## Gibbs Distribution

$$
\pi(\omega)=\frac{1}{z} e^{\frac{-U(\omega)}{T}}
$$

## Proving Gibbs Distribution Implies Markov Random Field

1. $P(X=\omega)>0$

Trivial because of exponential
2. $P\left(X_{s}=x_{s} \mid X_{r}=x_{r}, r \neq s\right)=P\left(X_{s}=x_{s} \mid X_{r}=x_{r}, r \in G_{s}\right)$

$$
\begin{aligned}
& P\left(X_{s}=x_{s} \mid X_{r}=x_{r}, r \neq s\right) \\
& =\frac{P\left(X_{s}=x_{s}, X_{r}=x_{r}, r \neq s\right)}{P\left(X_{r}=x_{r}, r \neq s\right)} \\
& \vdots \\
& =\frac{e^{-U(\omega) / T}}{\sum e^{-U\left(\omega^{*}\right) / T}}
\end{aligned}
$$

Look at the handout given in class for a detailed proof.
In summary, at the end of the day, sites without s cancel.

## Back to the Image Segmentation Problem

Our image model

line process $l$ not observed
Goal

- Given observed image g , find a probability distribution of true image $f$, and the line process $l$.

The line process estimate solves the image segmentation problem.
The true image estimate solves the image restoration problem.
Both problems are simultaneously solved!

- Note: only works for piecewise smoothe images $\Rightarrow$ no textures

Our model assumes:

$$
\begin{aligned}
& P(F=f, L=l)=\frac{e^{-U(f, l) / T}}{Z} \\
& g=f+\eta
\end{aligned}
$$

Our solution:

$$
\begin{aligned}
& P(X=\omega \mid G=g) \\
& =\frac{P(X=\omega, G=g)}{P(G=g)} \\
& =\frac{P(G=g \mid X=w) P(X=\omega)}{P(G=g)}
\end{aligned}
$$

Interesting term:

$$
\begin{aligned}
& P(G=g \mid X=\omega) \\
& =P(n=f-g) \\
& =\frac{1}{(\sqrt{2 \pi} \sigma)^{2}} \exp \left(\frac{-\|\mu-(g-f)\|^{2}}{2 \sigma^{2}}\right)
\end{aligned}
$$

Where we assume every pixel has independent noise $\eta \sim N(\mu, \sigma)$ e.g. Poisson process noise in CCDs

Result:

$$
\begin{aligned}
& P(X=\omega \mid G=g) \\
& =\frac{e^{-U(f, l)+t e r m / T}}{Z} \\
& =\frac{e^{-U_{p}(f, l) / T}}{Z}
\end{aligned}
$$

is the posterior probability of a particular $\mathrm{f}, \mathrm{l}$ given g . Note this is also a Gibbs distribution!

## MAP (maximum a posteriori) Estimate

If you insist on a single answer then return $\mathrm{f}^{*}, 1^{*}$ that maximizes

$$
P\left(X=\omega^{*} \mid G=g\right)=\frac{1}{Z} \exp \left(\frac{-U_{p}\left(f^{*}, l *\right) / T}{Z}\right)
$$

or equivalently, minimizes the energy function

$$
U_{p}\left(f^{*}, l^{*}\right)
$$

Problem: $\quad f, l$ space is very large!!
Solution: $\quad$ Construct samples of $f, l$ in this space with high probablity
Technique: Markov Chain Monte Carlo (MCMC) lets you sample the posterior distribution

## Sampling a Distribution

Q: How do we represent a probability distribution with sampling?
A: Create many samples drawn from that distribution and count!
Example:
Q: $\quad P(X>17)=$ ?
A: Create samples $X_{i}$ drawn from the distribution of X.

$$
P(X>17)=\frac{N\left(x_{i}>17\right)}{N_{\text {total }}}
$$

Count the number of samples greater than 17 and divide by total number of samples.

## Generating the Samples

Primitive random number generator $\mathrm{X} \sim \mathrm{U}(0,1)$.
To create $\mathrm{Y} \sim \mathrm{U}(\mathrm{a}, \mathrm{b})$ use
$Y=a+(b-a) X$
In general we can use the cumulative distribution function


1987 - Stochastic Simulation (Ripley) for generating samples for "standard stuff" in textbooks

## Markov Chain Monte Carlo (MCMC) Technique

- Define a suitable Markov Chain whose equilibrium distribution is the desired posterior distribution
- Generate samples from the Markov Chain


## Markov Chain Basics

$$
P\left(X_{t}=y \mid X_{t-1}, X_{t-2}, \cdots, X_{0}\right)=P\left(X_{t}=y \mid X_{t-1}\right)
$$

Example:

$$
\begin{aligned}
& P\left(X_{t}=3 \mid X_{t-1}=2\right)=1 / 2 \\
& P\left(X_{t}=1 \mid X_{t-1}=2\right)=1 / 2 \\
& P\left(X_{t}=4 \mid X_{t-1}=2\right)=0 \\
& P\left(X_{t}=2 \mid X_{t-1}=2\right)=0
\end{aligned}
$$

This transition probabilities can be written as a matrix
$P=\begin{aligned} & 1 \\ & 2 \\ & 3\end{aligned}\left[\begin{array}{cccc}1 / 2 & 1 / 2 & 0 & 0 \\ 1 / 2 & 0 & 1 / 2 & 0 \\ 0 & 1 / 2 & 0 & 1 / 2 \\ 0 & 0 & 1 / 2 & 1 / 2\end{array}\right]$

If we write the probability distribution at time $t$ as $p(t)$ then
$\mathrm{p}(\mathrm{t}+1)=\mathrm{p}(\mathrm{t}) \mathrm{P}$
For example if the drunk's walk starts at position 2 we denote $\mathrm{p}(\mathrm{t}=0)=\left[\begin{array}{lll}0 & 1 & 0\end{array}\right]$
$\mathrm{p}(\mathrm{t})$ is an evolving probability distribution which is a row vector that sums to one
The equilibrium distribution $p$ (infinity) $=\pi$ satisfies

$$
\pi \mathrm{P}=\pi
$$

and is a left eigenvector of P with eigenvalue one.
Finding the Markov Chain Corresponding to the Posterior Distribution
Metropolis Sampler - Rosenberg, Teller, Teller
Heat Bath ( Gibbs Sampler ) - "rapidly mixing" determines convergence rate
Metropolis Sampler
We are given that the posterior distribution is of the form $\mathrm{f}(\mathrm{x}) / \mathrm{Z}$

1. We have a proposal kernal satisfying $K(x, y)=K(y, x)$
2. Calculate $f(y)$
3. Accept transition with probability $=\min \{1, f(y) / f(x)\}$
4. This eventually converges to the "right thing"
