motion [18] with Hurst exponent $H$ is asymptotically $\frac{1}{f^{2H+1}}$ for $f \ll 1$. This is the characteristic spectral signature of a long range dependent time series. $H$ can therefore be estimated by the classical periodogram method for an individual time series by conducting a linear regression on the magnitude of the power spectrum about 0 in a log/log scale [21]. Wavelets are another family of orthogonal basis enabling a similar estimation [22]. One can therefore see the fractional differentiation operator of order $H + 1/2$ as a means to compensate for a pole of order $2H + 1$ in square magnitude in 0. Multiplying the Fourier transform of the signal by $(i f)^{H+1/2}$ eliminates the issue. It does not require any preprocessing of the data, no interpolation or re-ordering and we will show below that it has tremendous computational advantages in the context of distributed computing in terms of communication avoidance.

An approximation of differentiation in the case of discrete observations: It is noteworthy though that this method is intrinsically approximate in the practical context of discrete sampling. Indeed the multiplication rule for differentiation in frequency domain we proved in the context of discrete sampling:

3.2 Testing frequency domain LRD erasure

The example below considers two fractional Brownian motions $(X)$ and $(Y)$ Brownian motions with Hurst exponent $H = 0.4$ [1]. We compare the empirical distributions of cross-correlation estimates obtained over 100 trials with and without LRD erasure in frequency domain. In Figure 4 we showcase an experiment with 9998 uniformly random observations for $(X)$ and 6000 uniformly random observations for $(Y)$. While naive cross-correlation estimations lead to many spurious cross-correlation estimates with significantly high magnitudes of estimated correlation values for processes that are in fact independent, (90%) of the empirical distribution between $-0.9$ and $0.9$ the confidence interval we obtain with our novel frequency domain erasure method by fractional pole elimination is narrower (90% of the empirical distribution between $-0.05$ and $0.05$) and enables reliable analysis. The next section will expose the computational advantages of such a frequency domain based estimation as a communication avoidance mechanism.

4 Distribution

Scalable computation is essential to practical causal inference in real-world big data sets. Our proposed frequency domain approach provides a parallel communi-
culation avoiding mechanism to efficiently compress large time-series data sets while still enabling the estimation of cross-correlograms.

4.1 Computational setting

Cluster computing presents the opportunity to enable faster analysis by leveraging the scale-out compute resources in modern data-centers. However to leverage scale-out cluster computing it is essential to minimize communication across the network as network latency and bandwidth can be orders of magnitude slower than RAM [23, 24].

4.2 Computational advantages

One mechanism to distribute time domain analysis of time series is to construct overlapping blocks as described in [25]. However, this technique only works if there is no LRD. The need to specify the appropriate replication padding duration at preprocessing time makes it difficult to switch between the time scales at which cross-dependencies are computed.

The novel frequency domain based methods we propose can entirely be expressed as trivial map-reduce aggregation operations and do not require sorting or interpolating the data. Indeed, the use of projections on a subset of a Fourier basis only requires element-wise multiplication and then an aggregated sum to construct a unique concise signature in frequency domain. The empirical cross-correlation distribution on the left is affected by high magnitude spurious estimates. On the right, frequency domain fractional pole erasure eliminated the issue, considerably narrowing down the interval between the 5th and 95th percentiles.

Figure 4: Spurious cross-correlation is erased by pole elimination in frequency domain. The empirical cross-correlation distribution on the left is affected by high magnitude spurious estimates. On the right, frequency domain fractional pole erasure eliminated the issue, considerably narrowing down the interval between the 5th and 95th percentiles.

4.3 Fourier compression as a communication avoidance algorithm

The computation of Fourier projections is communication efficient in the distributed setting. The Fourier projection can be calculated by locally computing the sum of the mapping of multiplications by complex exponentials. Then, only the local partial sums need to be transmitted across the network to compute the projections of the entire data set. In this section, we study $d$ distinct processes with $N$ data points each. Let $V$ denote the desired variance for the cross-correlation estimator via the frequency domain.

Communication cost of aggregation with indirect frequency domain covariance estimates: Now consider the set of Fourier projections $\left( \hat{P}_f(x) = \sum_{n=1}^{N} x_n e^{-jft_n} \right)_{f=0,\Delta f,...,F\Delta f}$ which we aggregate on each single machine separately prior to sending them over the network. The number of projections needed to have an estimator for cross-correlation with variance $V$ is $O\left( \frac{V}{d} \right)$. Therefore, the size of the message sent out by each machine over the communication medium is now $O\left( \frac{dV}{d} \right)$ and representative of $O(dN)$ data points. If the user chooses $\frac{1}{d} \ll N$, our method effectively compresses the data prior to transmitting it over the network. It is noteworthy that the gain offered by this algorithm is system independent as long as communication between computing cores is the main bottleneck.

Distributed LRD erasure: The computational complexity of fractional differentiation (Eq. (14)) is $O(N^2 d)$ in the time domain. Furthermore, due to LRD, time domain fractional differentiation cannot be accomplished using the overlapping partitioning strategy proposed in [25]. Moreover, in distributed system, computing the fractional differentiation of a signal would require transmitting the entire data set across the network. As a consequence the bandwidth needed is $O(Nd)$.

Alternatively, fractional differentiation in the frequency domain is both computationally efficient and easily parallelizable. Once the Fourier transforms have been computed the now substantially compressed frequency domain representation can be collected on a single machine for further analysis. We then proceed with the elimination of fractional poles by a simple element-wise multiplication. No supplementary communication is needed to erase LRD and therefore the size of the data transmitted across the network is just $O\left( \frac{1}{d} \right)$ as opposed to $O(Nd)$. This remarkable improvement in communication requires only a modest computational cost of $O\left( \frac{1}{d} \right)$ projections per data point on slave machines.

The compute time therefore allows an interactive experience for the user and becomes even shorter with a distributed implementation on several machines. For example, on a single processor with a 2013 MacbookPro Retina we...
were able to compute 3000 projections on $10^5$ samples in roughly a minute.

<table>
<thead>
<tr>
<th>Method</th>
<th>Time complexity on slave</th>
<th>Communication size</th>
</tr>
</thead>
<tbody>
<tr>
<td>Time domain</td>
<td>$O(Nd^2)$</td>
<td>$O(Nd + d^2)$</td>
</tr>
<tr>
<td>Fourier projection</td>
<td>$O(Nd^{1/2})$</td>
<td>$O(d^{1/2})$</td>
</tr>
</tbody>
</table>

**Memory:** A potential concern with the frequency domain approach is that the aggregation of the Fourier projections to a single device could exceed the device’s memory. The device will have to store projections of size $O(\frac{1}{\sqrt{d}})$, compute element-wise products with time complexity $O(d^{2/3})$, and store the cross-correlation estimates in a memory container of size $O(d^2)$. In particular, the maximum size of the memory needed by the algorithm on the master is $O(\frac{1}{\sqrt{d}})$ which is small relative to the size of the data set in our current setting. Indeed, we assume that $N$ is large enough and therefore $\frac{1}{\sqrt{d}} << N$.

## 5 Causality estimation on actual data

Identifying leading components on the stock market is insightful in terms of assessing which stocks move the market and highlights the characteristic latency of trading reactions. Consider two stocks, for instance AAPL and IBM (shares of Apple Inc. and IBM traded in the New York stock exchange). The trade and quote table of Thomson Reuters records all bids, asks, trades in volume and price. It is therefore interesting to check if there is a causation link between the price at which AAPL is traded as compared to that of IBM. In particular, if we see an increase in the price of the former can we expect an increase shortly after in the later? With which delay? Weak causation or short delay indicates an efficient market with few arbitrage opportunities. Significant causation and longer delays would enable high frequency actors to take advantage of the causal empirical relationship in order to conduct statistical arbitrage [3].

### 5.1 Causal pairs of stocks

One critical application of the generic method we present is identifying which characteristic delays the NYSE stock market features as well as Lead-Lag ratios between pairs of stocks. Lead-Lag ratios that are significantly different from 1.0 indicate that changes in the price of one stock trigger changes in the price of another. This indicates pair arbitrageurs are most likely using high frequency arbitrage strategies on this pair of stocks.

### 5.2 Using Full Tick data

In order to highlight significant cross-correlation between pairs of stocks, one needs to consider high frequency dynamics. As we will show in the following, cross-correlation vanishes after a few milliseconds on most stocks and futures. In these settings it is then necessary to use full resolution data which in this instance comes in the form of Full Tick quote and trade tables (TAQ). These TAQ tables record bids, asks and exchanges on the stock market as they happen. The timestamps are therefore irregular and not common to different pairs of stocks. Also, stock prices are Brownian motions and therefore feature long memory. This context is therefore in the very scope of data intensive tasks we consider. We show our novel Fourier compression based cross-correlation estimator provides consistent estimates in this setting.

### 5.3 Checking the consistency of the estimator

Consider ask and bid quotes during one month worth of data. We create a surrogate noisy lagged version of AAPL with a 13ms delay and 91% correlation which is named AAPL-LAG. We study four pairs of time series: APPL/APPL-LAG, AAPL/IBM, AAPL/MSFT, MSFT/IBM. We study the changes in quoted prices (more exactly, volume averaged bid and ask prices). We obtained quote data for these stocks at millisecond time resolution representing several months of trading. We removed observations with redundant timestamps. The cross-correlograms obtained below are computed between 10 AM and 2PM for 61 days in January, February and March 2012. For each process, 3000 frequencies were used in the Fourier basis which is several orders-of-magnitude less than the number of observations that we get per day which ranges from $5 \times 10^4$ to $1 \times 10^5$. The estimate cross-correlograms in Figure 5 and their empirical significance intervals show that our estimator is consistent and does not suffer from non-vanishing variance as a result of LRD. We observe an 89% average peak cross-correlation with an 8ms delay for the surrogate pair of AAPL stocks which confirms our estimator is reliable with empirical data. While we observe the Fourier compression artifacts, these only occur because our surrogate data features a deterministic delay. They do not affect pairs of actual observed processes. In Figures 5 and 6 we highlight a taxonomy of causal relationships and show in particular that with our definition of causality anchored in linear predictions, a process may cause another one without any significant delay. This may also be symptomatic of a delay shorter than the millisecond resolution of our timestamps.

### 5.4 Choosing the number of projections

In order to guide practitioners in their choice of the number of Fourier basis elements to project onto, we conduct a numerical experiment on actual data. We compute an empirical standard deviation of the daily cross-correlogram obtained in January 2012 (19 days) for JPM (JP Morgan Chase) and GS (Goldman Sachs) with 10, 100, 1000 and 10000 projections. Figures 6 and 8 show that, as expected, the variance decreases linearly with the number of projec-
Figure 5: Average of daily cross-correlograms pairs of stock trade and quote data. Compression ratio is $< 5\%$. We retrieve lag and correlation accurately on surrogate data. The daily averaged cross-correlogram of AAPL and IBM is strongly asymmetric, therefore highlight that AAPL causes IBM. The symmetry between AAPL and MSFT shows there is no such relationship between them. Finally, symmetric and offset in correlation peak show that MSFT causes IBM with a millisecond latency.

5.5 Studying causality at scale

A primary goal of this work is to enable practical scalable causal inference for time series analysis. To evaluate scalability in a real-world setting in which $\frac{1}{V} << N$, we assess the relation between AAPL and MSFT over the course of 3 months. In contrast to our earlier experiments (shown in Figure 5), we no longer average daily cross-correlograms in and therefore only leverage concentration in the inverse Fourier transform step of the procedure. With only 3000 projections for $5 \times 10^6$ observations per time series, the results we obtain on Figure 7 reveals the causal relation between AAPL, AAPL-LAG, IBM and MSFT consistently with Figure 5.

Scalability: In order to assess the scalability of the algorithm in a situation where communication is a major bottleneck, we run the experiment with Apache Spark on Amazon Web Services EC2 machines of type r3.2xlarge. In Figure 8 we show that even with a large number of projections (10000) the communication burden is still low enough to achieve speed-up proportional to the number of machines used.

6 Conclusion

Time series analysis via the frequency domain presents several presents unique opportunities in terms of providing consistent causal estimates and scaling on distributed systems. We proposed a communication avoiding method to analyze causality which does not require any sorting or joining of data, works naturally with irregular timestamps without creating spurious causal estimates and makes the erasure of Long-Range dependencies embarrassingly parallel. Our approach is based on Fourier transforms as compression operators that do not modify the second order properties of stochastic processes. Applying an inverse Fourier transform to the resulting estimated spectra enables exploration of dependencies in the time domain. With the resulting consistent cross-correlogram, one can compute Lead-Lag ratios and characteristic delays between processes thereby infer linear causal structure. We show that projecting onto 3000 Fourier basis elements is sufficient to study stock market pair causality with tens of millions of high frequency recordings, thereby providing insightful analytics in a generic and scalable manner.
Figure 8: On the left we plot the empirical standard deviation of daily cross-correlograms (Figure 6) with respect to the number of projections showing that the variability decreases rapidly. On the right we plot the run time performance of our algorithm versus the number of Apache Spark EC2 machines demonstrating approximately linear speedup. The small number of projections ($10^4$) relative to the size of the data set ($10^7$ records) avoids communication.

7 Future work

Our future work will focus on using the Yule-Walker equations so as to turn the cross-correlation estimates we presented into linear predictive models. Confidence bands should also be obtained for the estimators thanks to an asymptotic distribution study akin to the proof of the Bartlett formula.

References


